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13. ABSTRACT (Maximum 200 Words) The goals of this IDEA award are two-fold, namely, to develop and refine statistical methodology for the analysis of a data set where multiple outcomes or disease incidence endpoints are of interest and to apply these methods to the analysis of a data set in prevention such as the National Surgical Adjuvant Breast and Bowel Project's Breast Cancer Prevention Trial. The software for the fitting of the B-spline models, the first type of model that was proposed in this study is now being used to fit these models and the simulation studies are also complete. Software development for the pseudospline approach is nearly complete. Extensions to the analysis of recurrent event data outcomes are now being developed.			
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INTRODUCTION

This goal of this work is to extend the current statistical methodology for the analysis of multiple outcome data, to include a more flexible class of modeling techniques. The method proposed by Wei, Lin, and Weissfeld (1989) has been found to be an extremely flexible and useful statistical method for the analysis of multiple outcome data. However, this method suffers from the limitation that each outcome must follow the proportional hazards model. In the extension of the methodology that is developed through the work on this grant, we present an extension of the method of Wei, Lin and Weissfeld (1989) that incorporates a spline based version of the Cox proportional hazards model that is proposed by Gray (1992). The advantage of this approach is that the proportionality assumption is not needed for the modeling of each outcome. In fitting this model to the outcome data, one obtains a more detailed view of the relationship between failure time and the given covariates of interest.

BODY:

The work included in the statement of work involves several components, the development of flexible marginal models for multiple time to event data using penalized B-spline based models, to extend these models using pseudosplines, and the development of regression diagnostics and goodness-of-fit tests for these models. Substantial progress has now been made on several of these aims.

The investigators, Dr. Kiros Berhane and Dr. Lisa Weissfeld, are at the University of Sourthern California and the University of Pittsburgh, respectively. There is a graduate student research at the University of Pittsburgh site who works closely with both Dr. Berhane and Dr. Weissfeld. This individual, Zekarias Berhane has been working on the grant since its beginning. He is now very experienced in the use of the software that is needed and has been instrumental in its development. There has been one meeting over the past year. This meeting took place in August at the Joint Statistical Meetings in Atlanta, GA. Dr. Berhane, Mr. Berhane and Dr. Weissfeld were all present at this meeting and spent approximately 20 hours together during this time period working on the software and refining the proposed methodology.

Throughout much of the past year the work on the grant has required a substantial portion of the investigators' time. There have been several problems in the implementation of the methodology that turned out to be more difficult to solve than was initially anticipated. These problems centered around the testing of the methodology, the estimation of the "meat" of the sandwich estimator for the variance covariance matrix of the proposed estimator, and the distribution of the test statistic needed for drawing inferences based on the model. A second Ph.D. student of Dr. Weissfeld, Mr. Zdenek Valenta has also been working on parts of the grant as a topic for his Ph. D. dissertation.

Specific Aim 1:

The goal of this aim is to develop flexible marginal models for multiple time to event data using penalized B-spline based models. We experienced several problems related to the completion of this aim over the past year. These problems centered on the refinement of the software and the methodology. We had to solve two major problems with respect to the methodology development over the past year: the estimator of the meat of the sandwich for the variance

covariance matrix and the form of the test statistic used for inferences drawn from the model. Based on the simulation results that are presented in the attached paper, the form of the estimator is now correct. As part of this development, we also relied on joint work that Dr. Weissfeld was doing with Mr. Valenta as part of his Ph.D. dissertation. This work is used in justifying the form of the variance covariance matrix as given in this paper. The goal of this work was the development of an appropriate survival estimator for Gray's model. This estimator is then used to estimate the cumulative hazard function, which is needed for the meat of the sandwich estimator. This work will appear in Statistics in Medicine, was presented at the ENAR meetings in Charlotte, NC in March of 2001 winning a student travel award. The second problem encountered in the development of the methodology is the derivation of the appropriate test statistic for forming inferences based on the model. This problem is now also solved. The software is now developed and running so that the method can be executed.

The major goal of this specific aim is now virtually complete and we are putting the final touches on the paper. In addition to the work initially proposed for this aim, there have been several additional pieces of work that are currently underway:

- a. the extension of the method of Wei, Lin and Weissfeld (1989) and Andersen and Gill (1982) for the modeling of recurrent event data. This work is being done by Zekarias Berhane as part of his Ph.D. dissertation and will be presented at the ENAR meetings in March 2002. This is potentially important work since the WLW method based on the Cox proportional hazards model does not perform well for the modeling of recurrent data. This poor performance is due to the lack of proportional hazards in the margins causing the Cox based approach to break down. Use of Gray's model for the margins should improve on this method.
- b. The development of an estimator for the survival curve. This piece of work is needed for the meat of the sandwich estimator. This work is complete and will appear in Statistics in Medicine.
- c. An examination of the power of the WLW method based on Gray's model. Some preliminary work has been done for this piece and may form the basis for a future grant.
- d. Inclusion of the model with time-varying coefficients. This work is analogous to that done for aim 1.

Specific Aim 2:

The goal of this aim is to develop flexible marginal models for multiple time to event data using pseudospline based models for the time to event data. This piece of work was delayed due to the problems encountered in implementing Aim 1. While preliminary software development and a draft of a manuscript are underway, much of this work was held up by the problems encountered in developing Aim 1. Since these problems have now been fixed, the work on Aim 2 will proceed in a timely fashion and should be completed in the very near future. Many of the issues discussed in Aim 1 are also important for the pseudospline based model.

Note that this work is now very close to completion.

Specific Aim 3:

We have begun some work on this aim; however, because of the work of aim 1, we have had little time to pursue this avenue of research. Mr. Berhane has met with Dr. Chang, who developed the projected and recursive residuals for the Cox model. As details are wrapped up from Aim 1, this work will take over part of Mr. Berhane's time as the GSR.

Specific Aim 4:

This work is currently underway and will serve as the second paper for Mr. Valenta's dissertation. Mr. Valenta is currently working on the extension of several different approaches for application to Gray's model.

KEY RESEARCH ACCOMPLISHMENTS:

The key research accomplishments to date from this work are:

- A program for running the multiple outcomes model based on the spline-based version of Cox's model as proposed by Gray.
- A preliminary version of a program for the multiple outcomes model based on the pseudo-spline based model.
- Software to run simulations for aim 1. The results of the simulation study are in the attached manuscript "Modeling Multiple Time-to-Event Data Using Penalized B-splines".
- Submission of the manuscript "Modeling Multiple Time-to-Event Data Using Penalized B-splines" to the NSABP group for preliminary approval before submission to a journal.
- Publication of the manuscript "Estimation of the Survival Function for Gray's Piecewise-Constant Time-Varying Coefficients Model" in Statistics in Medicine.
- Preliminary development of software for the modeling of recurrent event data using the approach of Andersen and Gill (1982).

REPORTABLE OUTCOMES:

- Attached manuscript "Modeling Multiple Time-to-Event Data Using Penalized B-splines" which has been submitted for internal NSABP review.
- Attached manuscript "Estimation of the Survival Function for Gray's Piecewise-Constant Time-Varying Coefficients Model" which is to appear in Statistics in Medicine.
- Presentation of "Flexible Models for Multiple Time-to-Event Data Using Penalized B-Splines" at the Joint Statistical Meetings in Atlanta in August 2001.
- Attached program for computing the estimators presented in the manuscript "Modeling Multiple Time-to-Event Data Using Penalized B-splines".

CONCLUSIONS:

This work provides researchers with another tool for the analysis of multiple outcome survival data. The advantage of this work is that the underlying modeling technique allows for greater flexibility when specifying the relationship between time to event and a given covariate. This is particularly applicable for the risk stratification variable used in the NSABP BCPT. For this variable the level of risk is quite different for individuals with a risk score of 10 or greater versus individuals with a risk score of less than 10. This illustrates the potential usefulness of this approach for the analysis of survival data. The analysis of the multiple outcomes verifies the fact that endometrial cancer is a significant side effect for women using tamoxifen for breast cancer prevention.

The work on Aim 1 for the grant is essentially complete. However, this work has lead to many new ideas that are being pursued through other venues. For example, the graduate student researcher, Zekarias Berhane, will be examining extensions to the recurrent event problem based on this work. Mr. Valenta also completed work on a survival function estimator that was used for formulating the variance-covariance estimator for the WLW extension. We will also spend time examining the properties of the proposed test statistics under various scenarios, focusing on power. The work for Aim 2 is now well along. Aim 4 will serve as a dissertation topic for Mr. Valenta. He is currently pursuing goodness-of-fit methodology for Gray's model and we are in the process of identifying the appropriate approach to this problem. We will be pursuing the residual analysis in the coming months.

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**APPENDIX 1. COPY OF “MODELING MULTIPLE TIME-TO-EVENT DATA USING
PENALIZED B-SPLINES”**

Modeling Multiple Time-to-Event Data Using Penalized B-splines

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Abstract

Penalized B-splines have been applied to time-to-event data, providing an extension of the proportional hazards model for a single outcome (Gray, 1994). We use this technique to extend the marginal models of Wei, Lin and Weissfeld (1989). This allows for greater flexibility in modeling the margins and makes formal development of inferential procedures possible. Applications to data from the NSABP-BCPT on the effectiveness of the drug Tamoxifen as a prevention tool against breast cancer will be discussed in detail. Results from extensive simulation studies on the small sample properties of the asymptotic tests will also be presented.

KEY WORDS: Survival analysis; Smoothing; Ridge regression; Additive models; Splines; Proportional hazards.

1 Introduction

The advent of promising drugs like tamoxifen in the treatment and/or prevention of breast cancer has ignited both hope and controversy in the scientific world and the general public. The controversy revolves around the issue of whether the benefits of the drug offset its known adverse side effects. One of the main studies that has been conducted to study the effectiveness of tamoxifen as a preventive agent for breast cancer is the Breast Cancer Prevention trial, hereafter referred to as BCPT (Fisher et al, 1998). It has been shown that tamoxifen, when used for at least 5 years, was effective in prolonging disease free survival and in reducing the rate of recurrences of second primary tumors in contralateral breast and ipsilateral breast tumor. It has also been shown that tamoxifen reduces the risk of invasive breast cancer in women that are at elevated risk due to various factors. But, there is also evidence that use of tamoxifen is positively associated with invasive endometrial cancer, ischemic heart disease, transient ischemic attack, deep vein thrombosis and/or pulmonary embolism. In order to demonstrate the positive or negative effectiveness of tamoxifen, one needs to compare the advantages of the drug to its disadvantages in a simultaneous and comprehensive manner. To do this, one needs to be able to make simultaneous inference on several time-to-event outcomes and also be able to flexibly model the effect of risk and/or prognostic factors that have non-linear effects. Considerable progress has been made over the years in the development of models that handle multiple time-to-event outcome data and models that allow for flexible modeling of effects of prognostic factors for single time-to-event outcome. But, to date, flexible methods do not exist that allow for simultaneous inference of multiple time-to-event outcomes. In this paper, we develop new inferential methods that allow for simultaneous inference on flexible models for multiple time-to-event outcomes.

The proportional hazards model (Cox 1972) has received considerable attention as a

popular way of modeling, possibly censored, time-to-event data. In addition to the proportionality of the hazards, the model assumes that the effects of the predictors (risk factors) on the response follow a parametric (mostly linear) form. Recently, this assumption has been relaxed to allow for data-dependent, and possibly non-linear, covariate effects by exploiting the flexibility of nonparametric regression techniques (Hastie and Tibshirani 1990). Fully non-parametric proportional hazards models (O’Sullivan (1988) and Hastie and Tibshirani (1990)), while attractively flexible, usually suffer from heavy computational load and lack of formal inferential procedures. Gray (1994) used the concept of pseudo-smoothers, with emphasis on penalized B-splines, to develop formal inference for proportional hazards models. Penalized B-splines provide an elegant compromise between regression splines and smoothing splines.

Another issue in the analysis of time-to-event data is the modeling of multiple outcomes. This problem has received considerable attention in the statistical literature. For example, Wei, Lin and Weissfeld (1989) propose the use of marginal modeling. However, most available methods have not been extended to include flexible and possibly nonlinear effects of prognostic factors. On the other hand, many researchers have demonstrated that important prognostic factors (e.g. BMI) have a markedly non-linear effect on breast cancer survival and/or prognosis (Gray, 1994). These methods, however, are limited to single outcomes and do not lend themselves to simultaneous inference of several time-to-event outcomes.

In this article, we extend the marginal models of Wei, Lin and Weissfeld (1989) to allow modeling flexibility via the use of penalized B-splines in the style of Gray (1994). See also Hastie (1996) for a detailed discussion on a more general class of pseudo-smoothers. The remainder of the paper is organized as follows. In §2, we introduce the spline based proportional hazards model that fits a separate marginal model for each of several time-to-event outcomes. In §3, we discuss theoretical and computational details of the proposed

simultaneous inferential procedures. In §4, we present results from an extensive simulations study on the empirical size of the proposed tests in small sample settings. In §5, we present results from a detailed analysis of the BCPT data. The last section discusses the main findings of the paper and various modeling and model checking issues (including diagnostics measures) that extend the additive model to allow for testing the proportionality of hazards and multi-dimensional modeling.

2 The model

To model marginal distributions of multivariate time-to-event data, let us consider a flexible proportional hazards model for each of the G failure types. For the g^{th} type of failure of the i^{th} , $i = 1, \dots, n$, subject, the model can be written as

$$\lambda_{gi}(t) = \lambda_{g0}(t) \exp\{\sum_j f_{jg}(Z_{jgi})\}, \quad t \geq 0, \quad (1)$$

where $\lambda_{g0}(t)$ is an unspecified baseline hazard function and f_{jg} , $j = 1, \dots, p$, denotes the unspecified smooth functions. In the usual setup (Cox, 1972), one observes data of the form $(X_{gi}, Z_{gi}, \Delta_{gi})$, where $X_{gi} = \min(\tilde{X}_{gi}, C_{gi})$, C_{gi} is the censoring time, $Z_{gi}(t) = (Z_{1gi}(t), \dots, Z_{pgi}(t))^T$ and $\Delta_{gi} = 1$ if $X_{gi} = \tilde{X}_{gi}$ and 0 otherwise.

Model (1) is fully non-parametric and quite general. Note also that the fully linear model of Wei, Lin and Weissfeld (1989) forms a special case of (1) where $f_{jg}(Z_{jgi}) = \beta_{jg} Z_{jgi}$. For this fully linear model, the partial likelihood is given as

$$PL_g(\beta) = \prod_{i=1}^n \left(\frac{\exp\{\beta_{g(T)} Z_{gi}(X_{gi})\}}{\sum_{l \in \mathcal{R}_g(X_{gi})} \exp\{\beta_{g(T)} Z_{gl}(X_{gl})\}} \right)^{\Delta_{gi}}, \quad (2)$$

where $\beta_g = (\beta_{1g}, \dots, \beta_{pg})^T$ and $\mathcal{R}_g(t) = \{l : X_{gl} \geq t\}$ denotes the set of subjects at risk just prior to time t with respect to the g^{th} type of failure. The solution to $\partial \log PL_g(\beta_g)/\partial \beta_g = 0$,

$\hat{\beta}_g$, can be shown to be a consistent estimator of β_g provided that the fully linear model is correctly specified (Anderson and Gill, 1982).

In practical applications, the effects of most covariates are known to have some parametric form, while some of them are best modeled via non-parametric smoothers. For simplicity of discussion, we discuss most details for a model with p parametric and one additional non-parametric term. We first let

$$\lambda_{gi}(t) = \lambda_{g0}(t) \exp\left\{\sum_j \beta_{jg} Z_{jgi} + f_g(h_{gi})\right\}, \quad t \geq 0, \quad (3)$$

where $j = 1, \dots, p$. We propose to estimate $f_g(h_{gi})$ using the penalized regression spline approach, *i.e.*,

$$f_g(h_g) = \gamma_{1g} h_g + \sum_{q=2}^{m+3} \gamma_{qg} B_{qg}(h_g). \quad (4)$$

Note that, we have dropped the constant term since it is accounted for by the baseline hazard, and only $(m+2)$ of the B-spline basis functions are used for identifiability (De Boor, 1974). Following Gray (1994), let $\gamma_g = (\gamma_{g2}, \dots, \gamma_{g(m+3)})$ and $\eta_g = (\gamma_{1g}, \gamma_g)$. Then, a penalized partial likelihood that includes a penalty function to allow for smoother alternatives would be defined as

$$PL_g^p(\beta_g, \eta_g) = PL_g(\beta_g, \eta_g) - 1/2\lambda \int [f_g''(u)]^2 du. \quad (5)$$

Recognizing that the penalty function given above is quadratic in the parameter vector $\gamma = (\gamma_0, \gamma_1, \dots, \gamma_{m+3})$, one could rewrite (5) as

$$PL_g^p(\beta_g, \eta_g) = PL_g(\beta_g, \eta_g) - 1/2\lambda_g \eta_g^T \mathbf{K}_g \eta_g. \quad (6)$$

where \mathbf{K} is a positive definite matrix that is a function of the covariate h_g . Note that \mathbf{K} is an $(m+3) \times (m+3)$ matrix with the first row and column as zeros, since the linear function passes unpenalized.

The hypotheses of interest with respect to the smooth function are then $\gamma_g = \mathbf{0}$ and $\eta_g = \mathbf{0}$, representing the hypotheses of “no effect” and “linear effect” respectively.

A model that is more focused towards testing proportionality of hazards via the use of time-varying coefficients could be considered as follows:

$$\lambda_{gi}(t) = \lambda_{g0}(t) \exp\left\{\sum_j \beta_{jg} Z_{jgi} + \phi_g(t) h_{gi}\right\}, \quad t \geq 0. \quad (7)$$

It is straightforward to extend either of the above two models to allow for multiple, say M , non-parametric terms. In this case, η_g would be a bigger vector that augments contributions from the basis functions of the M terms. Here, $\eta_g = (\eta_{g1} : \dots : \eta_{gM})$ would be of dimension $M(m+3) \times 1$ and the penalty term would be the sum of the M penalty functions where each non-parametric term has its own smoothing parameter, and penalty matrix. One could then test for the “overall” effect or “linearity” of the individual non-parametric terms or for a combination of them.

3 Inference

While making inference on each of the margins is important, this could be done easily by using developments in Gray (1994). Our interest here is mainly in being able to conduct simultaneous inference on several time-to-event outcomes in models that have non-parametric smooth terms. Once the marginal distributions are modeled, then the methods described in Wei, Lin and Weissfeld (1989) can be extended to test for trends across parameter estimates and to combine estimates across margins to test for covariate effects of interest.

Let us consider the case where we have p parametric terms and one additional non-parametric term as given by (3). Then, for outcome g , the unpenalized part of equation (6),

suppressing the dependence of the regression parameters on X_{gi} , can be written as

$$PL_g(\boldsymbol{\beta}_g, \boldsymbol{\eta}_g) = \prod_{i=1}^n \left(\frac{\exp\{\sum_{j=1}^p Z_{gj}\beta_{gj} + h_g\gamma_1 + \sum_{q=2}^{m+3} B_{qg}(h_g)\gamma_{qg}\}}{\sum_{s \in \mathcal{R}_g(X_{gi})} \exp\{\sum_{j=1}^p Z_{gj}\beta_{gj} + h_g\gamma_1 + \sum_{q=2}^{m+3} B_{qg}(h_g)\gamma_{qg}\}} \right)^{\Delta_{gi}}, \quad (8)$$

where all components are as defined in §2, for the g^{th} type of failure. Let $\boldsymbol{\psi}_g = (\boldsymbol{\beta}_g, \boldsymbol{\eta}_g)$ and $P_g = (Z_{1g} : \dots : Z_{pg} : h_g : B_{2g}(h_g) : \dots : B_{m+3,g}(h_g))$ with P_{gr} denoting the r^{th} column vector, $r = 1, \dots, (m+p+3)$. Letting \hat{A}_g be the unpenalized information matrix for the g^{th} outcome as a function of $\boldsymbol{\psi}$, it can be shown that

$$\sqrt{n}(\hat{\boldsymbol{\psi}}_g - \boldsymbol{\psi}_{g(T)}) = n(A_g + \lambda_n \tilde{K})^{-1} n^{-1/2} U_g(\boldsymbol{\psi}_{g(T)}) + o_p(1)$$

where $U_g(\boldsymbol{\psi}_{g(T)})$ is the score vector and $\boldsymbol{\psi}_{g(T)}$ is the vector of true parameter values for the g^{th} outcome (Gray, 1994) and \tilde{K} is the expanded penalty matrix that augments rows and columns of zeros to \mathbf{K} to account for the unpenalized terms in the model. Then, it follows from the asymptotic normality of $U_g(\boldsymbol{\psi}_{g(T)})$ that $\sqrt{n}(\hat{\boldsymbol{\psi}}_g - \boldsymbol{\psi}_{g(T)})$ is asymptotically normal with mean $\mathbf{0}$ and variance given as the limit of nV_g where

$$V_g = (A_g + \lambda_n \tilde{K})^{-1} A_g (A_g + \lambda_n \tilde{K})^{-1}, \quad (9)$$

To develop the simultaneous inferential procedures for several outcomes, we first note that the $\boldsymbol{\psi}_g$'s across the G multiple outcomes are generally correlated. Then, analogous to developments in Wei, Lin and Weissfeld (1989), the asymptotic covariance matrix between $\sqrt{n}(\hat{\boldsymbol{\psi}}_g - \boldsymbol{\psi}_g)$ and $\sqrt{n}(\hat{\boldsymbol{\psi}}_v - \boldsymbol{\psi}_v)$ can be consistently estimated by

$$\hat{D}_{gv}(\hat{\boldsymbol{\psi}}_g, \hat{\boldsymbol{\psi}}_v) = \hat{V}_g(\hat{\boldsymbol{\psi}}_g) \hat{C}_{gv}(\hat{\boldsymbol{\psi}}_g, \hat{\boldsymbol{\psi}}_v) \hat{V}_v(\hat{\boldsymbol{\psi}}_v), \quad (10)$$

where $\hat{C}_{gv}(\hat{\boldsymbol{\psi}}_g, \hat{\boldsymbol{\psi}}_v) = n^{-1} \sum_{i=1}^n W_{gi}(\hat{\boldsymbol{\psi}}_g) W_{vi}(\hat{\boldsymbol{\psi}}_v)^T$, and W_{gi} and W_{vi} are defined in terms of the unpenalized score contributions as discussed in §4.1. below. Based on these results from §4.1, the covariance matrix of $(\hat{\boldsymbol{\psi}}_1, \dots, \hat{\boldsymbol{\psi}}_G)$ can be consistently estimated by

$$\hat{Q} = n^{-1} \begin{pmatrix} \hat{D}_{11}(\hat{\psi}_1, \hat{\psi}_1) & \dots & \hat{D}_{1G}(\hat{\psi}_1, \hat{\psi}_G) \\ \vdots & \ddots & \vdots \\ \hat{D}_{G1}(\hat{\psi}_G, \hat{\psi}_1) & \dots & \hat{D}_{GG}(\hat{\psi}_G, \hat{\psi}_G) \end{pmatrix}. \quad (11)$$

3.1 Calculation of W_g

The robust variance estimator introduced by Wei, Lin and Weissfeld (1989) for inference across margins uses a plug-in estimator for covariances between the scores of the g^{th} and v^{th} margins.

For the g^{th} type of failure, let

$$N_{gi}(t) = I(X_{gi} \leq t, \Delta_{gi} = 1) ,$$

$$Y_{gi}(t) = I(X_{gi} \geq t)$$

and

$$M_{gi}(t) = N_{gi}(t) - \int_0^t Y_{gi}(u) \lambda_{gi}(u) du ,$$

where $I(\cdot)$ denotes the indicator function. Then, it is straightforward to show that the penalized score function has the form

$$U_g^{(p)}(\psi_g) = U_g(\psi_g) - \lambda_g \tilde{K}_g \psi_g$$

where

$$\begin{aligned} U_g(\psi_g) &= \sum_{i=1}^n \int_0^t P_{gi}(u) dM_{gi}(u) \\ &\quad - \int_0^t \frac{\sum_{i=1}^n Y_{gi}(u) P_{gi}(u) \exp\{\psi_g^T P_{gi}(u)\}}{\sum_{i=1}^n Y_{gi}(u) \exp\{\psi_g^T P_{gi}(u)\}} d\bar{M}_g(u) \end{aligned} \quad (12)$$

where $\tilde{M}_g(u) = \sum_{i=1}^n M_{gi}(u)$. Based on arguments that are parallel to those in Wei, Lin and Weissfeld (1989), the asymptotic covariance matrix between $\sqrt{n}(\hat{\psi}_g - \psi_g)$ and $\sqrt{n}(\hat{\psi}_v - \psi_v)$ is given by

$$\hat{D}_{gv}(\hat{\psi}_g, \hat{\psi}_v) = \hat{V}_g(\hat{\psi}_g) E\{w_{g1}(\hat{\psi}_g) w_{v1}(\hat{\psi}_v)^T\} \hat{V}_v(\hat{\psi}_v) ,$$

where

$$w_{gj} = \int_0^i n f\{P_{gi}(t) - s_g^{(1)}(\psi_g; t)/s_g^{(0)}(\psi_g; t)\} dM_{gi}(t) ,$$

$$s_g^{(1)}(\psi_g; t) = E[Y_{gi}(t) P_{gi}(t) \exp\{\psi_g^T P_{gi}(t)\}] ,$$

and

$$s_g^{(0)}(\psi_g; t) = E[Y_{gi}(t) \exp\{\psi_g^T P_{gi}(t)\}] .$$

We then use a plug in estimate for $E\{w_{g1}(\hat{\psi}_g) w_{v1}(\hat{\psi}_v)^T\}$ which takes the form of \hat{C} as in (10). This estimator is the same as the estimator proposed in Wei, Lin and Weissfeld (1989), since the penalty converges to zero under the null hypothesis. For this reason, the penalty term is dropped in the plug in estimate for $E\{w_{g1}(\hat{\psi}_g) w_{v1}(\hat{\psi}_v)^T\}$. We define,

$$W_{gi}(\psi_g) = \Delta_{gi} \left\{ P_{gi}(X_{gi}) - \frac{S_g^{(1)}(\psi_g; X_{gi})}{S_g^{(0)}(\psi_g; X_{gi})} \right\} - \sum_{l=1}^n \frac{\Delta_{gl} Y_{gi}(X_{gl}) \exp\{\psi_g^T P_{gi}(X_{gl})\}}{n S_g^{(0)}(\psi_g; X_{gl})}$$

$$\times \left\{ P_{gi}(X_{gl}) - \frac{S_g^{(1)}(\psi_g; X_{gl})}{S_g^{(0)}(\psi_g; X_{gl})} \right\} , \quad (13)$$

$$S_g^{(1)}(\psi; t) = n^{-1} \sum_{i=1}^n Y_{gi}(t) P_{gi}(t) \exp\{\psi_g^T P_{gi}(t)\} ,$$

and

$$S_g^{(0)}(\psi; t) = n^{-1} \sum_{i=1}^n Y_{gi}(t) \exp\{\psi_g^T P_{gi}(t)\} .$$

The above asymptotic results are based on the approach used in Wei, Lin and Weissfeld (1989). Note that \hat{Q} is constructed as a function of the information matrix, the penalty

matrix, the smoothing parameter and the individual elements of the unpenalized score vector, that is, a separate term is computed for each of the n observations. Note that, for the above approximation, the penalized versions of the likelihood and the score functions are used to compute the information matrix while the unpenalized score vector is used in the plug in estimator for the computation of W as given in (13). Note also that the penalty matrix \tilde{K}_g contributes to the penalized score and information matrix only for the last $(m + 2)$ components of ψ_g . Inferential procedures for the first p parametric terms are directly analogous to those outlined in Wei, Lin and Weissfeld (1989).

3.2 Testing statistical hypotheses

For the non-parametric term, one could conduct simultaneous inference on the “overall” effect and/or “linearity” of h across failure types. Let $\hat{\gamma}_g$ denote the components of $\hat{\psi}_g$ that correspond to the relevant components of the non-parametric term h_g . Let also $\hat{\Gamma}$ denote the relevant sub-matrix of \hat{Q} corresponding to $\hat{\gamma} = (\hat{\gamma}_1, \dots, \hat{\gamma}_G)$. Then, one could use the quadratic form $(\hat{\gamma}_1, \dots, \hat{\gamma}_G)\hat{\Gamma}^{-1}(\hat{\gamma}_1, \dots, \hat{\gamma}_G)^T$ to conduct a joint test on the null hypotheses given by $H_0 : \gamma_g = 0, g = 1, \dots, G$. Note that the tests for “overall” significance or “linearity” are done in the above setup by choosing the last $(m + 3)$ and $(m + 2)$ elements of ψ_g respectively. A testing procedure that is more in the spirit of Gray (1994) uses $(A_g + \lambda_n \tilde{K}_g)^{-1}$ and $(A_v + \lambda_n \tilde{K}_v)^{-1}$ in (11) instead of V_g and V_v respectively. Under the null hypothesis, the modified Wald test statistic would then have an asymptotic distribution of

$$\sum_{g=1}^G \sum_j \lambda_{gj} \phi_j^2$$

where the ϕ_j are independent standard normal random variables, and the λ_{gj} ’s are the eigenvalues of the matrix $\lim A_{\gamma\gamma|\psi}(A_{\gamma\gamma|\psi} + \lambda \tilde{K})^{-1}$, for the g^{th} outcome. The arguments that lead to this form are given in Gray (1994) for a single outcome. The extensions to

multiple margins are straightforward. Note that the use of penalized B-splines, as opposed to fully nonparametric smoothers such as smoothing splines, makes the computation of the λ_{gj} 's possible.

A linear contrast could be constructed to test a group of parameters (e.g. all parameters to a spline term on each margin) across outcomes. For example, one could test the hypothesis that $\gamma_1 = \dots = \gamma_G = \gamma$. One could then estimate the common γ by using a linear combination of the γ_g 's in a way that takes the appropriate variances-covariance matrix into account. Unlike the tests discussed in Wei, Lin and Weissfeld (1989), where one is concerned with a single parameter from each margin, spline terms usually involve multiple parameters and the multicollinearity among them should be taken into account in taking the linear combinations via the off-diagonal covariance terms. Trends in regression effects across margins could also be examined along the lines of Wei, Lin and Weissfeld (1989) via sequential multiple testing procedures as in Wei and Stram (1988).

3.3 Choice of smoothing parameters, degrees of freedom, and placement of knots

In the above setup, we assume that the amount of smoothing (*i.e.*, the value of the smoothing parameter) is fixed by the analyst via prior knowledge or through a grid search. It is also possible that one could develop automatic procedures for selecting the smoothing parameters by using criteria such as cross validation. While this could lead to optimal estimation of the functional forms, its implications for hypothesis testing are not obvious. Operationally, one specifies the degrees of freedom per a non-parametric term and the corresponding value of smoothing parameter is then calculated. As a general operating guide, we use a relatively small number of degrees of freedom (Gray, 1994). The number of the knots that determine the B-spline basis functions are generally set to be at least twice the number of the degrees

of freedom in order to avoid wild fluctuation in the smooth function estimates, and are usually set to be between 10 and 15, per outcome. We will discuss the potential effects of various choices of the number of knots in our simulation studies. In this paper, we follow Gray (1994) in putting the knots at locations that yield approximately equal numbers of failure observations between knots. The calculation of degrees of freedom is analogous to that given in Gray (1994) and Wei, Lin and Weissfeld (1989). For example, to test whether all parameters in a spline model are equivalent across G outcomes, we use $\sum_{g=1}^G df_g$, where

$$df_g = \text{trace}\{\lim A_{\gamma\gamma|\psi}^{(g)} (A_{\gamma\gamma|\psi}^{(g)} + \lambda_g \tilde{K}_g)^{-1}\} .$$

4 Simulation Study

Extensive simulation studies were conducted to examine the performances of the proposed procedures for conducting simultaneous inference on several time-to-event outcomes. We focused on the bivariate case, where two time-to-event outcomes are considered under various levels of dependence. To generate data, the family of bivariate exponential distributions of Gumbel (1960) was used. Consider two marginal distributions, say F_1 and F_2 , from the univariate exponential with hazard rates given by $\exp(\beta_1 Z)$ and $\exp(\beta_2 Z)$, respectively. Then, the distribution function of the bivariate exponential distribution is of the form

$$F(x_1, x_2) = F_1(x_1)F_2(x_2)[1 + \theta\{1 - F_1(x_1)\}\{1 - F_2(x_2)\}] .$$

The quantity $\theta/4$ measures the correlation between the two event times, where $-1 \leq \theta \leq 1$. In the above models, Z denotes any vector of covariates that may include binary indicators, or covariate effects that assume various functional forms.

In the simulations that test for overall significance, we set the covariate values in the two margins to be equal. Censoring indicators were generated independently using uniform

distributions gauged to depict various percentages of censoring (30%, 50%). Empirical sizes of the spline based tests, based on 2000 runs were examined under various specifications of sample sizes ($n = 200, 300, 400$), degrees of freedom ($df = 3, 5$), number of knots (10,15,20) and levels of dependence between the margins ($\alpha = 0.5, 1.0$). Note that the degree of correlation between the two outcomes is given by $\alpha/4$ and $\alpha = 1$ the maximum correlation allowed by the bivariate model of Gumbel (1960).

Table 1 gives results from simulation with low levels of dependence ($\alpha = 0.5$) between the outcomes. The results indicate that the empirical size is reasonably close to the corresponding nominal values only when the sample size is at least 200 per margin. Based on these simulation results and similar observations in Gray (1994), it would be advisable to use a smoother that has relatively small number of degrees of freedom, with number of knots not exceeding 15 for most practical applications.

(Table 1 around here)

Table 2 gives results from the simulation with high levels of dependence ($\alpha = 1.0$) between the outcomes. Here, due to the added level of dependence between the margins, the empirical sizes for $n = 200$ was still unacceptably high (results not shown). But, the empirical sizes for $n = 300, 400$ give reasonable results.

(Table 2 around here)

5 Example: The NSABP-BCPT Data

As an illustration of the proposed methods, we present results from a detailed analysis of data from the Breast Cancer Prevention Trial, hereafter referred to as BCPT, (Fisher et al, 1996). The BCPT was initiated in 1992 enrolling 13388 women that were at increased risk

for breast cancer due to their relatively old age (≥ 60 years of age), relatively high 5-year predicted risk for breast cancer (a risk of at least 1.66% for those 35-59 years of age) and history of lobular carcinoma *in situ*. Subjects were then randomly assigned to placebo or treatment groups (6707 subjects into a placebo group and 6681 subjects receiving 20mg/day of tamoxifen for up to 5 years). The main aim was to examine the effectiveness of tamoxifen in preventing the possible occurrences of invasive breast cancer in high-risk women. Data was also collected on other outcomes (some of them unwanted adverse side effects) such as invasive endometrial cancer, ischemic heart disease, transient ischemic attack, deep vein thrombosis and pulmonary embolism.

Analysis of data from the BCPT has shown (Fisher et al, 1998) that there was a 49% reduction in the risk of invasive breast cancer in those high risk women that received tamoxifen treatment (of up to five years) compared to those that received placebo. But, the benefits of tamoxifen were tempered by adverse side effects that significantly increased the risk of endometrial cancer, deep vein thrombosis, pulmonary embolism and some other cardiac effects. In fact, the issue of whether the benefits of tamoxifen outweighs the potential risk was controversial enough that the NCI sponsored a workshop on the subject in July, 1998, leading a risk-benefit analysis as reported in Gail et al. (1999).

The results indicate that age and baseline predicted risks for breast cancer play a significant role in determining whether the benefits of tamoxifen outweigh the associated risks. In this paper, we use the new developed techniques to simultaneously analyze several outcomes in a way that allows for risks that may not be constant across factors such as age. We focus on the invasive breast cancer (IBC), ischemic heart disease (IHD) and endometrial cancer (ENDO) as our outcomes of interest. The primary covariates of interest were treatment (TRT, placebo vs. tamoxifen), age at time of entry (AGE, in years), 5 year breast cancer risk at time of entry (based on a multivariate logistic model of Gail et al. (1989)) (PR5YR),

lobular carcinoma in situ (LCIS) and atypical hyperplasia of the breast (ATYPH, history at entry). The two continuous covariates that could be modeled using the spline approach, in order to examine non-linearity in their effects, were age and the five-year breast cancer probability from the Gail model.

The results from the marginal models on each of the three outcomes are given in Table 3 and the corresponding smooth function estimates for AGE and PR5YR are given in Figures 1-3. The results from the marginal models indicate that use of tamoxifen is associated with reduced risk of invasive of breast cancer ($p < 0.01$), but it was also associated with significantly increased risk of endometrial cancer. The increased risk in ischemic heart disease appeared to be marginal and not statistically significant. Age of the subjects appeared to be positively associated only with ischemic heart disease, but this association appeared to be linear (Figure 2). On the other hand, the 5yr probability of breast cancer (as estimated from Gail model) was non-linearly associated with onset of invasive breast cancer. Here, the estimated curve (Figure 1) indicates an initial rise in risk up to 6-7 units for the risk score with a decline in risk starting at about 10 units. The test for non-linearity was marginally significant indicating that a simple linear term may not suffice to control for this variable.

(Table 3 around here)

(Figures 1-3 around here)

The results from two bivariate models that simultaneously model invasive breast cancer with ischemic heart disease and endometrial cancer are given in Table 4. The results indicate that the benefits of tamoxifen as a preventive agent significantly outweighs the side effect of increased risk in ischemic heart disease. On the other hand, the significant increased risk in endometrial cancer that is associated with the use of tamoxifen warrants a closer look since it appears to wash out its benefit of reducing the risk of breast cancer. However, these results should be interpreted cautiously due to the small number of events in the data set. The

results also indicate a strong linear effect of age in the bivariate model for invasive breast cancer and ischemic heart disease. Additionally, PR5YR appears to have a strong non-linear effect in both bivariate models, indicating that it should be modeled as a non-linear term.

(Table 4 around here)

6 Discussion

The methods proposed here have the advantage of being able to estimate a relatively realistic functional form for the covariate effects of interest, while enabling formal inference on the overall significance or adequacy of a certain parametric form (e.g. linearity) across several time-to-event outcomes. This is made possible through the use of penalized B-splines that are known to offer an attractive compromise between fully non-parametric regression smoothers such as smoothing splines and flexible, but inherently parametric, techniques such as regression splines (Hastie and Tibshirani (1990), Gray (1994)).

In this paper, we have introduced a method for conducting simultaneous inference across several outcomes by extending the methods of Gray (1994) and Wei, Lin and Weissfeld (1989). The results from the analysis of the breast cancer data demonstrate its immediate usefulness in health related research. The simulated studies demonstrate that the asymptotic inferential procedures are reliable in finite sample settings and also provide rough guidelines on how to select realistic values for the degrees of freedom (hence smoothing parameters) and number and location of knots.

There are many open areas of research that would extend the methods in this paper, some of which are currently active areas of research for our group. Some of the most important areas of research include dealing with proportionality of hazards, diagnostic measures in the

multivariate setting, testing for trends in some parametric but monotonic subclass of the general spline approach (linearity has been explored here) and a more in depth examination of the issue of proportionality of hazards. A more general class of models that is based on the notion of pseudosplines as in Hastie (1996) is currently being developed by our group and results will be reported elsewhere. In this class of models, examination of adequacy of increasingly complex forms of polynomials would be natural due to the general structure of orthogonal-polynomial based pseudosplines, as opposed to the penalized B-splines discussed in this paper.

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FIGURE LEGENDS:

1. Spline based estimates of the log hazard ratio for breast cancer as functions of age and five year probability of breast cancer
2. Spline based estimates of the log hazard ratio for ischemic heart disease as functions of age and five year probability of breast cancer
3. Spline based estimates of the log hazard ratio for endometrial cancer as functions of age and five year probability of breast cancer

Table 1: Empirical sizes of robust inference on marginally correlated ($\alpha = 0.5$) bivariate time-to-event outcomes

Censoring Prob.	Deg. of freedom	Number of knots	$n = 200$			$n = 300$		
			Nominal level			Nominal level		
			0.01	0.05	0.10	0.01	0.05	0.10
0.3	3	10	0.012	0.038	0.069	0.018	0.055	0.092
		15	0.022	0.070	0.121	0.029	0.079	0.130
		20	0.047	0.112	0.167	0.035	0.084	0.134
	5	10	0.030	0.068	0.114	0.022	0.071	0.121
		15	0.052	0.129	0.184	0.027	0.089	0.146
		20	0.103	0.200	0.270	0.051	0.137	0.206
0.5	3	10	0.013	0.051	0.096	0.013	0.051	0.089
		15	0.032	0.098	0.151	0.023	0.073	0.130
		20	0.074	0.163	0.238	0.041	0.120	0.185
	5	10	0.016	0.042	0.081	0.008	0.031	0.061
		15	0.029	0.080	0.124	0.015	0.046	0.083
		20	0.068	0.152	0.216	0.035	0.078	0.123

Table 2: Empirical sizes of robust inference on moderatelyy correlated ($\alpha = 1.0$) bivariate time-to-event outcomes

Censoring Prob.	Deg. of freedom	Number of knots	$n = 300$			$n = 400$		
			Nominal level			Nominal level		
			0.01	0.05	0.10	0.01	0.05	0.10
0.3	3	10	0.015	0.062	0.122	0.009	0.037	0.076
		15	0.033	0.092	0.156	0.012	0.051	0.086
		20	0.056	0.140	0.210	0.016	0.061	0.096
	5	10	0.028	0.085	0.144	0.012	0.045	0.081
		15	0.048	0.119	0.174	0.016	0.066	0.112
		20	0.078	0.166	0.237	0.024	0.073	0.131
0.5	3	10	0.022	0.085	0.172	0.004	0.025	0.051
		15	0.044	0.125	0.206	0.007	0.030	0.057
		20	0.066	0.171	0.263	0.010	0.049	0.086
	5	10	0.013	0.052	0.095	0.024	0.077	0.119
		15	0.023	0.078	0.136	0.029	0.086	0.156
		20	0.040	0.123	0.198	0.040	0.096	0.170

Table 3: Marginal Proportional Hazards Models on Breast Cancer, Ischemic Heart Disease and Endometrial Cancer

Outcome	Covariate	Estimate	Test Statistic	df	P-value
Invasive Breast Cancer	TRT	-0.69	28.08	1	<0.01
	LCIS	0.19	0.40	1	0.53
	AGE (overall)		2.89	4	0.61
	AGE (Linearity)		2.78	3	0.44
	PR5YR (overall)		17.26	4	<0.01
	PR5YR (Linearity)		6.94	3	0.05
Ischemic Heart Disease	TRT	0.13	0.59	1	0.44
	LCIS	-0.95	2.00	1	0.16
	AGE (overall)		73.3	3.99	<0.01
	AGE (Linearity)		3.54	3	0.30
	PR5YR (overall)		5.33	4	0.24
	PR5YR (Linearity)		2.96	3	0.40
Endometrial Cancer	TRT	0.88	8.23	1	<0.01
	LCIS	0.60	0.32	1	0.57
	AGE (overall)		4.32	3.99	0.36
	AGE (Linearity)		3.84	3	0.26
	PR5YR (overall)		5.19	4	0.25
	PR5YR (Linearity)		2.50	3	0.50

Table 4: Bivariate Proportional Hazards Models on Breast Cancer, Ischemic Heart Disease and Endometrial Cancer

Outcome	Covariate	Test Statistic	df	P-value
IBC and IHD	TRT	28.92	2	<0.01
	LCIS	2.24	1.97	0.32
	AGE (overall)	419.50	8	<0.01
	AGE (Linearity)	5.61	6	0.48
	PR5YR (overall)	24.80	8	<0.01
	PR5YR (Linearity)	10.93	6	0.07
IBC and ENDO	TRT	36.57	2	<0.01
	LCIS	0.44	2	0.62
	AGE (overall)	7.96	8	0.44
	AGE (Linearity)	7.29	6	0.27
	PR5YR (overall)	27.26	8	<0.01
	PR5YR (Linearity)	13.75	6	0.02

**APPENDIX 2. COPY OF “ON THE USE OF PSEUDOSPLINES IN MODELING
MULTIVARIATE SURVIVAL DATA: WITH APPLICATIONS TO THE NSABP-BCPT”**

On the use of pseudosplines in modeling multivariate survival data: with applications to the NASBP-BCPT

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Abstract

Pseudosmoothers have been applied to time-to-event data, providing an extension of the proportional hazards model for a single outcome (Gray, 1994). We use this technique to extend the marginal models of Wei, Lin and Weissfeld (1989), by also allowing for a more flexible class of models than that allowed in Gray (1994) along the lines of Hastie (1996). This allows for greater flexibility in modeling the margins and makes formal development of inferential procedures possible. This method is illustrated with an example using data from the NSABP trial and small sample properties on the size and power of the proposed tests are studied via simulated data.

KEY WORDS: Survival analysis; Smoothing; Ridge regression; Additive models; Splines.

1 INTRODUCTION

The proportional hazards model (Cox, 1972), a widely popular method of analyzing censored time-to-event data, has been recently extended to allow data-dependent, and possibly non-linear, covariate effects by exploiting the flexibility of nonparametric regression techniques (Hastie and Tibshirani 1990a). These extensions include the fully non-parametric methods of O'Sullivan (1988) and Hastie and Tibshirani (1990b). These models, while attractively flexible, usually suffer from heavy computational load and lack of formal inferential procedures. Gray (1994) used the concept of penalized B-splines to develop formal inference for proportional hazards models. Gray's model sets up an inherently parametric model by using B-spline basis functions (De Boor, 1974), but then penalizes them to allow for smoother alternatives. The end result is a model that is flexible enough to capture non-linearities in covariate effects, but also allows for formal inference due to its quasi-parametric structure.

To date, non-parametric regression models have not been developed for multivariate time-to-event outcomes. In this article, we extend the marginal models of Wei, Lin and Weissfeld (1989) to allow data-dependent estimation of possibly nonlinear effects of covariates. Our extensions are analogous to those of Gray (1994). But, we use the more general class of pseudosplines that has been proposed by Hastie (1996). Specifically, we would like to exploit the ordered complexity of orthogonal polynomial basis functions in

order to more closely examine trends in covariate effects. This also naturally lends itself to testing of the adequacy of polynomial functions as opposed to fully non-parametric modeling techniques. An important subclass of this rich class of models is the one that has time-varying coefficients of covariate effects, as it allows for a check on the proportionality of hazards (Zucker and Karr, 1990).

The methods in this article are motivated by applications to data from the Breast Cancer Prevention Trial (BCPT, Fisher et al. 1998). The main purpose of this study was to examine the usefulness of the drug tamoxifen as an agent that could potentially prevent invasive breast cancer. But, the study also collected data on some of the drug's known adverse effects, such as ischemic heart disease, endometrial cancer and pulmonary embolism. We wanted to analyze the multiple time-to-event data that arose from this study in a way that allows for simultaneous inference across the various outcomes, but also allowing for nonlinear effects of important risk factors.

The remainder of the paper is organized as follows. In §2, we present the new flexible marginal models for multiple-to-event data. In §3, we present results from extensive simulation studies to study the small sample properties of the proposed inferential procedures. §4 summarizes the results from applications of the proposed methodology to data from the NSABP breast cancer prevention trial (BCPT). In §5, we summarize the main results and give details on future directions for research. The details on the theoreti-

cal development and asymptotic properties of the inferential procedures are given in the Appendix (?).

2 Proposed Model

2.1 Pseudo-smoothers

To fix ideas, we first consider linear smoothers in the univariate framework. Letting $(x_1, y_1), \dots, (x_n, y_n)$ denote a set of n independent observations, a scatterplot smoother is said to be linear if, concentrating on the computations of the function only at the design points in $\mathbf{x} = (x_1, \dots, x_n)$, it can be written as a linear map $S : R^n \rightarrow R^n$ defined by $\hat{\mathbf{y}} = S\mathbf{y}$, where $\mathbf{y} = (y_1, \dots, y_n)$ is the response vector. Here S is referred to as a smoother matrix and is analogous to the hat matrix in linear regression. From this point onwards, our discussion focuses on the smoothing spline, even though the idea of pseudo-smoothers applies, in principle, to any linear smoother.

A cubic smoothing spline minimizes the penalized least squares criterion

$$\sum_{i=1}^n \{y_i - g(x_i)\}^2 + \alpha \int_{-\infty}^{+\infty} g''(z)^2 dz , \quad (1)$$

over a suitable Sobolev space W_2 of functions (Wahba 1990). An equivalent form of (1), that is based on the fitted vector $\mathbf{f} = (f_1, \dots, f_n)$ and a penalty matrix \mathbf{K} , is

$$(\mathbf{y} - \mathbf{f})^T(\mathbf{y} - \mathbf{f}) + \alpha \mathbf{f}^T \mathbf{K} \mathbf{f} . \quad (2)$$

The solution to (2), which is a natural cubic spline with knots at each distinct x_i , can be shown to be

$$\hat{\mathbf{f}} = (\mathbf{I} + \alpha \mathbf{K})^{-1} \mathbf{y} , \quad (3)$$

where $S = (\mathbf{I} + \alpha \mathbf{K})^{-1}$ is the smoother matrix.

Based on observations from the eigendecomposition of S , Hastie (1996) proposed a class of pseudo-smoothers that can be constructed from a given set of orthonormal basis functions (which should be ordered in complexity) and a penalty sequence.

Letting $\mathbf{p}(x)$ be a $M \times 1$ vector of orthonormal basis functions and $\theta_m, m = 1, \dots, M$ the penalties, a pseudospline could be defined as a minimizer of

$$Q_\lambda = (\mathbf{y} - P\beta)^T(\mathbf{y} - P\beta) + \alpha \beta^T D_\theta \beta , \quad (4)$$

where P is the matrix of evaluations of \mathbf{p} at the data, $D_\theta = \text{diag}(\theta_1, \dots, \theta_M)$ and α is a smoothing parameter. Then, the solution can be shown to be

$$\hat{\mathbf{f}} = P\beta = P(I + \alpha D_\theta)^{-1} P^T \mathbf{y} \quad (5)$$

provided that the bases are orthonormal with respect to the observed \mathbf{x} , leading to a ridge type regression setting. To approximate the smoother matrix of a smoothing spline, S , we supply a pseudobasis P in order to define a pseudo-eigendecomposition of S , *i.e.*,

$$\hat{S}(P) = PD_\psi P^T . \quad (6)$$

The pseudo-eigenvalues that form the $M \times M$ diagonal matrix $D_\psi = \text{diag}(\psi_m)$ can be obtained by $\psi_m = \mathbf{p}^T S \hat{\mathbf{p}}_m$, where $\hat{\mathbf{p}}_m$ is a result of smoothing \mathbf{p}_m in $O(n)$ computations. Hastie (1996) proposes a better approximation, $\hat{S}(P^*)$, obtained by considering the $M \times M$ eigendecomposition $P^T S P = V D_\psi^* V^T$ and defining $P^* = PV$.

To define a pseudosmooth, one needs to supply a set of basis functions and a penalty sequence. The rank (M) required to approximate S increases with the degrees of freedom ($\text{tr}(S)$). The use of penalized partial likelihood for a model that is defined via B-spline basis functions (Gray, 1994) is one example. There are many other choices for defining the pseudobasis functions (see Hastie 1996).

For weighted smoothers, the idea of approximation remains the same while the level of complications depends on the structure of the weight matrix. For a diagonal weight matrix with positive elements, the algorithm remains the same following premultiplication of the \mathbf{x} and \mathbf{y} values by square root of the respective weights. For iteratively weighted additive models (where the

weights change at each iteration), we use the above algorithm to approximate the penalty matrix K (instead of S itself). We will discuss the mechanism for doing this (and reasons for it) in the next section.

2.2 The model

To model marginal distributions of multivariate time-to-event data, let us consider a flexible proportional hazards model for each of the G failure types. For the g^{th} type of failure of the i^{th} , $i = 1, \dots, n$, subject, the model can be written as

$$\lambda_{gi}(t) = \lambda_{g0}(t) \exp\left\{\sum_j f_{jg}(Z_{jgi})\right\}, \quad t \geq 0, \quad (7)$$

where $\lambda_{g0}(t)$ is an unspecified baseline hazard function and f_{jg} , $j = 1, \dots, p$, denoting unspecified smooth functions. In the usual setup (Cox, 1972), one observes data of the form $(X_{gi}, Z_{gi}, \Delta_{gi})$, where $X_{gi} = \min(\tilde{X}_{gi}, C_{gi})$, C_{gi} is the censoring time, $Z_{gi}(t) = (Z_{1gi}(t), \dots, Z_{pgi}(t))^T$ and $\Delta_{gi} = 1$ if $X_{gi} = \tilde{X}_{gi}$ and 0 otherwise. Then, the partial likelihood for the fully linear model, where $f_{jg}(Z_{jgi}) = \beta_{jg} Z_{jgi}$, is given as

$$PL_g(\beta) = \prod_{i=1}^n \left(\frac{\exp\{\beta_g^T Z_{gi}(X_{gi})\}}{\sum_{l \in \mathcal{R}_g(X_{gi})} \exp\{\beta_g^T Z_{gl}(X_{gl})\}} \right)^{\Delta_{gi}}, \quad (8)$$

where $\beta_g = (\beta_{1g}, \dots, \beta_{pg})^T$ and $\mathcal{R}_g(t) = \{l : X_{gl} \geq t\}$ denotes the set of subjects at risk just prior to time t with respect to the g^{th} type of failure.

The solution to $\partial \log PL_g(\beta_g)/\partial \beta_g = 0$, $\hat{\beta}_g$, can be shown to be a consistent estimator of β_g provided that the fully linear model is correctly specified (REF).

Then, a penalized partial likelihood, letting $\eta_g = \sum_j f_{jg}(Z_{jgi})$, would be given as

$$j_g(\eta_g) = PL_g(\eta_g) - 1/2 \sum_{i=1}^p \alpha_{gi} \int f_g''(s)^2 ds . \quad (9)$$

In terms of the fitted vectors, this can be given as

$$j_g(\eta_g) = PL_g(\eta_g) - 1/2 \sum_{i=1}^p \alpha_{gi} \mathbf{f}_{gi}^T \mathbf{K}_{gi} \mathbf{f}_{gi} . \quad (10)$$

The Newton-Raphson algorithm for maximizing $j_g(\eta_g)$ over $\mathbf{f}_{g1}, \dots, \mathbf{f}_{gp}$, letting $\mathbf{u}_g = dl/d\eta_g$ and $\mathbf{A}_g = -d^2l/d\eta_g \eta_g^T$ denote the unpenalized score vector and information matrix respectively, in order to get updates of $\mathbf{f}_{1g}^{(0)}, \dots, \mathbf{f}_{1g}^{(0)}$ to $\mathbf{f}_{1g}^{(1)}, \dots, \mathbf{f}_{1g}^{(1)}$ is (suppressing the subscript g for now)

$$\begin{pmatrix} \mathbf{A} + \alpha_1 \mathbf{K}_1 & \mathbf{A} & \dots & \mathbf{A} \\ \mathbf{A} & \mathbf{A} + \alpha_2 \mathbf{K}_2 & \dots & \mathbf{A} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{A} & \mathbf{A} & \dots & \mathbf{A} + \alpha_p \mathbf{K}_p \end{pmatrix} \begin{pmatrix} \mathbf{f}_1^{(1)} - \mathbf{f}_1^{(0)} \\ \mathbf{f}_2^{(1)} - \mathbf{f}_2^{(0)} \\ \vdots \\ \mathbf{f}_p^{(1)} - \mathbf{f}_p^{(0)} \end{pmatrix} = \begin{pmatrix} u - \alpha_1 \mathbf{K}_1 \mathbf{f}_1^{(0)} \\ u - \alpha_2 \mathbf{K}_2 \mathbf{f}_2^{(0)} \\ \vdots \\ u - \alpha_p \mathbf{K}_p \mathbf{f}_p^{(0)} \end{pmatrix} . \quad (11)$$

Here, \mathbf{K}_j represents a quadratic penalty matrix for the j^{th} predictor. From (11), one gets the $np \times np$ system of estimating equations

$$\begin{pmatrix} \mathbf{I} & \mathbf{S}_1 & \mathbf{S}_1 & \dots & \mathbf{S}_1 \\ \mathbf{S}_2 & \mathbf{I} & \mathbf{S}_2 & \dots & \mathbf{S}_2 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \mathbf{S}_p & \mathbf{S}_p & \mathbf{S}_p & \dots & \mathbf{I} \end{pmatrix} \begin{pmatrix} \mathbf{f}_1^{(1)} \\ \mathbf{f}_2^{(1)} \\ \vdots \\ \mathbf{f}_p^{(1)} \end{pmatrix} = \begin{pmatrix} \mathbf{S}_1 \mathbf{z}^{(0)} \\ \mathbf{S}_2 \mathbf{z}^{(0)} \\ \vdots \\ \mathbf{S}_p \mathbf{z}^{(0)} \end{pmatrix} , \quad (12)$$

where $\mathbf{z}_g^{(0)} = \boldsymbol{\eta}_g^{(0)} + \mathbf{A}_g^{-1(0)} \mathbf{u}_g^{(0)}$ and $\mathbf{S}_{jg} = (\mathbf{A}_g + \alpha_{jg} \mathbf{K}_{jg})^{-1} \mathbf{A}_g$ for the g^{th} type of failure.

In this paper, we use the pseudospline technology of Hastie (1996) to approximate the fully non-parametric smoothing splines. This allows us to make the computation manageable. More importantly, it allows us to develop formal techniques for statistical inference.

The backfitting algorithm that follows from (11) requires an expensive operation ($O((n)^3)$ operations) in order to update each smooth function in (9). For large data sets with many predictor variables, this may not be practical. We propose to approximate each smoother matrix S_{jg} , $j = 1, \dots, p$, by a rank M_{jg} approximation.

Following Hastie (1996), we choose to approximate the quadratic penalty matrices K_{jg} (instead of the smoother matrices S_{jg} 's) for each nonparametric term in (9). This approach has two main advantages. First, the approximation of the K_{jg} 's can be done in the unweighted setting to be followed by a weighted generalized ridge regression in order to accommodate the weights. Second, there is no need to make fresh approximations at each iterative step of the outer loop in the local scoring algorithm. From our discussions in §2.1 and details in Hastie (1996), it is easy to show that the steps that we outlined for approximating the S_{jg} 's are still applicable in approximating the K_{jg} 's. The resulting orthogonal basis matrices \mathbf{P}_{jg} 's for the S_{jg} 's are also the eigenvectors of the K_{jg} 's and the corresponding eigenvalues for K_{jg} 's could

be obtained by using $\psi_{jg} = (1/(1 + \alpha_{jg}\theta_{jg}))$, where ψ_{jg} and θ_{jg} are $M_{jg} \times 1$ vectors of penalties corresponding to S_{jg} and K_{jg} respectively.

Now, let us replace each of the penalty matrices, K_{jg} , in (11) by a rank M_{jg} approximation $\hat{K}_{jg} = P_{jg}D_{\theta_{jg}}P_{jg}^T$ and let $\mathbf{f}_{jg} = P_{jg}\beta_{jg}$. Then (suppressing the subscript g),

$$\begin{aligned} & \left(\begin{array}{cccc} \mathbf{A} + \alpha_1 P_1 D_{\theta_1} P_1^T & \mathbf{A} & \dots & \mathbf{A} \\ \mathbf{A} & \mathbf{A} + \alpha_2 P_2 D_{\theta_2} P_2^T & \dots & \mathbf{A} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{A} & \mathbf{A} & \dots & \mathbf{A} + \alpha_p P_p D_{\theta_p} P_p^T \end{array} \right) \begin{pmatrix} P_1\beta_1^{(1)} - P_1\beta_1^{(0)} \\ P_2\beta_2^{(1)} - P_2\beta_2^{(0)} \\ \vdots \\ P_p\beta_p^{(1)} - P_p\beta_p^{(0)} \end{pmatrix} \\ &= \begin{pmatrix} u - \alpha_1 P_1 D_{\theta_1} P_1^T P_1 \beta_1^{(0)} \\ u - \alpha_2 P_2 D_{\theta_2} P_2^T P_2 \beta_2^{(0)} \\ \vdots \\ u - \alpha_p P_p D_{\theta_p} P_p^T P_p \beta_p^{(0)} \end{pmatrix}. \end{aligned} \quad (13)$$

Letting $P_g = (P_{1g} : P_{2g} : \dots : P_{pg})$, $D_{\theta_g} = Bdiag(D_{\theta_{jg}})$ and $\beta_g = (\beta_{1g}^T : \dots : \beta_{pg}^T)$ for the g^{th} type of failure, it is easy to show that (13) is equivalent to

$$P_g\beta_g^{(1)} = (\mathbf{A}_g + P_g D_{\theta_g} P_g^T)^{-1} \mathbf{A}_g \mathbf{z}_g^{(0)},$$

where $\mathbf{z}_g^{(0)} = \mathbf{A}_g^{-1(0)} \mathbf{u}_g^{(0)} + P_g \beta_g^{(0)}$ and hence

$$\beta_g^{(1)} = (P_g^T \mathbf{A}_g^{(0)} P_g + D_{\theta_g})^{-1} P_g^T \mathbf{A}_g^{(0)} \mathbf{z}_g^{(0)}. \quad (14)$$

For unweighted generalized ridge regression, a well known trick (Golub and Van Loan 1983) could reduce the problem to that of an ordinary least

squares. This is done by creating pseudo-design matrix and pseudo-response vectors of the form

$$P_g^a = \begin{pmatrix} P_g \\ D_{\theta_g}^{1/2} \end{pmatrix}$$

and

$$\mathbf{X}_g^* = \begin{pmatrix} \mathbf{X}_g \\ \mathbf{0} \end{pmatrix}.$$

This allows for efficient computation following a Q-R decomposition of P_g^a . See Hastie (1996, Appendix A.2) for details. When the weight matrix is diagonal, the same trick could be used after premultiplying P_g and \mathbf{X}_g by square-root of the corresponding weights. For the fully weighted proportional hazards framework, however, this trick is less appealing since it requires the formation of matrix square roots of the weight matrix and new Q-R decompositions at each iteration. The use of matrix square roots for \mathbf{A} is also bound to lead to ill-conditioning and numerical instability (Thisted 1988, §3.10.3).

Following Hastie and Tibshirani (1990), one could use a diagonal approximation of the weight matrix. This is shown to give the correct penalized maximum likelihood estimator and justified via the delta algorithm (Jorgensen, 1984). So, considerable saving in computation could be achieved by using the above outlined trick after proper pre-multiplications of the P_g 's and the \mathbf{X}_g 's by the square root of the diagonal weights.

2.3 Inference

Once the marginal distributions are modeled, then the methods described in Wei, Lin and Weissfeld (1989) can be extended to test for trends across parameter estimates and to combine estimates across margins to test for covariate effects of interest. Specifically, based on equation (14) the partial likelihood can be written as

$$PL_g(\beta_g) = \prod_{i=1}^n \left(\frac{\exp\{P_{g1}\beta_{g1}(X_{gi}) + \dots + P_{gp}\beta_{gp}(X_{gi})\}}{\sum_{l \in \mathcal{R}_g(X_{gi})} \exp\{P_{g1}\beta_{g1}(X_{gl}) + \dots + P_{gp}\beta_{gp}(X_{gl})\}} \right)^{\Delta_{gi}}, \quad (15)$$

where $\beta_{gr} = (\beta_{gr1}, \dots, \beta_{grm})^T$ and P_{gr} is the $m \times m$ orthogonal basis matrix, which is also a function of X , for the g^{th} type of failure and the r^{th} covariate.

Define

$$\begin{aligned} W_{gir}(\hat{\beta}_{gr}) &= \Delta_{gi} \left(\sum_{j=1}^m P_{gr(jr)}(X_{gi}) - \frac{S_g^{(1)}(\hat{\beta}_{gr}; X_{gi})}{S_g^{(0)}(\hat{\beta}_{gr}; X_{gi})} \right) - \sum_{j=1}^n \frac{\Delta_{gj} Y_{gi}(X_{gj}) \exp(\sum_{k=1}^p P_{gk}\beta_{gk}(X_{gj}))}{n S_g^{(0)}(\beta_{gr}; X_{gj})} \\ &\quad \left(\sum_{j=1}^m P_{gr(jr)}(X_{gj}) - \frac{S_g^{(1)}(\beta_{gr}; X_{gj})}{S_g^{(0)}(\beta_{gr}; X_{gj})} \right), \end{aligned} \quad (16)$$

where

$$S_g^{(1)}(\beta_r; t) = n^{-1} \sum_{i=1}^n Y_{gi}(t) \left(\sum_{j=1}^m P_{gr(jr)} \right) \exp \left(\sum_{s=1}^m P_{gs} \beta_{gs}(t) \right),$$

$$S_g^{(0)}(\beta_r; t) = n^{-1} \sum_{i=1}^n Y_{gi}(t) \exp \left(\sum_{s=1}^m P_{gs} \beta_{gs}(t) \right),$$

and $Y_{gi}(t) = I(X_{gi} \geq t)$. Then, the asymptotic covariance matrix between $\sqrt{n}(\hat{\beta}_{ui} - \beta_{ui})$ and $\sqrt{n}(\hat{\beta}_{vh} - \beta_{vh})$ can be consistently estimated by

$$\hat{D}_{uv}(\hat{\beta}_{ui}, \hat{\beta}_{vh}) = \hat{I}_u^{-1}(\hat{\beta}_{ui})\hat{\beta}_{uv}(\hat{\beta}_{ui}, \hat{\beta}_{vh})\hat{I}_v^{-1}(\hat{\beta}_{vh}) , \quad (17)$$

where $\hat{I}_v^{-1}(\hat{\beta}_{vi})$ is the inverse of the negative of the second derivative matrix of the partial likelihood based on (14) and $\hat{\beta}_{uv}(\hat{\beta}_{ui}, \hat{\beta}_{vh}) = n^{-1} \sum_{j=1}^n W_{uji}(\hat{\beta}_{ui})W_{vji}(\hat{\beta}_{vh})^T$, where W_{uji} and W_{vji} are defined in (16). Thus, the covariance matrix of $(\hat{\beta}_1, \dots, \hat{\beta}_p)$ can be consistently estimated by $\hat{Q} = n^{-1}[D_{11}]$.

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**APPENDIX 3. COPY OF “ESTIMATION OF THE SURVIVAL FUNCTION FOR GRAY’S
PIECEWISE-CONSTANT TIME-VARYING COEFFICIENTS MODEL”**

**ESTIMATION OF THE SURVIVAL FUNCTION
FOR GRAY'S PIECEWISE-CONSTANT
TIME-VARYING COEFFICIENTS MODEL**

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SUMMARY

Gray's extension of Cox's proportional hazards (PH) model for right-censored survival data allows for a departure from the PH assumption via introduction of time-varying regression coefficients (TVC). For this model estimation of the conditional hazard rate relies on the inclusion of penalized splines. Cubic penalized splines tend to be unstable in the right tail of the distribution and thus quadratic, linear and piecewise-constant penalized splines may be a favorable choice. We derive a survival function estimator for one important member of the class of TVC models - a piecewise-constant time-varying coefficients (PC-TVC) model. Using the first-order Taylor series approximation we also derive an estimate for the variance of the log- and log(-log)-transformed survival function, which in turn leads to estimated confidence limits on the corresponding scales of the survival function. Accuracy in estimating underlying survival times and survival quantiles is assessed for both Cox's and Gray's PC-TVC model using a simulation study featuring scenarios violating the PH assumption. Finally, an example of the estimated survival functions and the corresponding confidence limits derived from Cox's PH and Gray's PC-TVC model, respectively, is presented for a liver transplant data set.

1. INTRODUCTION

The Cox proportional hazards (PH) model has played a prominent role in both the statistical literature and for the analysis of right-censored survival data since its first introduction by

Cox [1] in 1972. It has been widely used for the analyses of biomedical data from both longitudinal studies and clinical trials, mainly due to its appealing mathematical simplicity, as well as its general availability through most statistical packages. While the Cox PH model is relatively simple to present, it relies on the assumption of proportionality which may not be met in all data sets. To address this issue, models that allow for non-proportionality of the conditional hazards through the introduction of penalized splines have been proposed. A family of models which can be used to model non-proportional data, the time-varying coefficient (TVC) models, have been considered by Gamerman and West [2], and Zucker and Karr [3]. A general treatment of the first order asymptotic analysis of the penalized likelihood is due to Cox and O'Sullivan [4]. Building on the work of Tsiatis [5], Andersen and Gill [6] and Gill [7], O'Sullivan [8] treated nonparametric estimation in the Cox model using an approach complementary to that of Zucker and Karr [3]. The methodology of Zucker and Karr was further developed by Gray [9, 10]. Time-varying coefficient models were also studied by Hastie and Tibshirani [11] and the use of regression splines in modeling the conditional hazard rate is discussed in Sleeper and Harrington [12] and Gray [9]. The use of time dependence in Cox's PH model was also investigated by Pettitt and Daud [13], Hess [14] and Verweij and van Houwelingen [15]. One of the more useful spline-based extensions of the Cox proportional hazards model is that proposed by Gray [9]. Gray's TVC extension of the Cox PH model employs products of the covariates of interest with the spline functions of time. This allows for a flexible approach to the modeling of covariate effects without necessarily adhering to the assumption of proportional hazards, which may not be satisfied. The most

appealing model within the framework of models proposed by Gray is the piecewise-constant TVC (Gray's PC-TVC) model since this model is similar to the original Cox PH model and retains much of the mathematical simplicity of the Cox model. The advantage of the PC-TVC models is their flexibility, since the proportional hazards assumption is only required for each of the time intervals between the successive knots (i.e. time points allowing for a change in the regression coefficients). Gray's PC-TVC model may therefore be viewed as a piecewise proportional hazards model for the conditional hazard rate. The estimated survival function is often of interest when fitting a survival model to data, since this serves as a useful summary of the estimated survival experience of a given population. Gray's work on TVC models has focused on estimation of the model coefficients, inference and residual analysis and to date, no estimator for the survival function has been presented. Andersen *et al.* [16] show that confidence limits for the survival function estimated from the Cox PH model are optimal when the estimates are based on a log-transformed or log(-log)-transformed scale for the survival curve. In this paper we present an estimator of the survival function under Gray's PC-TVC model. Estimation is based on the observation that between the successive knots, where the hazard regression coefficients are assumed to remain constant, the integration with respect to a differential of the cumulative hazard rate may proceed in a manner similar to that for the original Cox PH model. The estimated variance of the predicted survival function under Gray's PC-TVC model is derived for both the log- and log(-log)-transformed scale of the survival function and corresponding estimates of the confidence limits are presented.

2. ESTIMATED SURVIVAL FOR GRAY'S PC-TVC MODEL

Within the TVC family of models we assume that the hazard function can be modeled as follows:

$$d\Lambda(t|x) = d\Lambda_0(t) \exp\{x'\beta(t)\}, \quad (1)$$

where $\Lambda(\cdot)$ denotes the cumulative hazard function and $\Lambda_0(\cdot)$ denotes the cumulative baseline hazard. Here $\beta'(t) = (\beta_1(t), \beta_2(t), \dots, \beta_p(t))$, where $\beta_j(t) = \sum_k \theta_{jk} B_{jk}(t)$, $j = 1, \dots, p$ [9] are modeled with a full set of B-spline basis functions, $B_{jk}(t)$ [17]. Unlike Cox's proportional hazards model where the hazard regression coefficients, $\beta(t)$, in (1) are fixed, they are a function of time under Gray's PC-TVC model. Specifically, the coefficients are assumed to be constant only for values of $t \in [\tau_j, \tau_{j+1})$, $j = 0, \dots, q$. Here τ_j , $j = 1, \dots, q$, denote the internal knots, $\tau_0 = 0$, and $\tau_{q+1} = T$ represent the maximum observed (survival or censoring) time. Under Gray's PC-TVC model, the coefficients, $\beta(t)$, are therefore right-continuous step functions of time with jumps possibly occurring at the knots τ_j , $j = 1, \dots, q$. Estimation of the regression parameters in Gray's PC-TVC model proceeds by maximizing the penalized partial likelihood, which involves a partial likelihood term as in the Cox model, plus the following penalty term: $\frac{1}{2} \lambda_j \sum_{k=2}^{q+1} (\theta_{jk} - \theta_{j,k-1})^2$, where q is the number of internal knots for modeling the splines [9]. An essential component of the survival function estimate under Gray's PC-TVC model is based on the corresponding estimate of the cumulative baseline hazard. We extend Breslow's estimator [18] of the cumulative baseline hazard function to derive an estimator of the baseline hazard function for the TVC model. We assume that the

coefficients, β , in Breslow's formula can simply be replaced with their corresponding time-varying counterparts, $\beta(t)$. Consequently, under the TVC model (1) for the conditional hazard rate the estimated cumulative baseline hazard function is of the form:

$$\hat{\Lambda}_0(t) = \int_0^t \frac{1}{\sum_i Y_i(s) \exp\{z'_i \hat{\beta}(s)\}} \sum_{i=1}^n dN_i(s) , \quad (2)$$

where $Y_i(t)$ is an indicator function for the i -th patient's risk status at time t (i.e., $Y_i(t) = 1$ if the i -th patient is in the risk set at time t , and 0 otherwise).

For Gray's PC-TVC model the formula for the estimated survival function of a patient with p -variate covariate vector, \mathbf{z}_0 , will be:

$$\begin{aligned} \hat{S}(t|z_0) &= \exp\left\{-\int_0^t d\hat{\Lambda}(s|z_0)\right\} = \exp\left\{-\int_0^t \exp\{z'_0 \hat{\beta}(s)\} d\hat{\Lambda}_0(s)\right\} \\ &= \exp\left\{-\int_0^T I(s \leq t) \exp\{z'_0 \hat{\beta}(s)\} d\hat{\Lambda}_0(s)\right\}. \end{aligned} \quad (3)$$

On the log-transformed scale of the survival function we obtain:

$$\log \hat{S}(t|z_0) = -\int_0^T I(s \leq t) \exp\{z'_0 \hat{\beta}(s)\} d\hat{\Lambda}_0(s) = -\sum_{j=0}^q \exp\{z'_0 \hat{\beta}(\tau_j)\} \hat{\Lambda}_{0j}(t), \quad (4)$$

where

$$\hat{\Lambda}_{0j}(t) = \int_{[\tau_j, \tau_{j+1})} I(s \leq t) d\hat{\Lambda}_0(s) = \int_{[\tau_j, \tau_{j+1})} I(s \leq t) \frac{\sum_{i=1}^n dN_i(s)}{\sum_i Y_i(s) \exp\{z'_i \hat{\beta}(s)\}} \quad (5)$$

represents a contribution to the estimated (total) cumulative baseline hazard $\hat{\Lambda}_0(t)$ corresponding to an interval $[\tau_j, \tau_{j+1})$. Since $\beta(\tau_j)$ remains constant on $[\tau_j, \tau_{j+1})$, we will make use of the following notation: $\beta_j = \beta(\tau_j)$, where β_j is a vector of length p . Given a covariate vector \mathbf{z}_0 , we thus obtain an estimate of the survival function, $S(t|z_0)$, as follows:

$$\hat{S}(t|z_0) = \exp\left\{-\sum_{j=0}^q \exp\{z'_0 \hat{\beta}_j\} \hat{\Lambda}_{0j}(t)\right\}. \quad (6)$$

3. CONFIDENCE LIMITS BASED ON THE LOG-TRANSFORMATION

Based on (4), the formula for the variance of the log-transformed estimator of the survival function is as follows:

$$\begin{aligned} \text{Var}(\log \hat{S}(t|z_0)) &= \text{Cov}\left(-\sum_{j=0}^q \hat{\Lambda}_{0j}(t) \exp(z'_0 \hat{\beta}_j), -\sum_{j=0}^q \hat{\Lambda}_{0j}(t) \exp(z'_0 \hat{\beta}_j)\right) = \\ &= \sum_{k=0}^q \sum_{l=0}^q \text{Cov}(\hat{\Lambda}_{0k}(t) \exp(z'_0 \hat{\beta}_k), \hat{\Lambda}_{0l}(t) \exp(z'_0 \hat{\beta}_l)) . \end{aligned} \quad (7)$$

Note that (7) requires an estimator of the covariance which can be derived from a Taylor series approximation. We also define the following functions:

$$g(\hat{\beta}_j, t) = \hat{\Lambda}_{0j}(t) \exp(z'_0 \hat{\beta}_j), \quad j \in \{0, \dots, q\}. \quad (8)$$

The vector of the corresponding partial derivatives may be evaluated as follows:

$$\frac{\partial}{\partial \hat{\beta}_j} g(\hat{\beta}_j, t) = \exp(z'_0 \hat{\beta}_j) \left(z_0 \hat{\Lambda}_{0j}(t) + \frac{\partial}{\partial \hat{\beta}_j} (\hat{\Lambda}_{0j}(t)) \right).$$

Now, the first order Taylor series approximation of $g(\hat{\beta}_j)$ about the expected value of $\hat{\beta}_j$ (which we will denote by β_j) can be written as:

$$g(\hat{\beta}_j, t) \approx g(\beta_j, t) + \left[\frac{\partial}{\partial \hat{\beta}_j} (g(\hat{\beta}_j, t)) \Big|_{\hat{\beta}_j=\beta_j} \right]' (\hat{\beta}_j - \beta_j) . \quad (9)$$

The covariance terms in (7) can be approximated at time t using the Delta method as follows:

$$\text{Cov}\{g(\hat{\beta}_k, t), g(\hat{\beta}_l, t)\} \approx W_k(t)' \text{Cov}(\hat{\beta}_k, \hat{\beta}_l) W_l(t), \quad (10)$$

where

$$W_j(t) = \left[\exp(z'_0 \hat{\beta}_j) \left(z_0 \hat{\Lambda}_{0j}(t) + \frac{\partial}{\partial \hat{\beta}_j} \hat{\Lambda}_{0j}(t) \right) \right]_{|\hat{\beta}_j=\beta_j}, \quad j \in \{k, l\} \quad (11)$$

and

$$\frac{\partial}{\partial \hat{\beta}_j} \hat{\Lambda}_{0j}(t) = \int_{[\tau_j, \tau_{j+1})} I(s \leq t) \frac{-\sum_i Y_i(s) z_i \exp\{z_i' \hat{\beta}_j\}}{\left\{ \sum_i Y_i(s) \exp\{z_i' \hat{\beta}_j\} \right\}^2} \sum_{i=1}^n dN_i(s), \quad j \in \{k, l\} \quad (12)$$

is a p-variate vector of partial derivatives of $\hat{\Lambda}_{0j}(t)$.

At time t we also have:

$$z_0 \hat{\Lambda}_{0j}(t) = \int_{[\tau_j, \tau_{j+1})} I(s \leq t) \frac{\sum_i Y_i(s) z_0 \exp\{z_i' \hat{\beta}_j\}}{\left\{ \sum_i Y_i(s) \exp\{z_i' \hat{\beta}_j\} \right\}^2} \sum_{i=1}^n dN_i(s), \quad (13)$$

so that

$$W_j(t) = \left[\int_{[\tau_j, \tau_{j+1})} I(s \leq t) \frac{\sum_i Y_i(s) (z_0 - z_i) \exp\{(z_0 + z_i)' \hat{\beta}_j\}}{\left\{ \sum_i Y_i(s) \exp\{z_i' \hat{\beta}_j\} \right\}^2} \sum_{i=1}^n dN_i(s) \right]_{|\hat{\beta}_j = \beta_j}. \quad (14)$$

Consequently, the formula for the estimated variance of the predicted survival function will take the following form:

$$\hat{Var}(\log \hat{S}(t|z_0)) = \sum_{k=0}^q \sum_{l=0}^q W_k(t)' Cov(\hat{\beta}_k, \hat{\beta}_l) W_l(t). \quad (15)$$

Finally, the $100(1-\alpha)\%$ Confidence Limits for the survival function estimated under Gray's PC-TVC model are calculated as follows:

$$\exp\left(\log \hat{S}(t|z_0) \pm z_{1-\alpha/2} \sqrt{\hat{Var}(\log \hat{S}(t|z_0))}\right), \quad (16)$$

where $z_{1-\alpha/2}$ denotes an upper $\alpha/2$ -quantile of the standard normal distribution, $\hat{Var}(\log \hat{S}(t|z_0))$ is given by (15) and $\log \hat{S}(t|z_0)$ is estimated based on equation (4).

4. CONFIDENCE LIMITS BASED ON THE LOG-(LOG)-TRANSFORMATION

On the log(-log)-scale of the estimated survival function we obtain the following:

$$\log(-\log(\hat{S}(t|z_0))) = \log\left(\sum_{j=0}^q \hat{\Lambda}_{0j}(t) \exp(z'_0 \hat{\beta}_j)\right). \quad (17)$$

Let us denote the complete vector of time-varying coefficient estimates from Gray's PC-TVC model by $\hat{\beta} = (\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_q)$. Note that each component of the vector is itself a vector of length p (where p stands for the number of covariates being modeled by splines). Also, let $\frac{\partial}{\partial \hat{\beta}} \tilde{g}(\hat{\beta}) = \left(\frac{\partial}{\partial \hat{\beta}_0} \tilde{g}(\hat{\beta}), \frac{\partial}{\partial \hat{\beta}_1} \tilde{g}(\hat{\beta}), \dots, \frac{\partial}{\partial \hat{\beta}_q} \tilde{g}(\hat{\beta}) \right)$, where each of the q components of the vector of partial derivatives of $\tilde{g}(\hat{\beta})$ is itself a vector of length p. Using this notation we write at time t:

$$\tilde{g}(\hat{\beta}, t) = \log\left(\sum_{j=0}^q \hat{\Lambda}_{0j}(t) \exp(z'_0 \hat{\beta}_j)\right). \quad (18)$$

Thus the k-th component of the vector of partial derivatives (being itself a vector of length p) will be:

$$\left(\frac{\partial}{\partial \hat{\beta}} \tilde{g}(\hat{\beta}, t)\right)_k = \frac{\exp(z'_0 \hat{\beta}_k) \left(z_0 \hat{\Lambda}_{0k}(t) + \frac{\partial}{\partial \hat{\beta}_k} (\hat{\Lambda}_{0k}(t))\right)}{\sum_{j=0}^q \exp(z'_0 \hat{\beta}_j) \hat{\Lambda}_{0j}(t)}. \quad (19)$$

It follows from (14) that:

$$\left(\frac{\partial}{\partial \hat{\beta}} \tilde{g}(\hat{\beta}, t)\right)_k = \int_{[\tau_k, \tau_{k+1}]} I(s \leq t) \frac{\sum_i Y_i(s)(z_0 - z_i) \exp\{(z_0 + z_i)' \hat{\beta}_k\}}{\left\{ \sum_{j=0}^q \exp\{z'_0 \hat{\beta}_j\} \hat{\Lambda}_{0j}(t) \right\} \left\{ \sum_i Y_i(s) \exp\{z'_i \hat{\beta}_k\} \right\}^2} \sum_{i=1}^n dN_i(s). \quad (20)$$

Let us write:

$$\tilde{W}(t) = \left[\frac{\partial}{\partial \hat{\beta}} \tilde{g}(\hat{\beta}, t) \right]_{|\hat{\beta}=\beta}. \quad (21)$$

Using the first-order Taylor series approximation of the log(-log)-transformed survival function we can estimate the variance as follows:

$$\hat{Var}(\log(-\log(\hat{S}(t|z_0)))) \approx \tilde{W}(t)'Var(\hat{\beta})\tilde{W}(t), \quad (22)$$

where $Var(\hat{\beta})$ is the covariance matrix of the complete vector of time-varying coefficients with the partial derivatives in expression (22) evaluated as in (20). Consequently, the $100(1-\alpha)\%$ confidence limits for the survival function estimated under Gray's PC-TVC model based on the log(-log) transformation of the survival function will be given by:

$$\exp\left\{-\exp\left\{\log(-\log\hat{S}(t|z_0)) \mp z_{1-\alpha/2}\sqrt{\hat{Var}(\log(-\log\hat{S}(t|z_0)))}\right\}\right\}, \quad (23)$$

where $z_{1-\alpha/2}$ denotes an upper $\alpha/2$ -quantile of the standard normal distribution, $\log\hat{S}(t|z_0)$ is obtained from (4) and $\hat{Var}(\log(-\log\hat{S}(t|z_0)))$ is estimated using (22).

5. SIMULATION STUDIES

In order to assess the accuracy of both Cox's and Gray's survival estimators we designed two simulation studies allowing for comparison of the estimated survival quantiles and probabilities of survival obtained from Cox's and Gray's model with the true underlying values. We considered scenarios that violate the assumption of proportionality. In all instances throughout this article, Gray's PC-TVC model was fitted with 10 knots selected automatically so that approximately the same number of events was observed between the successive knots,

and 4 degrees of freedom that fully specify the choice of the corresponding value of the smoothing parameter.

We have generated survival data from the piecewise-exponential distribution with two time-points allowing for a change in the hazard at .3 and .8 years. For a set of survival probabilities $\{.99, .95, .90, .75, .50, .25, .10, .05, .01\}$, the corresponding time-points were estimated using both Cox's and Gray's models based on 1000 samples of size 150. Also, for a set of time points of 3, 7, 14 and 30 days and .5, 1, 1.5 and 3 years, estimates of the corresponding probabilities of survival were calculated from each of the models. For this simulation study, all of the data are complete. The results we obtained for censored data were very similar to those for complete data. The introduction of censoring, however, leaves some quantities related to the right tail of the distribution inestimable (e.g. time points corresponding to small survival probabilities).

In the first study, one third of each sample (associated with the first covariate being an indicator function for that group) was generated with hazards of $(1.5, 1, 2)$, the second third of the sample (associated with the second covariate) was generated with the hazards reversed (i.e. $(2, 1, 1.5)$), and the baseline hazards were all taken from $\text{Exp}(1)$. In the second study, hazards of $(2, 1, .5)$ were associated with the first covariate, those reversed $((.5, 1, 2))$ were associated with the second covariate and a constant hazard of 1 was again assumed for the baseline.

We wrote two simple Splus functions to compute the true survival quantiles and probabil-

ties for the piecewise-exponential distribution. Figure 1 (consisting of 4 panels numbered clockwise beginning from the top left quadrant) presents plots of the differences between the estimated and the true quantities (i.e. probabilities and survival quantiles respectively), as determined in both of the above studies. In both studies the survival curves were estimated at the covariate values (1,0) and (0,1) respectively, indicating a patient exhibiting the hazards specified by the first or second (i.e. reversed) set of hazards used in each example.

Panels 1 and 4 of Figure 1, based on 1000 samples, reveal that the differences between the true and estimated survival quantiles (times) were consistently smaller for Gray's model (denoted by circles in the plot). For this model the four corresponding trends in the hazard implied average departures from the true underlying quantiles of less than 20 days with the exception of the 1% quantile, for which the average departures ranged from 21 to 60 days. For the Cox model (denoted by triangles in the plot), however, departures from the true values greater than 50 days were observed for the 75, 50, 10, 5 and 1% survival quantiles. Panel 4 reveals that the estimates of the two smallest survival quantiles based on the Cox model were actually off by more than 1 year for both trends in the hazard. The magnitude of error observed was generally higher for the hazard rates of (2,1,.5) or reversed, than for those of (1.5,1,2) or reversed.

Similarly, panels 2 and 3 of Figure 1 illustrate the superior performance of Gray's PC-TVC model over that of the Cox model in terms of the accuracy of the estimated survival probabilities associated with several pre-determined time points. For 1000 samples simulated

with hazards (1.5,1,2) and (2,1,1.5) respectively, the probability estimates based on Gray's model were all within a distance of .01 from the true underlying values. For hazard rates of (2,1,.5) and (.5,1,2), estimates obtained from Gray's model exceeded the .01 distance in 3 of 18 cases with the maximum departure from the true value being .017 (associated with the time point of 6 months). Based on the Cox model, however, departures below .01 were observed in only 10 of 36 cases. In 16 of the 36 cases the magnitude of error associated with the Cox model exceeded the level of .025. The magnitude of error was again generally higher for the hazard rates (2,1,.5) or reversed, than for those of (1.5,1,2) or reversed.

The averaging effect of the Cox model is well documented in panels 2 and 3. Since the simulated hazard rates stabilized after .8 years, we observe that the departures from the true underlying values decreased dramatically after 1 year. As a result of the lack of flexibility on the part of Cox's model, however, this lead to subsequent departures in the opposite directions at the right tail of the distribution.

Results obtained from the simulation studies indicate that a high level of accuracy is maintained by the survival function estimates based on Gray's model, even in the tails of the distribution. Estimates obtained using Gray's model were generally close to the true values, while those derived from the Cox model occasionally showed large departures from the true underlying values. This resulted from a violation of the proportionality assumption in the data. The lack of precision in Cox's model was caused by the averaging of the time-varying effects, which is a built-in feature of Cox's model. In contrast, a high level of accuracy has

been maintained by Gray's survival estimator, even in the tails of the distribution.

6. UNOS DATA EXAMPLE

In this section we present a real data example comparing survival function estimators derived from Cox's and Gray's model, respectively. It features a dataset from the UNOS (United Network for Organ Sharing) database of cancer patients who underwent a liver transplant.

Here we estimate the graft survival for a subject whose covariate values are set to the median sample values. In graft survival analysis a failure is defined as an organ failure or a death of the recipient. We compare the best Cox and Gray models found for the data. The best models featured the following covariates (with corresponding sample median values listed in the parentheses): donor's anti CMV IGG result (dcmvgr, 1), indicator of whether the recipient had any prior transplant (priortx, 0), log-serum creatinine (lcreat, 0), log-total serum bilirubin (ltbili, 1.224), blood match indicator (abo.mtch, 1) and log-prothrombin time (lptp, 2.695). A summary of the modeling results may be found in Table 1. Covariates found to be significant under the best Cox model for the liver transplant graft survival of UNOS cancer patients were "lcreat", "ltbili", "dcmvgr" and "abo.mtch", with log-total serum bilirubin (ltbili) being identified as marginally non-proportional with regard to the effect on the hazard rate (p-value 0.0499). The best Gray's model included "lcreat", "lptp", "abo.mtch" and "priortx". Here the log-prothrombin time (lptp) was identified as having a highly non-proportional effect on the hazard rate (p-value 0.007).

Survival functions and 95% confidence limits estimated by the two models at the sample median covariate values are presented in Figure 2. Although the confidence bands for the two survival curves overlap (Gray's estimated survival function actually follows closely the upper confidence band estimated by the Cox's model), we can still observe a notable difference between the two survival estimates. The real data example of this section further illustrates the differences in survival estimates that might be obtained for data which does not follow the proportional hazards assumption.

7. CONCLUSIONS

Gray's piecewise-constant time-varying coefficients model for right-censored survival data is a flexible alternative to the Cox proportional hazards model in scenarios where the PH assumption may not be satisfied. The survival function estimator that we derived for this model provides a useful summary of the modeling results based on the patient's covariate values.

Simulation studies presented earlier have shown a lack of accuracy on the part of the Cox model with regard to estimating survival probabilities and predicting survival quantiles when the survival distribution does not satisfy the PH assumption.

Finally, based on Cox and Gray's model, respectively, a differing graft survival experience was demonstrated for a UNOS cancer patient after a liver transplant.

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We thank Dr. Robert J. Gray for making his own implementation of the modeling routine available to us. The survival function estimator has been written as an Splus function based on Gray's routine and is available from the authors upon request.

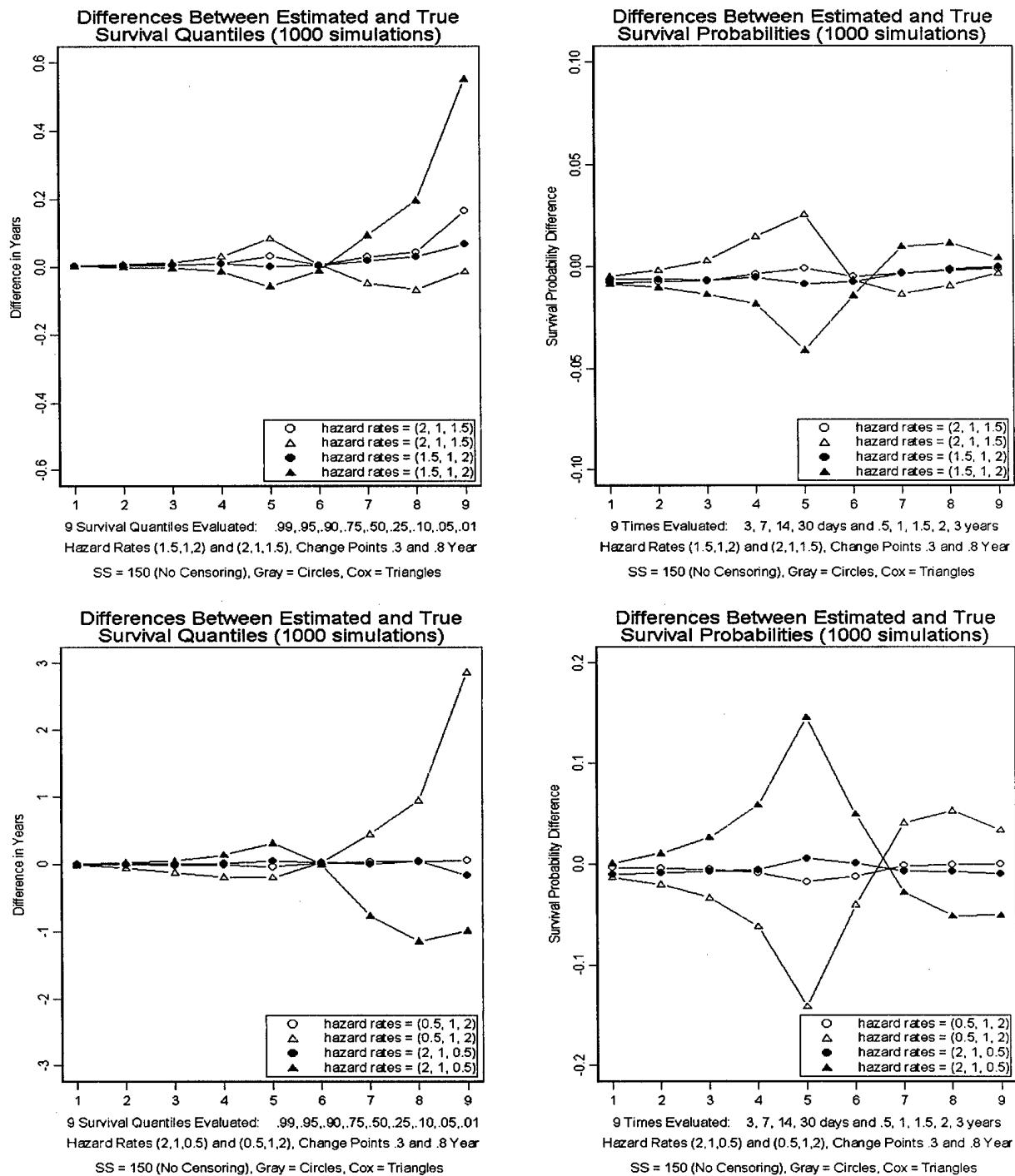


Figure 1: Simulation studies results summary

TABLE 1: Results Summary for UNOS Cancer Patients
 (502 observations with 278 failures)

Covariates	Cox's Model			Gray's Model		
	Coeff	p-value	n.prop.	Coeff (Range)	p-value	n.prop.
lcreat	.266	.014	.789	(.154:.555)	.001	.277
ltbili	.182	.001	.050	-	-	-
dcmvgr	.307	.011	.936	-	-	-
abo.mtch	-1.147	.049	.642	(-2.936:.205)	.007	.142
lptp	-	-	-	(-.244:1.688)	.000	.007
priortx	-	-	-	(-5.999:3.228)	.040	.769

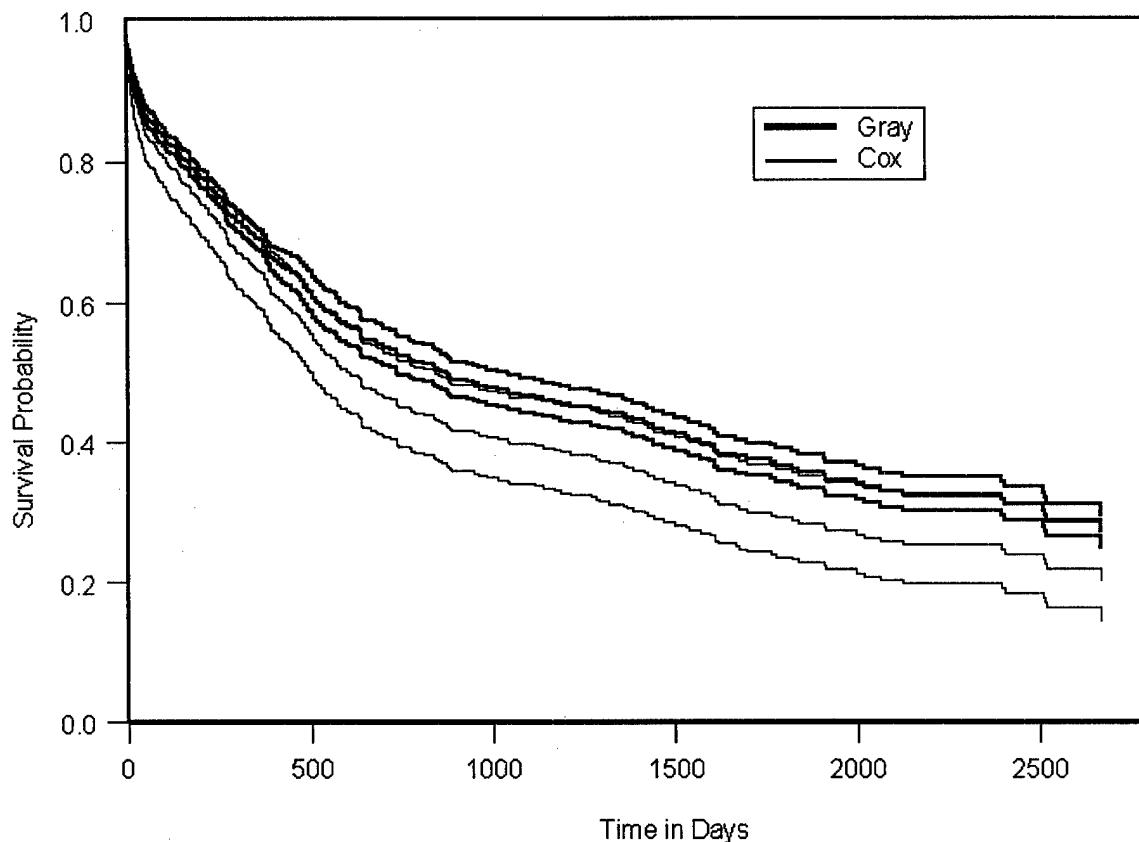


Figure 2: UNOS cancer patients post-liver transplant graft survival with 95% C.L.'s, Cox and Gray model results for a subject with median-valued covariates

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APPENDIX 4. COPY OF S-PLUS CODE FOR EXECUTING GRAY'S MULTIPLE OUTCOME SURVIVAL MODEL

```

#####
# file contains s code defining functions to fit WLW (1989) type
# flexible Cox models for multiple outcome data as in Berhane and
# Weissfeld (2000?). The functions are: mcox.spline, wlwcov,
# mcox.spline.test, spline.test cox.spline.int2
# cox.spline.plot cox.spline.mresid and cox.spline. [and to read
#jasa.data]
#
# This software comes with absolutely no guarantees. It was modified
# from the software by Robert Gray (1992), Harvard Univeristy.
# You have permission to
# use it for any noncommercial purpose, and to modify it as needed.
#
#
#####
#
# An S-plus function to fit a cox.spline model on each of the multiple
# margins and calculate the variance covariance matrix (as in Berhane
# and Weissfeld, 2000?) for simultaneous Wald type inference across
# the margins.
#
# Input variables are:
#
# gind      --- Outcome type indicator vector
# model.type --- as in cox.spline (only "a" [for additive] allowed)
# time       --- vector of failure times (sorted by outcome type)
# status      --- 1 if failure, 0 if censored (by outcome type)
# spline.cov --- list of outcome-specific matrices of covaraites
#                  to be modeled with splines
# linear.cov --- list of outcome-specific matrices of covaraites
#                  to be modeled as linear terms
# strata     --- stratification variable (numeric,
#                  character, or category) sorted by outcome type
#                  Default is all observations in one strata
# df         --- matrix of degrees of freedom for the spline fits
#                  one column per outcome type.
#
# Note: Ignored if smooth.param is specified.
# Default is 3 for nonlinearity (about 4 total) for each variable when
# model.type="a", 3 for each variable when model.type="t", and 8 for the
# tensor product term when model.type="i".
# If length of df is not of the appropriate length (1 for "i", the number
# of columns in spline.cov for "a" or "t"), a warning message is printed,
# and df replicated to the appropriate length.
#
# nknot      --- number of knots for spline variables (length 1 for
#                  "a" or "t", since the same value is used for all the
#                  covariates and outcome types, and length 2 for "i").
# spline.knot --- locations of knots for splines
#
# Note: Each set of knots must be an augmented
# knot sequence of length nknot+6, with something <= min for the first
# three entries, then the
# interior knots, then something >= the max for the next 3 entries. If
# "a" then a set of knots is needed for each column of spline.cov. If
# "i", 2 sets of knots are needed, and if "t" only one is needed. Default
# is for the program to choose knot locations based on equal numbers of
# observations between knots. (Default for now)
#

```

```

# smooth.opt --- smoothing parameter option.
# Note: If smooth.param is not specified, the smoothing parameters
# in the penalty function are chosen to give the specified degrees
# of freedom. If smooth.opt>0 smoothing parameters adjusted after every
# iteration, if <= 0 they are calculated after the first iteration only,
# and then held fixed. This can speed convergence, but the degrees of
# freedom at the final estimates may not match the specified df.
#
# smooth.param --- matrix of smoothing parameters.
# Note: If specified must have dimension equal to
# G (# outcome type) columns and raws equal to the number
# of columns in spline.cov ("a", "t") or 1 ("i").
# If specified, the smoothing parameters are kept fixed at these
# values throughout the iteration.
#
# Output variables are:
#
# m.out --- A list of results from marginal fits of
# of every outcome type.
#
# mbeta --- vector of regression parameter estimates
# sorted by outcome types.
#
# Qhat --- Variance-covariance matrix between outcome types
# for simultaneous inference.
#
# test.list --- summary of simultaneous inference across outcome
# types for overall significance and for non-linearity
#
mcoxmp.spline_function(gind, m.type, time, status, spline.cov, linear.cov,
  tstrata, dof, nknot, spline.knot, smooth.opt = 1, smooth.param,
  output.opt = "tests", nest, maxiter = 30, eps = 0.0001,
  rescale = T, ord = 0)
{
  g <- length(unique(gind))
  m.out <- vector("list", g)
  mbeta <- NULL
  w.time <- cbind(gind, time)
  w.status <- cbind(gind, status)
  if(!missing(spline.cov))
    wspline.cov <- cbind(gind, spline.cov)
  if(!missing(linear.cov))
    wlinear.cov <- cbind(gind, linear.cov)
  if(!missing(tstrata))
    w.strata <- cbind(gind, tstrata)
  for(i in 1:g) {
    t.time <- w.time[gind == i, -1]
    o.time <- order(t.time)
    i.time <- t.time[o.time]
    id <- rep(1:length(i.time), 1)
    o.id <- id[o.time]
    t.status <- w.status[gind == i, -1]
    i.status <- t.status[o.time]
    if(!missing(spline.cov)) {
      if(!missing(linear.cov)) {
        tspline.cov <- as.matrix(
          wspline.cov[gind == i, -1])
      }
    }
  }
}

```

```

        ispline.cov <- tspline.cov[o.time,]
        tlinear.cov <- as.matrix(
            wlinear.cov[gind == i, -1])
        ilinear.cov <- tlinear.cov[o.time,]
    }
    else {
        tspline.cov <- as.matrix(
            wspline.cov[gind == i, -1])
        ispline.cov <- tspline.cov[o.time,]
        #print(nrow(ispline.cov))
    }
}
else {
    tlinear.cov <- as.matrix(wlinear.cov[gind == i, -1])
    ilinear.cov <- tspline.cov[o.time,]
    #print(nrow(ilinear.cov))
}
t.strata <- w.strata[gind == i, -1]
i.strata <- t.strata[o.time]
#cat("ready for model\n")
if(!missing(smooth.param)) {
    if(!missing(spline.cov)) {
        if(missing(linear.cov)) {
            m.out[[i]] <- cox.spline(model.type=
                =m.type, time=i.time,
                status=i.status,
                spline.cov = ispline.cov,
                strata = i.strata,
                df = dof, nknot = nknot,
                smooth.opt=1, smooth.param=
                smooth.param[, g], output.opt
                = "tests", maxiter = 30, eps
                = 0.0001, rescale = T, ord = 0
            )
        }
        else {
            m.out[[i]] <-cox.spline(model.type=
                m.type, time = i.time,
                status=i.status,
                linear.cov = ilinear.cov,
                spline.cov = ispline.cov,
                strata = i.strata,
                df=dof, nknot=nknot,
                smooth.opt=1, smooth.param=
                smooth.param[, g],
                output.opt = "tests",
                maxiter = 30, eps=0.0001,
                rescale = T, ord = 0)
        }
    }
    else {
        m.out[[i]] <-cox.spline(model.type=m.type,
            time=i.time, i.status,
            linear.cov = ilinear.cov,
            strata = i.strata, df = dof,
            nknot = nknot, smooth.opt = 1,
            smooth.param = smooth.param[, g],

```

```

        output.opt = "tests",
        maxiter = 30, eps = 0.0001, rescale = T,
        ord = 0)
    }
}
else {
    if(!missing(spline.cov)) {
        if(!missing(linear.cov)) {
            #cat("spl.lin\n")
            m.out[[i]] <- cox.spline(model.type
                = m.type, time=i.time,
                status=i.status,
                linear.cov = ilinear.cov,
                spline.cov = ispline.cov,
                strata = i.strata,
                df = dof, nknot = nknot,
                smooth.opt = 1,
                output.opt = "tests", maxiter
                = 30, eps = 0.0001, rescale =
                T, ord = 0)
        }
        else {
            #cat("spl.lin.mis\n")
            m.out[[i]] <- cox.spline(
                model.type=m.type,
                time=i.time,
                status=i.status,
                spline.cov = ispline.cov,
                strata = i.strata,
                df = dof,
                nknot = nknot,
                smooth.opt = 1,
                output.opt = "tests",
                maxiter = 30, eps = 0.0001,
                rescale = T, ord = 0)
        }
    }
    else {
        #cat("spline missing\n")
        m.out[[i]] <-cox.spline(model.type=m.type,
            time=i.time,
            strata = i.status,
            linear.cov = ilinear.cov,
            strata = i.strata,df=dof,
            nknot=nknot, smooth.opt=1,
            output.opt = "tests", maxiter = 30,
            eps = 0.0001, rescale = T, ord = 0)
    }
}
mbeta <- c(mbeta, m.out[[i]]$coef)
tw = m.out[[i]]$w
m.out[[i]]$w = tw[order(o.id),]
twp = m.out[[i]]$wp
m.out[[i]]$wp = twp[order(o.id),]
}
gp <- length(mbeta)
nump <- gp/g

```

```

p <- nump
Qhat <- matrix(rep(0, gp * gp), ncol = gp)
for(i in 1:g) {
  for(j in 1:g) {
    i.ind <- (((i - 1) * p) + 1):(((i - 1) * p) + p)
    j.ind <- (((j - 1) * p) + 1):(((j - 1) * p) + p)
    Qhat[i.ind, j.ind] <- wlwcov.wt(m.out[[i]]$d2i,
                                         m.out[[j]]$d2i,
                                         m.out[[i]]$var,
                                         m.out[[j]]$var,
                                         m.out[[i]]$wp, m.out[[j]]$wp,
                                         m.out[[i]]$coef, m.out[[j]]$coef)$dij
  }
}
#cat('Qhat is:\n ')
mcox.out_mcox.spline.intm(m.out,Qhat)
return(Qhat,m.out, mbeta, g, mcox.out)
}

wlwcov.wt_function(pvi,pvj,uvj,wi,wj,betai,betaj){
# An S-plus auxilliary function to calculate the middle (meat) part of the
# variance-covariance for the multiple outcome data (called by
# mcox.spline)
#
# Input variables are:
#
#       vi, vj      --- Var-Cov matrices for ith and jth outcome types
#       wi, wj      --- W matrices for ith and jth outcome types as in
#                      Berhane and Weissfeld (2000?, Eqn. 10)
#
# Output variables are:
#
#       Phi      --- Middle part of the var-cov matrix between ith
#                      and jth outcome types
#       dij      --- Var-cov matrix for ith and jth outcome types
#
nobs _ nrow(na.omit(wi))
phi_t(wi) %*% wj
#phi_phi/nobs
uvi_solve(uvi)
uvj_solve(uvj)
uvi_(uvi+t(uvi))/2
uvj_(uvj+t(uvj))/2
vi_pvi%*%uvi%*%pvi
vj_pvj%*%uvj%*%pvj
dij_vj%*%phi%*%vj
#dij_dij/nobs
return(phi,dij)
}

mcox.spline.intm _ function(zlist,Qhat) {
# zlist is a list pf the lists created by cox.spline on several outcomes
# (margins) but with the same variables and knots (options "a" and "t" only).
if (any(zlist[[1]]$nvar != zlist[[2]]$nvar) |
  (length(zlist[[1]]$coef) != length(zlist[[2]]$coef)))
  stop("models must include the same variables")
}

```

```

if (zlist[[1]]$type == "a" & zlist[[2]]$type == "a") {
  nsplin _ zlist[[1]]$nvar[2]
  nknot _ length(zlist[[1]]$knot)/nsplin-6
  nfixed _ zlist[[1]]$nvar[1]
  if (nfixed>0) {
    g_length(zlist)
    #cat("g is: ")
    #print(g)
    out.fn _ vector("list",nfixed)
    out.fr _ vector("list",nfixed)
    #names(out.fn) _ names(z1$coef)[1:nfixed]
    names(out.fr) _ names(out.fn)
    for (i in 1:nfixed) {
      np1.t _ length(zlist[[1]]$coef)
      np.t_np1.t*g
      theta.f_NULL
      np.f_0
      for(j in 1:g){
        theta.fup_ zlist[[j]]$coef[i]
        theta.f_c(theta.f,theta.fup)
        np.f_np.f+1
      }
      v.f _ matrix(0,nrow=np.f,ncol=np.f)
      h.f _ v.f
      q.f_v.f
      npt_0
      for(j in 1:g){
        v.f[j,j]_zlist[[j]]$var[i,i]
        h.f[j,j]_zlist[[j]]$d2i[i,i]
        q.f[j,j]_Qhat[npt+i,npt+i]
        npt_npt+np1.t
      }
      ct.f _ diag(np.f)
      out.fn[[i]]_spline.test.naive(theta.f,h.f,v.f,q.f,ct.f)
      out.fr[[i]]_spline.test.rob(theta.f,h.f,v.f,q.f,ct.f)
    }
  }
  else {
    out.fn _ NULL
    out.fr_NULL
  }
}

if (nsplin>0) {
  g_length(zlist)
  #cat("g is: ")
  #print(g)
  out.an _ vector("list",nsplin)
  out.ln_vector("list",nsplin)
  out.ar _ vector("list",nsplin)
  out.lr_vector("list",nsplin)
  names(out.an) _ names(zlist[[1]]$test)
  names(out.ln) _ names(zlist[[1]]$test)
  names(out.ar) _ names(zlist[[1]]$test)
  names(out.lr) _ names(zlist[[1]]$test)
  for (i in 1:nsplin) {
    l1 _ nfixed+nsplin+(i-1)*(nknot+2)+1
    l2 _ l1+nknot+1
  }
}

```

```

np1.t _ length(zlist[[1]]$coef)
np.t_np1.t*g
theta.a_NULL
np.a_0
for(j in 1:g){
theta.up_ zlist[[j]]$coef[c(nfixed+i,11:12)]
theta.a_c(theta.a,theta.up)
np.a_np.a+length(theta.up)
}
np1.a_np.a/g
v.a _ matrix(0,nrow=np.a,ncol=np.a)
h.a _ v.a
q.a_v.a
lim.at_NULL
npa.j_0
npt.j_0
for(j in 1:g){
lim.tt_c(npt.j+nfixed+i,(npt.j+l1):(npt.j+l2))
lim.t_c(nfixed+i,11:12)
lim.a_c((npa.j+1):(npa.j+np1.a))
lim.at_c(lim.at,lim.a)
v.a[lim.a,lim.a]_zlist[[j]]$var[lim.t,lim.t]
h.a[lim.a,lim.a]_zlist[[j]]$d2i[lim.t,lim.t]
npa.k_0
npt.k_0
for(k in 1:g){
    mult.ka_((k-1)*npa.k)
    mult.ktt_((k-1)*npt.k)
    mult.ja_((j-1)*npa.j)
    mult.jtt_((j-1)*npt.j)
    klim.a_c((mult.ka+1):(mult.ka+np1.a))
    klim.tt_c(mult.ktt+nfixed+i, (mult.ktt+l1):(mult.ktt+l2))
    jlim.a_c((mult.ja+1):(mult.ja+np1.a))
    jlim.tt_c(mult.jtt+nfixed+i, (mult.jtt+l1):(mult.jtt+l2))
    q.a[jlim.a,klim.a]_Ohat[jlim.tt,klim.tt]
    npa.k_npa.k+np1.a
    npt.k_npt.k+np1.t
}
npa.j_npa.j+np1.a
npt.j_npt.j+np1.t
}
ct.a _ diag(np.a)
theta.l_NULL
np.l_0
for(j in 1:g){
theta.up_zlist[[j]]$coef[c(11:12)]
theta.l_c(theta.l,theta.up)
np.l_np.l+length(theta.up)
}
np1.l_np.l/g
v.l _ matrix(0,nrow=np.l,ncol=np.l)
h.l _ matrix(0,nrow=np.l,ncol=np.l)
q.l_matrix(0,nrow=np.l,ncol=np.l)
lim.lt_NULL
npl.j_0
npt.j_0
for(j in 1:g){

```

```

lim.tt_c((npt.j+l1):(npt.j+l2))
lim.t_c(l1:l2)
lim.l_c((npl.j+1):(npl.j+np1.l))
lim.lt_c(lim.lt,lim.l)
v.l[lim.l,lim.l]_zlist[[j]]$var[l1:l2,l1:l2]
h.l[lim.l,lim.l]_zlist[[j]]$d2i[l1:l2,l1:l2]
npl.k_0
npt.k_0
for(k in 1:g){
    mult.kl_((k-1)*npl.k)
    mult.ktt_((k-1)*npt.k)
    mult.jl_((j-1)*npl.j)
    mult.jtt_((j-1)*npt.j)
    klim.l_c((mult.kl+1):(mult.kl+np1.l))
    klim.tt_c((mult.ktt+l1):(mult.ktt+l2))
    jlim.l_c((mult.jl+1):(mult.jl+np1.l))
    jlim.tt_c((mult.jtt+l1):(mult.jtt+l2))
    q.l[jlim.l,klim.l]_Qhat[jlim.tt,klim.tt]
    npl.k_npl.k+np1.l
    npt.k_npt.k+np1.t
}
npl.j_npl.j+np1.l
npt.j_npt.j+np1.t
}
ct.l _ diag(np.l)
out.an[[i]]_spline.test.naive(theta.a,h.a,v.a,q.a,ct.a)
out.ln[[i]]_spline.test.naive(theta.l,h.l,v.l,q.l,ct.l)
out.ar[[i]]_spline.test.rob(theta.a,h.a,v.a,q.a,ct.a)
out.lr[[i]]_spline.test.rob(theta.l,h.l,v.l,q.l,ct.l)
}
}
else {
out.an _ NULL
out.ln _ NULL
out.ar _ NULL
out.lr _ NULL
}
}
else stop("lists must be of same type")
return(out.an,out.ln,out.ar,out.lr,out.fn,out.fr)
}

spline.test.naive_function(theta, h, v, q, ct)
{
    # routine for wald like test for the gen lin hypoth ctxtheta=0
    #using estimates from a penalized likelihood
    # theta_parameter estimates
    # h=inverse of second derivative matrix from penalized likelihood
    # v=var-cov matrix of parameter estimates
    # ct=contrast matrix. tests ct theta = 0
    # first compute cthct' ctvct'
    h <- ct %*% h %*% t(ct)
    v <- ct %*% v %*% t(ct)
    q <- ct %*% q %*% t(ct)
    theta.t <- ct %*% theta
    # calc eigenvalues
    # first correct h for possible nonsymmetry due to roundoff:
}

```

```

h <- (t(h) + h)/2
h <- solve(chol(h))
q <- (t(q) + q)/2
#q <- solve(chol(q))
#q <- solve(q)
theta <- t(h) %*% theta
stat <- t(theta) %*% theta
#theta.t <- t(q) %*% theta.t
#stat <- t(theta.t) %*% theta.t
#stat <- t(theta.t) %*% q %*% theta.t
h <- t(h) %*% v %*% h
h <- eigen(h)$value
k <- length(theta)
df <- rep(1, k)
v <- rep(0, k)
v <- .Fortran("dwch",
               as.double(h),
               as.double(v),
               as.integer(df),
               as.integer(k),
               as.double(stat),
               double(1),
               as.integer(df))[[6]]
df <- sum(h)
v <- c(stat, v, df)
names(v) <- c("stat", "pv", "df")
v
}

spline.test.rob_function(theta, h, v, q, ct)
{
  # routine for wald like test for the gen lin hypoth ctxtheta=0
  #using estimates from a penalized likelihood
  # theta_parameter estimates
  # h=inverse of second derivative matrix from penalized likelihood
  # v=var-cov matrix of parameter estimates
  # ct=contrast matrix. tests ct theta = 0
  # first compute cthct' ctvc't'
  h <- ct %*% h %*% t(ct)
  v <- ct %*% v %*% t(ct)
  q <- ct %*% q %*% t(ct)
  theta.t <- ct %*% theta
  # calc eigenvalues
  # first correct h for possible nonsymmetry due to roundoff:
  h <- (t(h) + h)/2
  h <- solve(chol(h))
  q <- (t(q) + q)/2
  q <- solve(chol(q))
  #q <- chol(q)
  #q <- solve(q)
  #theta <- t(h) %*% theta
  #stat <- t(theta) %*% theta
  theta.t <- t(q) %*% theta.t
  stat <- t(theta.t) %*% theta.t
  #stat <- t(theta.t) %*% q %*% theta.t
  h <- t(h) %*% v %*% h
  h <- eigen(h)$value
}

```

```

#h <- t(h) %*% q %*% h
#h <- eigen(h)$value
k <- length(theta)
df <- rep(1, k)
v <- rep(0, k)
v <- .Fortran("dwch",
               as.double(h),
               as.double(v),
               as.integer(df),
               as.integer(k),
               as.double(stat),
               double(1),
               as.integer(df))[[6]]
df <- sum(h)
#df <- k
#v <- 1-pchisq(stat,df)
v <- c(stat, v, df)
names(v) <- c("stat", "pv", "df")
v
}

mgrayp1.simu_function<-(n,m,pcens,bhaz,alpha=0.25,dfr,nknots,conf.level){
cens_runif<-function(2*n,0,1)
cens_ifelse<-function(cens>pcens,1,0)
#z1 <- 100*runif(n)
#z2 <- 100*runif(n)
z1_runif<-function(n)
z1_(z1-mean(z1))/(sqrt(var(z1)))
#z1_seq<-function(from=-1.71,to=1.71,length=n)
z2_runif<-function(n)
z2_(z2-mean(z2))/(sqrt(var(z2)))
#z1 <- ifelse(z1<0.5,1,0)
#z2 <- ifelse(z2<0.5,1,0)
m.mat <- matrix(0,1,2)
#lamda1 <- bhaz*exp( 3 * z1 + 3 * z2)
#lamda2 <- bhaz*exp( 3 * z1 + 5 * z2)
#lamda1 <- bhaz*exp( 3 * z1 )
#lamda2 <- bhaz*exp( 3 * z1 )
lamda1 <- bhaz
lamda2 <- bhaz
count_0
mcox.fit.list_vector("list",m)
for(i in 1:m){

bivdata <- simu.bivexp(n, lamda1, lamda2, alpha)
biv.data <- as.data.frame(cbind(c(bivdata$x,bivdata$y),c(z1,z1),
                                c(z2,z2),cens,c(rep(1,n),rep(2,n)),rep(1,2*n)))
names(biv.data) <- c("eventt","z1","z2","cens","groups","stratas")
#cat("dim(biv.data) is:" )
#print(dim(biv.data))
#print(biv.data[1:10,])
}

```

```

#names(biv.data) <- c("x", "y", "z1", "z2", "cen1", "cen2")
cv1_biv.data$z1
cv2_biv.data$z2
#fail <- as.data.frame(cbind(cens))
#linear.cov <- array(c(cv1, cv2), dim=c(400, 2))
#spline.cov <- as.matrix(bivexp.dt1[,c("cv1", "cv2")])
##cov <- as.data.frame(cbind(cv1, cv2))
##names(cov)_c("x1", "x2")
cov <- as.data.frame(cbind(cv1, cv2))
names(cov)_c("x1", "x2")
#linear.cov <- array(c(cv1))
#spline.cov <- array(c(cv1, cv2), dim=c(400, 2))
cov_as.matrix(cov)
mcox.fit_mcoxmp.spline(gind=biv.data$groups, m.type="a",
                        time=biv.data$eventt, status=biv.data$cens,
                        spline.cov=cov[,1], tstrata=biv.data$stratas,
                        dof=dfr[1], nknot=nknots)
out.an_as.vector(mcox.fit$mcox.out$out.an$spl1[2])
out.ar_as.vector(mcox.fit$mcox.out$out.ar$spl1[2])
##cat("spl1[2] is:")
##print(c(out.an, out.ar))
##print(ifelse(out.an < conf.level, 1, 0))
##print(as.numeric(ifelse(out.ar < conf.level, 1, 0)))
m.mat[1,1]_m.mat[1] + as.vector(ifelse(out.an < conf.level, 1, 0))
m.mat[1,2]_m.mat[2] + as.vector(ifelse(out.ar < conf.level, 1, 0))
#m.mat[1,3]_m.mat[1,3] + ifelse(mcox.fit$mcox.out$out.ln$x1[2] < 0.05, 1, 0)
#m.mat[1,4]_m.mat[1,4] + ifelse(mcox.fit$mcox.out$out.ln$x1[2] < 0.05, 1, 0)
mcox.fit.list[[i]]_mcox.fit$mcox.out
count_count+1
#x <- as.matrix(bivexp.dt1[,c("cv1", "cv2")])
#z <- cox.spline("a", tevt, cens1, spline.cov=x, strata=strata1,
#                 df=c(3,3), nknot=10)
}
##cat("m.mat is:")
##print(m.mat)
##print(is.matrix(m.mat))
m.mat_as.data.frame(m.mat)
names(m.mat)_c("gla.n", "gla.r")
m.mat_m.mat/count
return(count, m.mat, mcox.fit.list, mcox.fit$mcox.out)
}

simu.bivexp <- function(n, lamda1, lamda2, alpha = 0.25)
{
# n ..... number of observations per strata
# x,y.. ..... event time for the bivariate
# lamda1, lamda2 corresponding hazards

#u1 <- rep(0,n)
#u2 <- rep(0,n)
#x <- rep(0,n)
u1 <- runif(n)
u2 <- runif(n)
x <- -log(1 - u1)/lamda1
a <- rep(0,n)
b <- rep(0,n)
c <- rep(0,n)

```

```

v <- alpha*(2*exp(-x*lamda2)-1)
a <- v
b <- -(1+v)
c <- (1-u2)
w2 <- (-b - sqrt(b**2-4*a*c))/(2*a)
y <- -log(1-w2)/lamda2
return(x,y)
}

#####
# file contains s code defining functions spline.test cox.spline.int2
# cox.spline.plot cox.spline.mresid cox.spline, and to read jasa.data
#
# This software comes with absolutely no guarantees. You have permission to
# use it for any noncommercial purpose, and to modify it as needed.
# Copyright 1992 by Robert Gray
#
#####

spline.test _ function(theta,h,v,ct) {
# routine for wald like test for the gen lin hypoth ctxtheta=0
#using estimates from a penalized likelihood
# theta_parameter estimates
# h=inverse of second derivative matrix from penalized likelihood
# v=var-cov matrix of parameter estimates
# ct=contrast matrix. tests ct theta = 0
# first compute cthct' ctvct'
h _ ct%*%h%*%t(ct)
v _ ct%*%v%*%t(ct)
theta _ ct%*%theta
# calc eigenvalues
# first correct h for possible nonsymmetry due to roundoff:
h _ (t(h)+h)/2
h _ solve(chol(h))
theta _ t(h)%*%theta
stat _ t(theta)%*%theta
h _ t(h)%*%v%*%h
h _ eigen(h)$value
k _ length(theta)
df _ rep(1,k)
v _ rep(0,k)
v _
.Fortran("dwch",as.double(h),as.double(v),as.integer(df),as.integer(k),
         as.double(stat),double(1),as.integer(df))[[6]]
df _ sum(h)
v _ c(stat,v,df)
names(v) _ c("stat","pv","df")
v
}

cox.spline.int2 _ function(z1,z2) {
# z1 and z2 are lists created by cox.spline on different subgroups

```

```

# but with the same variables and knots (options "a" and "t" only).
if (any(z1$nvar != z2$nvar) | (length(z1$coef) != length(z2$coef)))
  stop("models must include the same variables")
if (z1$type == "a" & z2$type == "a") {
  nsplin _ z1$nvar[2]
  nknot _ length(z1$knot)/nsplin-6
  nfixed _ z1$nvar[1]
  if (nfixed>0) {
    outf _ vector("list",nfixed)
    names(outf) _ names(z1$coef)[1:nfixed]
    for (i in 1:nfixed) {
      stat _ (z1$coef[i]-z2$coef[i])/sqrt(z1$var[i,i]+z2$var[i,i])
      pv _ 1-pchisq(stat*stat,1)
      outf[[i]] _ c(stat,pv)
      names(outf[[i]]) _ c("stat","pv")
    }
  }
  else outf _ NULL
  if (nsplin>0) {
    out _ vector("list",nsplin)
    names(out) _ names(z1$test)
    for (i in 1:nsplin) {
      l1 _ nfixed+nsplin+(i-1)*(nknot+2)+1
      l2 _ l1+nknot+1
      theta _ c(z1$coef[c(nfixed+i,l1:l2)],z2$coef[c(nfixed+i,l1:l2)])
      np _ length(theta)
      np2 _ np/2
      h _ matrix(0,nrow=np,ncol=np)
      h[1:np2,1:np2] _ z1$d2i[c(nfixed+i,l1:l2),c(nfixed+i,l1:l2)]
      h[(np2+1):np,(np2+1):np] _ z2$d2i[c(nfixed+i,l1:l2),c(nfixed+i,l1:l2)]
      v _ matrix(0,nrow=np,ncol=np)
      v[1:np2,1:np2] _ z1$var[c(nfixed+i,l1:l2),c(nfixed+i,l1:l2)]
      v[(np2+1):np,(np2+1):np] _ z2$var[c(nfixed+i,l1:l2),c(nfixed+i,l1:l2)]
      ct _ cbind(diag(np2),-diag(np2))
      out[[i]] _ spline.test(theta,h,v,ct)
    }
  }
  else out _ NULL
  c(outf,out)
}
else if (z1$type == "t" & z2$type == "t") {
  nsplin _ z1$nvar[2]
  nfixed _ z1$nvar[1]
  nb _ (length(z1$coef)-nfixed)/nsplin
  if (nfixed>0) {
    outf _ vector("list",nfixed)
    names(outf) _ names(z1$coef)[1:nfixed]
    for (i in 1:nfixed) {
      stat _ (z1$coef[i]-z2$coef[i])/sqrt(z1$var[i,i]+z2$var[i,i])
      pv _ 1-pchisq(stat*stat,1)
      outf[[i]] _ c(stat,pv)
      names(outf[[i]]) _ c("stat","pv")
    }
  }
  else outf _ NULL
  if (nsplin>0) {
    out _ vector("list",nsplin)
  }
}

```

```

names(out) _ names(z1$test)
for (i in 1:nsplin) {
  l1 _ nfixed+(i-1)*nb+1
  l2 _ l1+nb-1
  theta _ c(z1$coef[l1:l2],z2$coef[l1:l2])
  np _ length(theta)
  np2 _ np/2
  h _ matrix(0,nrow=np,ncol=np)
  h[1:np2,1:np2] _ z1$d2i[l1:l2,l1:l2]
  h[(np2+1):np,(np2+1):np] _ z2$d2i[l1:l2,l1:l2]
  v _ matrix(0,nrow=np,ncol=np)
  v[1:np2,1:np2] _ z1$var[l1:l2,l1:l2]
  v[(np2+1):np,(np2+1):np] _ z2$var[l1:l2,l1:l2]
  ct _ cbind(diag(np2),-diag(np2))
  out[[i]] _ spline.test(theta,h,v,ct)
}
}
else out _ NULL
c(outf,out)
}
else stop("lists must be of same type")
}

cox.spline.mresid _ function(time,status,spline.cov,coef,spline.knot,
linear.cov,strata)  {

no _ length(time)
if (length(status) != no) stop("status wrong length")

# note that splin covs are not first in cov matrix if linear covs are:
if (missing(linear.cov)) {
  linear.cov _ as.matrix(spline.cov)
  if (no != nrow(linear.cov)) stop("nrow mismatch in spline.cov")
  nsx _ ncol(linear.cov)
}
else {
  linear.cov _ as.matrix(linear.cov)
  if (no != nrow(linear.cov)) stop("nrow mismatch in linear.cov")
  spline.cov _ as.matrix(spline.cov)
  if (no != nrow(spline.cov)) stop("nrow mismatch in spline.cov")
  nsx _ ncol(spline.cov)
  linear.cov _ cbind(linear.cov,spline.cov)
}
nfx _ ncol(linear.cov)

if (missing(strata)) {
  strata _ rep(1,no)
} else {
  if (length(strata) != no) stop("strata vector is the wrong length")
}
strata _ codes(factor(strata))

nomiss _ !(is.na(time) | is.na(status) | is.na(strata) |
           is.na(linear.cov %*% rep(1,nfx)))
# check status variable
if (any(status[nomiss] != 0 & status[nomiss] != 1))
  stop("invalid status values")

```

```

time _ time[nomiss]
status _ status[nomiss]
strata _ strata[nomiss]
linear.cov _ linear.cov[nomiss,]

no _ length(time)
nostr _ table(strata)
nstr _ length(nostr)
nostr _ cumsum(nostr)

nk1 _ length(spline.knot)/nsx - 6
nknot <- (length(coef)-nfx)/nsx - 2
nknot _ length(spline.knot)/nsx - 6

elp _ .Fortran("cests",as.double(linear.cov),
               as.integer(no),as.integer(nfx),as.integer(nsx),
               as.double(coef),as.integer(length(coef)),
               as.integer(nknot),
               as.double(spline.knot),as.integer(nk1),double(no))[[10]]

o _ order(strata,time)
mresid _ .Fortran("mresd",as.double(time[o]),as.double(status[o]),
                   as.double(elp[o]),as.integer(no),
                   as.integer(nstr),as.integer(nostr))[[3]]
mresid[o] _ mresid
mresid
}

cox.spline.plot _ function(z,ncovplot,xlab,main="",
                           ylab,zlab="Log Hazard Ratio",ylim,plotopt=T,pvar=T,knots=F,font=3,lwd=1,...)
{
if (z$type == "i") {           # plot for interaction
  x_sort(unique(z$est[,1]))
  y_sort(unique(z$est[,2]))
  u_z$est[,3]
  dim(u)_c(length(y),length(x))
  u_t(u)
  if (plotopt) {
    if (missing(xlab)) xlab _ dimnames(z$est)[[2]][1]
    if (missing(ylab)) ylab _ dimnames(z$est)[[2]][2]
    if (!exists(".Device")) stop("Graphics device must be active")
    persp(x,y,u,xlab=xlab,ylab=ylab,zlab=zlab,font=font,lwd=lwd,...)
    mtext(main,cex=1.5)
  }
}
else if (z$type == "t") {      # plots for tvc's
  ec _ 2*ncovplot
  if (pvar) {
    u _ ifelse(z$est[,ec+1]<0,0,z$est[,ec+1])
    u _ 2*sqrt(u)
    uu _ z$est[,ec]+u
    u _ z$est[,ec]-u
    if (missing(ylim)) ylim _ range(c(u,uu))
  }
  else if (missing(ylim)) ylim _ range(z$est[,ec])
  if (plotopt) {
}

```

```

if (missing(xlab)) xlab _ "Years"
if (missing(ylab)) ylab _ "Log Hazard Ratio"
if (!exists(".Device")) stop("Graphics device must be active")
plot(z$est[,1],z$est[,ec],main=main,xlab=xlab,ylab=ylab,ylim=ylim,
    type="l",font=font,lwd=lwd,...)
if (pvar) {
  lines(z$est[,1],uu,lty=2,lwd=lwd)
  lines(z$est[,1],u,lty=2,lwd=lwd)
}
}
ylim
}
else if (z$type == "a") { # plots for covs in additive model
  nsplinecov _ z$nvar[2]
# nsplinecov is the # spline covs in the model, ncovplot the # of the
# one being plotted
  o _ order(z$est[,ncovplot])
  ec _ nsplinecov+2*ncovplot-1
  if (pvar) {
    u _ ifelse(z$est[,ec+1]<0,0,z$est[,ec+1])
    u _ 2*sqrt(u)
    uu _ z$est[,ec]+u
    u _ z$est[,ec]-u
    if (missing(ylim)) ylim _ range(c(u,uu))
  }
  else if (missing(ylim)) ylim _ range(z$est[,ec])
  if (plotopt) {
    if (!exists(".Device")) stop("Graphics device must be active")
    if (missing(xlab)) xlab _ dimnames(z$est)[[2]][ncovplot]
    if (missing(ylab)) ylab _ "Log Hazard Ratio"
    plot(z$est[o,ncovplot],z$est[o,ec],main=main,xlab=xlab,ylab=ylab,ylim=ylim,
        type="l",font=font,lwd=lwd,...)
    if (pvar) {
      lines(z$est[o,ncovplot],uu[o],lty=2,lwd=lwd)
      lines(z$est[o,ncovplot],u[o],lty=2,lwd=lwd)
    }
    abline(h=0)
    if (knots) {
      ht _ .02*(ylim[2]-ylim[1])
      tk _ z$knots[4:(length(z$knots)-3)]
      ht_rep(ht,length(tk))
      segments(tk,ht,tk,-ht)
    }
  }
  ylim
}
}

cox.spline _ function(model.type,time,status,spline.cov,linear.cov,
strata,df,nknot,spline.knot,smooth.opt=1,smooth.param,output.opt="tests",
nest,maxiter=30,eps=1.e-4,rescale=T,ord=0,nstimes,stimes)  {

# nov: parameters in call to fortran routines
#   1=no, 2=nrx(=no), 3=ncx, 4=istr (col # of x for strata, -1 if none,
#   5=maxiter, 6=knot option (1 use knots provided, 0 program calculates),
#   7=smoothing option (<0 use input smoothing params, 0 calc smoothing
#   param only in first iteration, >0 recalc after each iteration)

```

```

# 8=nfx (nfx+nsx in sph, nfx is #col in linear.cov, nsx in spline.cov),
# 9=nsx, 10=analysis option (<=0 estimates only, 1 est & var, >1 est,
# var & test),
# sph: 12=nknot, 13=nest
# sphi: 11,12=nknx,nkny, 13,14=nestx,nesty, 15=rescale opt (0=yes)
# stvc: 12=order of spline, 13=nest, 14=nknot, 15=nstimes

nov _ rep(0,15)
nov[1] _ length(time)
if (length(status) != nov[1]) stop("status wrong length")

nov[5] _ maxiter

if (missing(linear.cov)) {
  linear.cov _ as.matrix(spline.cov)
  namspl _ dimnames(linear.cov)[[2]]
  namlin _ NULL
  if (nov[1] != nrow(linear.cov)) stop("nrow mismatch in spline.cov")
  nov[9] _ ncol(linear.cov)
  nov[8] _ 0
}
else {
  linear.cov _ as.matrix(linear.cov)
  namlin _ dimnames(linear.cov)[[2]]
  if (nov[1] != nrow(linear.cov)) stop("nrow mismatch in linear.cov")
  nov[8] _ ncol(linear.cov)
  spline.cov _ as.matrix(spline.cov)
  if (nov[1] != nrow(spline.cov)) stop("nrow mismatch in spline.cov")
  nov[9] _ ncol(spline.cov)
  linear.cov _ cbind(linear.cov,spline.cov)
  namspl _ dimnames(spline.cov)[[2]]
}
if (length(namspl) != nov[9] & nov[9]>0)
  namspl_paste("spl",(1:nov[9]),sep="")
if (length(namlin) != nov[8] & nov[8]>0) namlin_paste("lin",(1:nov[8]),sep="")

if (missing(strata)) {
  nov[4] _ -1
}
else {
  if (length(strata) != nov[1]) stop("strata vector is the wrong length")
  strata _ as.factor(strata)
  linear.cov _ cbind(linear.cov,codes(strata))
  nov[4] _ ncol(linear.cov)
}
nov[3] _ ncol(linear.cov)

nomiss _ !(is.na(time) | is.na(status) |
           is.na(linear.cov %*% rep(1,nov[3])))
# check status variable
if (any(status[nomiss] != 0 & status[nomiss] != 1))
  stop("invalid status values")
if (nov[4]<1) {
  status.table _ table(status[nomiss])
}
else {
  status.table _ table(strata[nomiss],status[nomiss])
}

```

```

}

nov[1] _ length(time[nomiss])
nov[2] _ nov[1]

if (nov[4] > 0) nstr _ length(unique(linear.cov[nomiss,nov[4]]))
else nstr _ 1

dsub _ rep(0,3*nov[9])
if (nov[9]>0) {
  if (!missing(smooth.param)) {
    if (model.type == "i") dsub[1] _ smooth.param
    else dsub[(nov[9]*2+1):(nov[9]*3)] _ smooth.param
    nov[7] _ -1
  }
  else {
    if (missing(df)) {
      if (model.type == "i") dsub[2] _ 8
      else dsub[1:nov[9]] _ rep(3,nov[9])
    }
    else {
      if (model.type == "i") {
        if (length(df) != 1) cat("warning:  length of df wrong\n")
        dsub[2] _ df
      } else {
        if (length(df) != nov[9]) cat("warning:  length of df wrong\n")
        dsub[1:nov[9]] _ df
      }
    }
    nov[7] _ max(smooth.opt,0)
  }
}

if (output.opt == "est") nov[10]_0
else if (output.opt == "var") nov[10]_1
else nov[10]_2

if (model.type == "a") {
  nfxx _ nov[8]
  nov[8] _ nov[8]+nov[9]
  linear.cov _ c(linear.cov[nomiss,],rep(0,nov[1]*nov[9]*4))

  if (missing(nknot)) nov[12] _ 10
  else nov[12] _ nknot[1]
  nknot _ nov[12]
  if (missing(spline.knot)) {
    nov[6] _ 0
    spline.knot _ rep(0,(nknot+6)*max(1,nov[9]))
  }
  else {
    if (length(spline.knot) != (nknot+6)*nov[9]) {
      nov[6] _ 0
      spline.knot _ rep(0,(nknot+6)*max(1,nov[9]))
    }
    else {
      nov[6] _ 1
    }
  }
}

```



```

        as.double(s0),
        as.double(s1),
        as.double(ts1),
        as.double(w),
        as.double(wp),
        as.double(ts1p),
        as.integer(id),
        as.double(slp),
        as.double(z11),
        as.double(z22),
        as.integer(gn))
#           as.integer(nf))
#cat('After sph\n')

# print (dsub[[22]])

if (dsub[[10]][1] <= 0) {
  if (dsub[[10]][1] < -2.5)
    stop("Possible overflow problems (No convergence)")
  else if (dsub[[10]][1] < -1.5)
    stop("Singular Penalized Information Matrix")
  else if (dsub[[10]][1] < -0.5) stop("Iteration did not converge")
  else stop("????")
}

# things in output vector:
# lik, beta, var, inv 2nd deriv, knots, tests, eigenvalues, estimates
# dsub is dsub[[12]]
# 2nd deriv mat in dsub(3*nov[9]+1)-dsub(3*nov[9]+np*np)

if (nknot != dsub[[1]][12]) {
  nknot <- dsub[[1]][12]
  np <- dsub[[1]][8]+(nknot+2)*dsub[[1]][9]
  dsub[[10]] <- dsub[[10]][1:(np*np)]
  dsub[[6]] <- dsub[[6]][1:np]
}
nsx _ dsub[[1]][9]
#cat('nsx is:\n')
#####
# unpenalized information matrix
  indunp _ dsub[[13]][5]
  maxind _ indunp+np*np-1
  unpinf _ dsub[[12]][indunp:maxind]
  dim(unpinf) _ c(np,np)
#####
  penmind _ dsub[[13]][4]
#  penmaxi _ penmind + np*np-1
#  penmaxi _ penmind + 4*(nknot+4)*nsx-1
#  penmaxi _ penmind + 4*nk3*nsx
#  penmatrix _ dsub[[12]][penmind:penmaxi]
  penm _ dsub[[12]][penmind:(indunp-1)]
***** cat('isub3-isub5 are')
***** print(c(dsub[[13]][3],penmind,dsub[[13]][5]))
***** cat('4*(nknot+4)*nsx\n')
***** print(4*(nknot+4)*nsx)
***** cat('length of pen matrix\n')

```

```

***** print(indunp-penmind)
dim(penm) _ c(nknot+4,4,nsx)
***** cat('*****penalty matrix*****\n')
***** print(penm)
***** cat('*****\n')
#####
print(nsx)
print(c((2*nsx+1):(3*nsx)))
print(np)
smoothp _ dsub[[12]][c((2*nsx+1):(3*nsx))]
***** cat('*****smoothing parameters*****\n')
print(smoothp)
cat('dim of dsub[[12]]\n')
print(length(dsub[[12]]))
cat('for dft,dfa,alpha\n')
print(3*nsx)
cat('penalty matrix\n')
print(4*(nknot+4)*nsx)
cat('unpen inf\n')
print(np*np)
# cat('***** DSUB[[12]] *****\n')
# print(dsub[[12]])

***** cat('*****\n')
#####

nvar _ c(nfxx,nov[9])
names(nvar) _ c("linear","spline")
dim(dsub[[2]]) _ c(nov[1])
dim(dsub[[3]]) _ c(nov[1])
dim(dsub[[17]]) _ c(nov[1],np)
dim(dsub[[18]]) _ c(nov[1])
dim(dsub[[19]]) _ c(nov[1])
dim(dsub[[20]]) _ c(nov[1],np)
dim(dsub[[22]]) _ c(nov[1],np)
dim(dsub[[23]]) _ c(nov[1],np)
dim(dsub[[27]]) _ c(nov[1],np)
dim(dsub[[28]]) _ c(nov[1],np)
dim(dsub[[29]]) _ c(nov[3],3)
#print(dsub[[17]][1:10,])
#cat("cens indicator:\n")
#print(dsub[[3]][1:20])
#print(dsub[[2]][1:10])

#cat("z :\n")
#print(dsub[[17]][1:20,])
#cat("z11 :\n")
#print(dsub[[27]][1:20,])
#cat("\n z22 :\n")
#print(dsub[[28]][1:20,])
#cat("\n gn :\n")
#print(dsub[[29]])
dim(dsub[[10]])_c(np,np)
####print(dsub[[10]])
tempval _ as.data.frame(c(dsub[[2]],dsub[[3]],dsub[[17]][,c(1,2)]),
# ncol=4)
cat("event time, status, and covariate values\n")

```

```

#print(tempval)
dim(dsub[[17]]) _ c(nov[1],np)
#cat("covariate values\n")
#print(dsub[[17]][,c(1,2)])
#cat("event time and status\n")
#print(c(dsub[[2]],dsub[[3]]))
names(dsub[[6]]) _ c(namlin,namspl,rep(namspl,rep(nknot+2,nov[9])))
dsub[[4]]_matrix(dsub[[4]],nrow=nov[1])[1:nov[13],c((nfxx+1):(nfxx+nov[9])),
(nov[3]+1):(nov[3]+2*nov[9]))]
v2_dsub[[12]][(3*nov[9]+1):(3*nov[9]+np*np)]
dim(v2)_c(np,np)
dim(dsub[[14]])_c(3,2*nov[9])
dim(dsub[[15]])_c(nk3,2*nov[9])
dsub[[5]]_matrix(dsub[[5]],ncol=max(nov[9],1))
dimnames(dsub[[14]])_list(c("stat","pv","df"),rep(c("overall","linear"),
nov[9]))
dimnames(dsub[[4]])_list(NULL,c(namspl,rep(c("est","var"),nov[9])))
dimnames(dsub[[15]])_list(NULL,rep(c("overall","lin"),nov[9]))
if (output.opt=="est") {
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],
  est=dsub[[4]],knots=dsub[[5]],nknot = nknot,
  smooth.param=dsub[[12]][(nov[9]*2+1):(nov[9]*3)],table=status.table,
  type="a",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
  isx=dsub[[14]], infom=unpinf,
  z=dsub[[17]], exz=dsub[[18]], s0=dsub[[19]], s1=dsub[[20]],
  tent = dsub[[2]], csin = dsub[[3]], penm = penm) }
else if (output.opt=="var") {
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],var=dsub[[10]],d2i=v2,
  est=dsub[[4]],knots=dsub[[5]],nknot = nknot,
  smooth.param=dsub[[12]][(nov[9]*2+1):(nov[9]*3)],table=status.table,
  type="a",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
  isx=dsub[[14]], infom=unpinf,
  z=dsub[[17]], exz=dsub[[18]], s0=dsub[[19]], s1=dsub[[20]],
  tent = dsub[[2]], csin = dsub[[3]], penm = penm) }
else {
  test _ vector("list",nov[9])
  names(test) _ namspl
  for (i in 1:nov[9]) test[[i]] _ dsub[[14]][,c(2*i-1,2*i)]
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],var=dsub[[10]],d2i=v2,
  est=dsub[[4]],knots=dsub[[5]],tests=test,eigs=dsub[[15]],nknot=nknot,
  smooth.param=dsub[[12]][(nov[9]*2+1):(nov[9]*3)],table=status.table,
  type="a",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
  isx=dsub[[14]], infom=unpinf,
  z=dsub[[17]], exz=dsub[[18]], s0=dsub[[19]], s1=dsub[[20]],
  tent = dsub[[2]], csin = dsub[[3]], penm = penm) }
}
else if(model.type == "i") {
  if (ncol(spline.cov) != 2)
    stop("Only 2 covariates allowed for tensor product spline")

linear.cov _ c(linear.cov[nomiss,],rep(0,nov[1]*8))

if (missing(nknot)) nov[11:12] _ 3
else nov[11:12] _ nknot
if (missing(spline.knot)) {
  nov[6] _ 0
  spline.knot _ rep(0,nov[11]+nov[12]+12)
}

```

```

}
else {
  if (length(spline.knot) != (nov[11]+nov[12]+12)) {
    nov[6] _ 0
    spline.knot _ rep(0,nov[11]+nov[12]+12)
  }
  else {
    nov[6] _ 1
  }
}

if (missing(nest)) nov[13:14] _ 20
else nov[13:14] _ nest

nb _ (nov[11]+4)*(nov[12]+4)-1
np _ nov[8]+nb
iwork _ rep(0,3*nov[1]+max(np,nov[1])+1+4*nstr)
isub _ rep(0,6+nov[8]+2*nov[1])
for (i in 1:nov[8]) isub[6+i] _ i
isub[1] _ nov[8]+1
isub[2] _ nov[8]+2
lgsub _ max(nov[13]*nov[14]*4+np*np,16*(nov[11]+4)*(nov[12]+4)+nb*nb+np*np+
  max(max(nov[1],4*np+np*np),np*np+nb*nb))
dsub _ c(dsub[1:3],rep(0,lgsub))

if (rescale) nov[15] _ 0
else nov[15] _ 1

u1 _ double(np)
u2 _ double(1)
v2 _ double(np*np)
*****
z _ double(nov[1]*np)
z1 _ double(nov[1]*np)
exz _ double(nov[1])
s0 _ double(nov[1])
s1 _ double(nov[1]*np)
ts1 _ double(np)
w _ double(nov[1]*np)
wp _ double(nov[1]*np)
ts1p _ double(np)
id _ integer(nov[1])
s1p _ double(nov[1]*np)
nf _ sum(status, na.rm=T)
*****
dsub _ .Fortran("sphi",as.integer(nov),
  as.double(time),
  as.double(status),
  as.double(linear.cov),
  as.double(spline.knot),
  u1,u2,c(u2,u2),u1,
  as.double(v2),
  as.integer(iwork),
  as.double(dsub),
  as.integer(isub),
  double(12),
  nov[13:14],
  nov[15])

```

```

double(4*nb),
as.double(eps),
as.double(z),
as.double(exz),
as.double(s0),
as.double(s1),
as.double(ts1),
as.double(w),
as.double(wp),
as.double(ts1p),
as.integer(id),
as.double(s1p),
as.integer(z1))
#      as.integer(nf))

if (dsub[[10]][1] <= 0) {
  if (dsub[[10]][1] < -2.5)
    stop("Possible overflow problems (No convergence)")
  else if (dsub[[10]][1] < -1.5)
    stop("Singular Penalized Information Matrix")
  else if (dsub[[10]][1] < -0.5) stop("Iteration did not converge")
  else stop("????")
}

# things in output vector:
# lik, beta, var, inv 2nd deriv, knots, tests, eigenvalues, estimates
# dsub is dsub[[12]]
# estimates in dsub(4)-dsub(3+nestx*nesty*4)
# 2nd deriv mat in dsub(4+nestx*nesty)-dsub(3+nestx*nesty+np*np)

nvar _ nov[8:9]
names(nvar) _ c("linear","spline")

names(dsub[[6]])_c(namlin,paste("spline",1:nb,sep=""))
ne2 _ nov[13]*nov[14]
linear.cov _ dsub[[12]][4:(3+ne2*4)]
dim(linear.cov) _ c(ne2,4)
v2 _ dsub[[12]][(4+4*ne2):(3+4*ne2+np*np)]
dim(v2) _ c(np,np)
dim(dsub[[10]]) _ c(np,np)
dim(dsub[[14]]) _ c(3,4)
dim(dsub[[15]]) _ c(nb,4)
dimnames(dsub[[14]]) _ list(c("stat","pv","df"),
  c("overall",namspl,"interact"))
dimnames(linear.cov) _ list(NULL,c(namspl,"est","var"))
dimnames(dsub[[15]]) _ list(NULL,c("overall",namspl,"interact"))
if (output.opt=="est") {
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],
    est=linear.cov,nknot=nknot,
    knots=dsub[[5]], smooth.param=dsub[[12]][1],table=status.table,
    type="i",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
    isx=dsub[[14]], infom=unpinf)
}
else if (output.opt=="var") {
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],var=dsub[[10]],
    d2i=v2,est=linear.cov,nknot=nknot,
    knots=dsub[[5]],smooth.param=dsub[[12]][1],table=status.table,
    type="v",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
    isx=dsub[[14]], infom=unpinf)
}

```

```

    type="i",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
    isx=dsub[[14]], infom=unpinf)
}
else {
  list(loglik=c(dsub[[7]],dsub[[8]]),coef=dsub[[6]],var=dsub[[10]],
    d2i=v2,est=linear.cov,knots=dsub[[5]],nknot=nknot,
    tests=dsub[[14]],eigs=dsub[[15]],smooth.param=dsub[[12]][1],
    table=status.table,type="i",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
    isx=dsub[[14]],infom=unpinf)
}
}
else if (model.type == "t") {

if (ord != 0 & ord != 2 & ord != 3) stop("ord must be 0,2 or 3 for tvc")

nov[12] _ ord

if (missing(nknot)) nknot _ 10
  else nknot _ nknot[1]
nov[14] _ nknot
if (missing(spline.knot)) {
  nov[6] _ 0
  spline.knot _ rep(0,(nknot+6))
}
else {
  if (length(spline.knot) != (nknot+6)) {
    nov[6] _ 0
    spline.knot _ rep(0,(nknot+6))
  }
  else {
    nov[6] _ 1
  }
}
iord1 _ ord+1
ncp _ iord1
if (missing(nest)) nest_100
if (ord==0) {
  nstimes _ nknot
# if ord==0 use interior knots for switch times
# so if knots specified need to use them.
  if (nov[6]==1) stimes _ spline.knot[4:(nknot+3)]
  else stimes _ rep(0,nknot)
  ntud _ nstimes+1
  ncp _ 2
  nww _ ntud
  nest _ 2*ntud
}
else {
  if (!missing(stimes)) {
    nstimes _ length(stimes)
    ntud _ nstimes+1
    nww _ ntud
    nest _ 2*ntud
  }
  else {
    if (missing(nstimes)) {
      nstimes _ nov[1]+1
    }
  }
}

```

```

        stimes _ 0
        ntud _ 1
        nww _ nov[1]
    }
    else {
        stimes _ rep(0,nstimes)
        ntud _ nstimes+1
        nww _ ntud
        nest _ 2*ntud
    }
}
nov[15] _ nstimes
nov[13] _ nest
nbas _ nknot+iord1
np _ nov[8]+nov[9]*nbas
iwork _ rep(0,3*nov[1]+max(np,nov[1])+1+(2+2*ntud)*nstr)
isub _ rep(0,8+nov[8]+nov[9]+nww)
l _ 0
if (nov[8]>0) for (i in 1:nov[8]) {
    l _ l+1
    isub[8+1] _ 1
}
for (i in 1:nov[9]) {
    l _ l+1
    isub[8+1] _ 1
}
if (smooth.opt>0) {
    nkk _ max(np-nbas,nbas)
    kmd _ max((np-nbas)*(np-nbas),nbas*nbas+2*nbas*nkk)
}
else {
    kmd _ nbas*nbas
}
ldsub _ ncp*(nknot+4)+iord1*nww+np*np+kmd+max(nov[1],4*np+np*np)
ldsub _ max(ldsub,iord1*nww+nest*(2*nov[9]+1))
dsub _ c(dsub,rep(0,ldsub))
u1 _ double(np)
u2 _ double(1)
v2 _ double(np*np)
*****
z _ double(nov[1]*np)
z1 _ double(nov[1]*np)
exz _ double(nov[1])
s0 _ double(nov[1])
s1 _ double(nov[1]*np)
ts1 _ double(np)
w _ double(nov[1]*np)
wp _ double(nov[1]*np)
ts1p _ double(np)
id _ integer(nov[1])
s1p _ double(nov[1]*np)
nf _ sum(status, na.rm=T)
*****
dsub _ .Fortran("stvc",as.integer(nov),
                 as.double(time[nomiss]),

```

```

as.double(status[nomiss]),
as.double(linear.cov[nomiss,]),
as.double(spline.knot),
as.double(stimes),
u1,u2,c(u2,u2),u1,
as.double(v2),
as.integer(iwork),
as.double(dsub),
as.integer(isub),
as.double(3*2*nov[9]),
as.double(nbas*2*nov[9]),
as.double(v2),
as.double(eps),
as.double(z),
as.double(exz),
as.double(s0),
as.double(s1),
as.double(ts1),
as.double(w),
as.double(wp),
as.double(ts1p),
as.integer(id),
as.double(s1p),
as.integer(z1))
#      as.integer(nf))

if (dsub[[11]][1] <= 0) {
  if (dsub[[11]][1] < -2.5)
    stop("Possible overflow problems (No convergence)")
  else if (dsub[[11]][1] < -1.5)
    stop("Singular Penalized Information Matrix")
  else if (dsub[[11]][1] < -0.5) stop("Iteration did not converge")
  else stop("????")
}

# things in output vector:
# lik, beta, var, inv 2nd deriv, knots, tests, eigenvalues, estimates
# dsub is dsub[[13]]

nvar _ nov[8:9]
names(nvar) _ c("linear","spline")

names(dsub[[7]]) _ c(namlin,rep(namspl,rep(nbas,nov[9])))
dim(dsub[[11]]) _ c(np,np)
dim(dsub[[17]]) _ c(np,np)
j1 _ dsub[[14]][5]
if (ord==0) nest_2*dsub[[1]][14]+2
linear.cov _ dsub[[13]][j1:(j1-1+nest*(2*nov[9]+1))]
dim(linear.cov) _ c(nest,2*nov[9]+1)
dim(dsub[[15]]) _ c(3,2*nov[9])
dim(dsub[[16]]) _ c(nbas,2*nov[9])
dimnames(dsub[[15]]) _ list(c("stat","pv","df"),
  rep(c("overall","nonprop"),nov[9]))
dimnames(linear.cov) _ list(NULL,c("times",rep(c("est","var"),nov[9])))
dimnames(linear.cov)[[2]][2*(1:nov[9])] _ namspl
dimnames(dsub[[16]]) _ list(NULL,rep(c("overall","nonprop"),nov[9]))

```

```

test _ vector("list",nov[9])
names(test) _ namspl
for (i in 1:nov[9]) test[[i]] _ dsub[[15]][,c(2*i-1,2*i)]

if (output.opt=="est") {
list(loglik=c(dsub[[8]],dsub[[9]]),coef=dsub[[7]],
  est=linear.cov,knots=dsub[[5]],nknot=nknot,
  smooth.param=dsub[[13]][(nov[9]*2+1):(nov[9]*3)],stimes=dsub[[6]],
  table=status.table,type="t",nvar=nvar,w=dsub[[22]],wp=dsub[[23]],
  isx=dsub[[14]],infom=unpinf)
}
else if (output.opt=="var") {
list(loglik=c(dsub[[8]],dsub[[9]]),coef=dsub[[7]],var=dsub[[11]],d2i=dsub[[17]],
  est=linear.cov,knots=dsub[[5]],nknot=nknot,
  smooth.param=dsub[[13]][(nov[9]*2+1):(nov[9]*3)],stimes=dsub[[6]],
  table=status.table,type="t",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
  isx=dsub[[14]],infom=unpinf)
}
else {
list(loglik=c(dsub[[8]],dsub[[9]]),coef=dsub[[7]],var=dsub[[11]],d2i=dsub[[17]],
  est=linear.cov,knots=dsub[[5]],tests=test,eigs=dsub[[16]],nknot=nknot,
  smooth.param=dsub[[13]][(nov[9]*2+1):(nov[9]*3)],stimes=dsub[[6]],
  table=status.table,type="t",nvar=nvar,w=dsub[[22]], wp = dsub[[23]],
  isx=dsub[[14]], infom=unpinf)
}
}
}

#jasa.data <- read.table("jasa.data",header=T,row.names="id")

```

APPENDIX 5. COPY OF FORTRAN CODE NEEDED TO RUN GRAY'S MULTIPLE OUTCOME SURVIVAL MODEL

```

c
c This software comes with absolutely no guarantees. You have permission to
c use it for any noncommercial purpose, and to modify it as needed.
c Copyright 1992 by Robert Gray
c

      subroutine cests(x,no,nfx,nsx,beta,np,nknot,tknot,nkl,elp)
c routine for calculating function estimates and variances
c from output of sph/wsph.
c x is covariate values including spline terms. for info on ordering
c etc see sph.f
      double precision x(no,nfx),w(4),elp(no)
      double precision beta(np),tknot(-2:(nkl+3),nsx)
      isx=nfx-nsx
      nk2=nknot+2
      do 88 i=1,no
        elp(i)=0
        do 10 j=1,nfx
          10   elp(i)=elp(i)+x(i,j)*beta(j)
          do 5 k=1,nsx
            call spln(nknot,3,tknot(-2,k),x(i,isx+k),1,intbl,w,1)
            13=nfx+nk2*(k-1)+intbl-1
            14=min(nk2,intbl+3)-intbl+1
            do 6 iq=1,14
              6    elp(i)=elp(i)+w(iq)*beta(13+iq)
              5  continue
              5    elp(i)=exp(elp(i))
88  continue
      return
      end

      subroutine mresd(s,c,x,n,nstr,nostr)
      double precision s(n),c(n),x(n),xb,dd,ch
      integer nostr(nstr)

c
c subroutine to calculate martingale residuals in proportional
c hazards models
c                                     and s the survival times
c on input c is the censoring indicator ^ and x is exp(x'b) for each case
c on output x is overwritten with resids for each case.
c data is assumed sorted on survival (ascending) within strata
c

      ll=1
      do 10 k=1,nstr
        xb=0
        do 19 i=ll,nostr(k)
19      xb=xb+x(i)
        ch=0
        dd=0
        13=ll
        do 20 i=ll,nostr(k)
          if (s(i).eq.s(13)) then
            dd=dd+c(i)
          else
            if (dd.gt.0) ch=ch+dd/xb
            dd=c(i)
            do 22 ii=13,i-1
              xb=xb-x(ii)
              x(ii)=c(ii)-x(ii)*ch
22      continue
            13=i
          endif
20      continue
          if (dd.gt.0) ch=ch+dd/xb

```

```

      do 23 ii=l3,nostr(k)
23    x(ii)=c(ii)-x(ii)*ch
      l1=nostr(k)+1
10   continue
      return
end

      subroutine calvar(vv,np,lvv,v3,lv3,var,lvar,work)
c input: vv actual 2nd deriv matrix, v3 var cov matrix of the scores
c np matrices are npxnp, with actual row dimensions given by l* terms
c output: var is var cov matrix, vv^-1 v3 vv^-1. vv and var can be the
c same. work (entries 1 to np*np) will have vv^-1
      double precision vv(lvv,np),v3(lv3,np),var(lvar,np),work(np,np)
c first compute full variance covariance matrix:
      call pdi(np,vv,work,lvv,np)
      do 10 i=1,np
      do 10 j=i,np
      var(i,j)=0
      do 11 ii=1,np
      do 11 jj=1,np
11    var(i,j)=var(i,j)+work(i,ii)*v3(ii,jj)*work(jj,j)
      var(j,i)=var(i,j)
10   continue
      return
end

      subroutine coxrg(s,c,x,n,nrx,ncx,istrat,np,mit,iflg,ntime,
     &times,beta,lik0,lik,scv,vinf,iwork,dwork,act,eval,sub2,
     &isub,dsub,eps,icnv,z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1,gn)
c input:
c      s      survival times
c      c      censoring indicator (1=failure, 0=censoring)
c      x      matrix, ith row contains covariate information on
c              the ith case.
c      n      sample size (# rows used in s,c and x)
c      nrx,ncx actual row dim of x in calling program, and number of
c              cols of x actually used
c      istrat column of x which contains the strata variable.  istrat<=0
c              means only one stratum.
c      np      total # parameters in the model
c      mit     max # iterations, included primarily so you
c              can specify 1 to get info for score tests.
c      iflg    this program creates a variety of indexing information
c              which is stored in portions of iwork.
c              and also does some permuting of rows of s,c, and
c              x. On subsequent calls with the same values of
c              s,c,x, can save recalculating this by setting
c              iflg>0. If iflg<=0 then this info is calculated
c      ntime   # of times at which time varying covariates (tvc) are allowed
c              to change values. If ntime=0 assumes no tvc. Set ntime to
c              a value>n to use all failure times.
c      times   a vector of length ntime containing the times where
c              covariates are allowed to change values. If ntime>n or
c              ntime<=0 then times is not used. If times(ntime)<0 then
c              ntime timepoints will be generated and stored in times,
c              with approx equal #'s of failures between the times.
c      beta    initial value of parameters. Usually should set this
c              to 0 before the call. Can be set to other values to
c              generate info for score tests, etc.
c      act,eval,sub2 names of user supplied subroutines called by this
c              program (see below)
c      isub,dsub integer and double precision arrays containing
c              information that is passed to the subroutines act,eval

```

```

c      & sub2
c      eps convergence criteria, stops when change in loglikelihood
c          is < eps
c
c      output:
c          on output rows of s c and x may have been permuted
c          beta values of parameters after termination of iteration
c          lik0 value of likelihood at initial parameter values
c          lik value of log likelihood after final iteration
c          scv value of the score vector components after termination
c          vinf information matrix after termination
c          icnv 0=converge, 1=did not converge, 2=sing inf matrix
c              3=could not find a better point, or overflow probs
c
c      workspace:
c          iwork integer, of length at least 3*n+max(n,np)+1+2*nstr+2*ntud*nstr,
c              where ntud is ntime+1, or ntud=1 (when ntime>n)
c          dwork double precision, of length at least max(n,4*np+np*np)
c
c      subroutines:
c      act, eval, and sub2 are the names of subroutines that must be
c      provided by the user. The names of these routine must be declared
c      external in the calling program. act is a routine that specifies
c      which model terms are active in the current time interval. The
c      reason for including this is that if your model includes tvc that
c      are products of fixed covariates with functions of time (created to
c      give flexible hazard ratio models over time), then if the functions
c      of time have local support properties (ie are 0 outside of a fixed
c      interval) then considerable efficiency can be gained by indicating
c      this through act. The call is
c          call act(lci,11,12,idx,n,sv,ce,xx,nrx,ncx,iact,nact,np,
c          &lc2,isub,dsub)
c      where lci indexes the time interval, and lc2 the strata. 11 and 12
c      and idx are such that sv(idx(11,1)) to sv(idx(11,2)) are the
c      survival times of the subjects who fail at the smallest failure time
c      in the interval, or are censored between this time and the next. 12
c      gives the same thing for the largest time. sv are the survival
c      times, ce the censoring times, and xx the cov info, n the sample size,
c      iact as output will need to contain which of the np covariates are active
c      on this time interval, eg if nact=4 and iact=c(1,3,4,6), it would
c      mean that there are contributions from the 1st, third, fourth and sixth
c      covariates in this time interval, but that the score and information
c      contributions of the other covariates are identically 0.
c
c      eval is a routine which specifies the value of the covariate vector
c      for subject i at time sv(1). The call is
c          call eval(g, gp, b, lci, i, 1, sv, ce, xx, nrx, n, np, lc2, nact, iact,
c          &isub, dsub),
c      where gp(j) will be the value of the jth cov (only gp(iact(j)), j=1,
c      nact will actually be used). lci gives the number of the time
c      interval, i the observation number, sv(1) the failure time of the
c      current likelihood contribution, with the other arguments as
c      specified in act. This routine also needs to calculate g=x'b for
c      case i.
c
c      sub2 is a routine called after each iteration. This can be a null
c      routine, or could be used to apply a penalty function to the scores
c      and information (for example). The call is
c          call sub2(np,beta,lik,scv,vinf,svi,isub,dsub,nit)
c      where nit is the current iteration count. One additional call
c      is made after convergence, with nit=-1. (svi is a working matrix
c      of size np*np, included because the extra space is sometimes needed)
c          change 4-18-91: extra call after convergence no longer made.

```

```

c
c In addition to the routines in this file (coxft, coxg, strat, ut,
c uft, tint) and the user
c supplied routines, calls are made to tint4 cholg, solve, dperm and sortg,
c which are in ~gray/src/util
c
c     double precision s(n),c(n),x(nr,ncx),scv(np),lik0,lik,
c     &vinf(np,np),dwork(1),times(1),beta(np),dsub(1),eps
C*****
c     double precision z(n,np),exz(n),s0(n),s1(n,np),z1(n,np)
c     double precision ts1(np),w(n,np)
C*****
C>>>>
c     double precision wp(n,np),ts1p(np),s1p(np,n)
c     integer id(n),gn(1)
C<<<<
c
c     integer iwork(1),isub(1)
c     external act,eval,sub2
C>>>>
c         do 15 i = 1, n
c             do 10 j = 1, np
c                 z1(i,j) = x(i,j)
c 10         continue
c 15         continue
C<<<<
c         mnstr=1
c         miuti=mnstr+1
c         mtduti=miuti+2*n
c miwk needs lenght n
c         miwk=mtduti+n
c         mnostr=miwk+max(n,np)
c         mdwk=1
c         if (ntime.lt.n) then
c ntud is the # time intervals
c         ntud=ntime+1
c         if (ntud.gt.1.and.times(ntime).le.0) then
c             do 5 i=1,n
c 5             iwork(miwk+i-1)=i
c             call sortg(s,iwork(miwk),1,n)
c             call dperm(c,iwork(miwk),1,n,dwork(mdwk))
c             do 6 j=1,ncx
c 6             call dperm(x(1,j),iwork(miwk),1,n,dwork(mdwk))
c             do 15 i = 1, n
c                 do 10 j = 1, np
c                     z1(i,j) = x(i,j)
c 10                 continue
c 15                 continue
c                 call tint4(ntud,s,c,times,n,iwork(miwk))
c                 write (6,*) (times(j),j=1,ntime)
c             endif
c             else if (ntud.lt.1) then
c                 ntud=1
c             endif
c index information:
c             if (iflg.le.0) then
c                 call strat(istrat,s,c,x,iwork(mnostr),nstr,iwork(miwk),
c &dwork(mdwk),n,ncx,nrx,iwork(miuti),iwork(mtduti))
c                 iwork(mnstr)=nstr
c             else
c                 nstr=iwork(mnstr)
c             endif
c             write (6,*) nstr,n

```

```

c      write (6,*)(iwork(mnostr+i-1),i=1,nstr)
mnuti=mnostr+nstr
c      write (6,*)(iwork(mnuti+i-1),i=1,nstr)
c      call pr1(iwork(miuti),iwork(mtduti),n)
mitud=mnuti+nstr
c length of itud is ntud*nstr*2
mdvi=mdwk+4*np
c time intervals:
if (ntime.lt.n) then
  l1=1
  do 225 l1=1,nstr
  l2=(l1-1)*2*ntud
    call tint(ntud,times,s,c,iwork(mitud+l2),n,iwork(miuti),
& iwork(mnuti+l1-1),l1)
    l1=iwork(mnostr+l1-1)+1
225   continue
endif
CC>>>>
C      do 15 i = 1, n
C          do 10 j = 1, np
C              z1(i,j) = x(i,j)
C 10      continue
C 15      continue
CC<<<<
c      call pr2(iwork(mitud),ntud,nstr)
c fit model:
call coxft(s,c,x,ncx,nrx,n,np,beta,lik0,lik,scv,vinf,mit,
&iwork(miuti),iwork(mtduti),iwork(mnuti),iwork(mitud),ntud,
&ntime,nstr,iwork(mnostr),iwork(miwk),dwork(mdvi),dwork(mdwk),
&act,eval,sub2,isub,dsub,eps,icnv,z,exz,s0,s1,ts1,w,
&wp,ts1p,id,s1p,z1,gn)
return
end

subroutine coxft(s,c,x,ncx,nrx,n,np,beta,lik0,lik,sb,vv,mit,
&iuti,tduti,nuti,itud,ntud,ntime,nstr,nostr,iwork,vvi,work,
&act,eval,sub2,isub,dsub,eps,icnv,z,exz,s0,s1,ts1,w,
&wp,ts1p,id,s1p,z1,gn)
double precision s(n),c(n),x(nrx,ncx),beta(np),lik0,liko,lik
double precision sb(np),vv(np,np),vvi(np,np),work(1),dsub(1)
double precision eps,shrnk,eps2
C*****
double precision z(n,np),exz(n),s0(n),s1(n,np),z1(n,np)
double precision ts1(np),w(n,np)
C*****
C>>>>
double precision wp(n,np),ts1p(np),s1p(np,n)
integer id(n),gn(ncx,3)
C<<<<
integer iuti(n,2),tduti(n),nuti(nstr),itud(ntud,2,nstr)
integer iwork(1),isub(1),nostr(nstr)
external act,eval
do 10 i = 1, n
  do 5 j = 1, np
    w(i,j) = 0.d0
    wp(i,j) = 0.d0
5 continue
10 continue
eps2=1.d+0
shrnk=.5d0
mshrk=4
icnv=1
npl=np+1

```

```

    call coxg(ntime,n,np,beta,s,c,x,ncx,nrx,iuti,itud,ntud,
    &tduti,nuti,sb,vv,lik,
    &vvi,work(2*np+1),work(3*np+1),nstr,nostr,iwork,
    &act,eval,isub,dsub,iprb,z,exz,s0,s1,ts1,w,id,gn,nfx)

C>>>>>
    do 2 i = 1, n
        do 1 j = 1, np
            z1(i,j) = x(i,j)
1         continue
2         continue
C<<<<<

c      write (6,100) ((vv(i,j),j=1,np),i=1,np)
c100  format (4e18.8)
    if (iprb.gt.0) then
        icnv=3
        return
    endif
    call sub2(np,beta,lik,sb,vv,vvi,isub,dsub,1)
    lik0=lik
    liko=lik
c      write (6,*) lik
    do 15 i=1,np
15   work(i)=beta(i)
    call cholg(np,vv,vvi,np,np)
    if (vvi(1,1).lt.0) then
        icnv=2
        return
    endif
    call solve(np,vvi,sb,work(np1),np)
    do 58 i=1,np
58   beta(i)=beta(i)+work(np+i)
    if (mit.le.1) return
    nshrk=0
    do 188 nit=2,mit
        call coxg(ntime,n,np,beta,s,c,x,ncx,nrx,iuti,itud,ntud,
        &tduti,nuti,sb,vv,lik,
        &vvi,work(2*np+1),work(3*np+1),nstr,nostr,iwork,
        &act,eval,isub,dsub,iprb,z,exz,s0,s1,ts1,w,id,gn,nfx)
        call sub2(np,beta,lik,sb,vv,vvi,isub,dsub,nit)
c      write (6,*) lik
        if (iprb.gt.0.or.lik.lt.liko-eps2) then
            if (nshrk.le.mshrk) then
                nshrk=nshrk+1
c      write (6,110)
c      format('shrinking')
        do 56 i=1,np
            work(np+i)=shrnk*work(np+i)
56   beta(i)=work(i)+work(np+i)
        go to 188
        else
            icnv=3
            return
        endif
c      else
c      nshrk=0
    endif
    do 455 i=1,np
455  work(i)=beta(i)
    call cholg(np,vv,vvi,np,np)
    if (vvi(1,1).lt.0) then
        icnv=2

```

```

        return
    endif
    call solve(np,vvi,sb,work(np1),np)
    do 456 i=1,np
456  beta(i)=beta(i)+work(np+i)
     if (abs(lik-liko).le.eps.and.nshrk.le.0) then
c         liko=0
c         do 668 i=1,np
c 668  liko=liko+work(np+i)*work(np+i)
c         if (liko.lt.np*1.d-6) then
            icnv=0
            go to 920
c         endif
c     endif
nshrk=0
liko=lik
188  continue
920  continue

CCCCCCCCCCCCC          ****
CC      call realpr('Covariate values')
do 925 i = 1, n
    if(.NOT.(c(i).gt.0.d-10))then
        do ii=1, nfx
            z(i,ii) = z1(i,ii)
        enddo
        do ii=1, ncx
            l1 = gn(ii,1)
            l3 = gn(ii,2)
            l4 = gn(ii,3)
            do l = 1,l4
                z(i,l3+l)=z1(i,l1+l)
            enddo
        enddo
    else
        goto 930
    endif
925 continue

930  ict = 0
    do kk=1,n
        ict=ict+c(kk)
    enddo
    do i = 1, n
        exz(i) = 0.d0
        do j = 1, np
            exz(i) = exz(i) + beta(j)*z(i,j)
            s1(i,j) = 0.d0
            s1p(j,i) = 0.d0
        enddo
    enddo
CC    call realpr(z(i,1),1)
    exz(i) = dexp(exz(i))
    s0(i) = 0.d0
    enddo
    do i = n, 1, -1
        if(i.eq.n)then
            s0(i) = exz(i)
            do j = 1, np
                s1(i,j)=z(i,j)*exz(i)
                s1p(j,i) = z(i,j)-s1(i,j)/s0(i)
            enddo
        else
            s0(i) = s0(i+1) + exz(i)
        endif
    enddo

```

```

        do j = 1, np
            s1(i,j) = s1(i+1,j) + z(i,j)*exz(i)
            s1p(j,i) = z(i,j)-s1(i,j)/s0(i)
        enddo
    endif
    call sub8n(np,beta,s1p(1,i),isub,dsub,n)
enddo
CCCCCCC *****
do 9 I = 1, n
    do 6 j = 1, np
        ts1(j) = 0.d0
        ts1p(j) = 0.d0
        mf = 0
        do l = 1, i
            mf = mf + c(l)
        enddo
        do 3 l = 1, i
            if(c(l).gt.1.d-10) then
                ts1(j) = ts1(j) + (exz(i)/s0(l))*(z(i,j) -
                & s1(l,j)/s0(l))
                & ts1p(j) = ts1p(j)+(exz(i)/s0(l))*(z(i,j) +
C           & (s1p(j,l)+s1(l,j)/s0(l) - z(l,j))*ict/mf-s1(l,j)/s0(l))
                ts1p(j)=ts1p(j)+(exz(i)/s0(l))*(z(i,j)-s1(l,j)/s0(l))
            endif
3 continue
6 continue
if(c(i).gt.1.d-10)then
    do j=1, np
        w(i,j) = (z(i,j)-s1(i,j)/s0(i)) - ts1(j)
        wp(i,j) = s1p(i,j) - ts1p(j)
    enddo
else
    do j = 1, np
        w(i,j) = -ts1(j)
        wp(i,j) = -ts1p(j)
    enddo
endif
c         write(6,*)(w(i,11),11=1,26)
9 continue
CCCCCCCCCCCCC *****
      return
end

subroutine sub8n(np2,beta,sb,isub,dsub,nf)
double precision beta(np),sb(np),dsub(1)
integer isub(1)
common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
call sub9n(beta,sb,dsub(isub(4)), dsub(malph),nf)
return
end

subroutine sub9n(beta,sb,penm,alpha,nf)
double precision beta(np),sb(np)
double precision penm(nk4,4,nsx),alpha(nsx)
common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
do 535 j=1,nsx
    l1=nfx+1+(j-1)*nk2
    call penscor(nk2,beta(l1),penm(1,1,j),nk4,4,sb(l1),alpha(j),
    &np,no,nf)

```

```

535 continue
      return
      end

      subroutine penscor(nb,beta,penm,mr,md,s,alpha,np,n,nf)
c given a penalty matrix penm and parameter values beta,
c subtracts penalty terms from the scores
c the beginning index of beta and s (the score) in the
c call should correspond to where the spline terms begin
c ie call penlik(...beta(7)...s(7)...) if the spline is
c stored consecutively beginning with the 7th component
c nk is the number of knots, nb is the number
c of basis fcns actually in use, alpha is the value of the
c smoothing parameter. Aug. 6' 2001
      double precision penm(mr,md),beta(1),s(1),alpha,u
      do 10 i=1,nb
         u=0
         m1=max(1,i-md+1)
         m2=min(nb,i+md-1)
         l=0
         do 12 j=i,m2
            l=l+1
            u=u+penm(i,l)*beta(j)
12     continue
         if (m1.lt.i) then
            l=i-m1+2
            do 11 j=m1,i-1
               l=l-1
               u=u+beta(j)*penm(j,l)
11     continue
         endif
         s(i)=s(i)-(alpha*u/nf)
10     continue
      return
      end

      subroutine coxg(nt,n,np,b,sv,ce,xx,ncx,nrx,idx,it,ntud,nd,
     &ndf,s,v,lik,vt,xb,gp,nstr,nostr,iact,act,eval,isub,dsub,iprb,
     &z,exz,s0,s1,ts1,w,id,gn,nfx)
c subroutine to calculate cox likelihood, score, and inf
C nt is # time intervals= #switch points +1.
c n is the # of observations, np the # parameters, b the current
c parameter values, idx and it are index information: idx should be
c from uft, giving for each strata the increment in the risk set
c at each unique failure time. it should be from tint, and gives
c the rows of idx which correspond to failures in each of the time
c intervals. nd is the # deaths at each unique failure time in each
c strata. ndf is the # of unique failure times (nuti in other programs)
c in each strata
c On output, s v and lik are the score, information (-second deriv)
c and likelihood. vt is an npxnp matrix xb and g are np dimensional
c vectors.
c also, iact is an integer working vector of length np.
c act and eval need to be names of subroutines that will return
c the active covariates within each time interval (act), and
c the covariate values for a specified obs within a specified interval
c (eval)
      double precision b(np),s(np),v(np,np),lik,u,dsub(1)
      double precision xb(np),xb1,xblo,u1,u2,vt(np,np),g, gp(np)
      double precision sv(n),ce(n),xx(nrx,ncx)
C*****
      double precision z(n,np),exz(n),s0(n),s1(n,np)
      double precision ts1(np),w(n,np)

```

```

C*****
      integer idx(n,2),nd(n),it(ntud,2,nstr),ndf(nstr),id(n)
      integer iact(np),nostr(1),isub(1),gn(ncx,3)
      inct = n
      iprb=0
      lik=0.d0
      do 210 i = 1, n
         do 205 j = 1, np
            z(i,j) = 0.d0
205  continue
210  continue
      do 1 j=1,np
      s(j)=0.d0
      do 2 j1=1,np
      v(j,j1)=0.d0
2     continue
1     continue
c   loop over times (lci) within strata (lc2)
      nost=0
      do 996 lc2=1,nstr
         if (nt.le.n) then
            ntime=ntud
         else
            ntime=ndf(lc2)-nost
         endif
         do 98 lci=ntime,1,-1
c get which covariates are structurally 0 on time interval i
c nact is the # nonzero, iact(i) is the # of the ith nonzero covar
c ie iact(1)=2 means that the 2nd covariate is not structurally 0.
         if (nt.le.n) then
            l2=it(lci,2,lc2)
            l1=it(lci,1,lc2)
         else
            l2=lci+nost
            l1=lci+nost
         endif
         if (l1.le.0) go to 98
         call act(lci,l1,l2,idx,n,sv,ce,xx,nrx,ncx,iact,nact,np,
&lc2,isub,dsub)
         xb1=0.d0
         xb1o=1
         ktmp=0
         do 11 j=1,np
            xb(j)=0.d0
            do 12 j1=j,np
               vt(j,j1)=0.d0
12       continue
11       continue
            l3=idx(ndf(lc2),2)
        4    do 10 k=l2,l1,-1
c k is ranging over unique failure times (largest to smallest)
c l to l4 are the failures at the kth unique failure time
c l to l3 are all obs at the kth unique failure but smaller than
c the next largest failure.
            l=idx(k,1)
            l4=l+nd(k)-1
CCC ****
            ncens = l3 - l4
            nf = l4 - l + 1
CCC ****
            15=l4+1
            do 26 i=1,l4
c get covariates for obs i on time interval lci (gp)

```

```

        call eval(g,gp,b,lci,i,1,sv,ce,xx,nrx,ncx,n,np,lc2,nact,iact,
&      isub,dsub)
        call eval8n(gn,lci,i,1,sv,ce,xx,nrx,ncx,n,np,lc2,nact,iact,
&      isub,dsub,nfx)
c g=x'b
c      g=0
c      do 842 j=1,nact
c 842  g=g+gp(iact(j))*b(iact(j))
         lik=lik+g
         if (abs(g).gt.600) then
            iprb=1
            return
         endif
         u=dexp(g)
         xb1=xb1+u
         do 88 j=1,nact
88   xb(iact(j))=xb(iact(j))+gp(iact(j))*u
         u=ktmp*u*(xb1/xb1o)
         do 19 j=1,nact
            s(iact(j))=s(iact(j))+gp(iact(j))
            u1=u*(gp(iact(j))-xb(iact(j))/xb1)
            do 19 j1=j,nact
               u2=gp(iact(j1))-xb(iact(j1))/xb1
               vt(iact(j),iact(j1))=vt(iact(j),iact(j1))+u1*u2
19   continue
         xb1o=xb1
         ktmp=1
         do j = 1, nact
            z(inct-ncens,iact(j))=gp(iact(j))
C           z1(inct-ncens,iact(j))=gp(iact(j))
         enddo
         inct=inct-1
26   continue
         if (15.le.13) then
C ****
      ncsi = inct - ncens
      ict = 1
C ****
         do 6 i=15,13
            call eval(g,gp,b,lci,i,1,sv,ce,xx,nrx,ncx,n,np,lc2,nact,iact,
&          isub,dsub)
CCC       call eval8n(gn,lci,i,1,sv,ce,xx,nrx,ncx,n,np,lc2,nact,iact,
CCC       &      isub,dsub)
c      g=0
c      do 841 j=1,nact
c 841  g=g+gp(iact(j))*b(iact(j))
         if (abs(g).gt.600) then
            iprb=1
            return
         endif
         u=dexp(g)
         xb1=xb1+u
         do 8 j=1,nact
8     xb(iact(j))=xb(iact(j))+gp(iact(j))*u
         u=ktmp*u*(xb1/xb1o)
         do 9 j=1,nact
            u1=u*(gp(iact(j))-xb(iact(j))/xb1)
            do 9 j1=j,nact
               u2=gp(iact(j1))-xb(iact(j1))/xb1
               vt(iact(j),iact(j1))=vt(iact(j),iact(j1))+u1*u2
9   continue
         xb1o=xb1
         ktmp=1

```

```

        do j = 1, nact
C ****
      z(ncsi+ict+nf,iact(j))=gp(iact(j))
C      z1(ncsi+ict+nf,iact(j))=gp(iact(j))
C ****
      enddo
      ict = ict + 1
6     continue
C ****
      inct = inct - ncens
C ****
      endif

c update score, likelihood, and inf:
c      if (xb1.le.0) write (6,118) xb1
c118    format ('xb1=',e12.4)
c110    format (6e12.4)
      lik=lik-nd(k)*dlog(xb1)
      do 18 j=1,nact
        s(iact(j))=s(iact(j))-nd(k)*xb(iact(j))/xb1
        do 20 j1=j,nact
20      v(iact(j),iact(j1))=v(iact(j),iact(j1))+nd(k)*
          &vt(iact(j),iact(j1))/xb1
      18  continue
      13=1-1
      10  continue
      98  continue
      101 format(i4,e20.10)
      nost=nostr(lc2)
      996 continue
c fill out symmetric information matrix:
      if (np.eq.1) go to 99
      do 56 j1=2,np
        do 56 j2=1,j1
56      v(j1,j2)=v(j2,j1)
      99  return
      end

      subroutine strat(istr,s,c,d,nostr,nstr,it1,t1,no,nv,nnob,
      &iuti,tduti)
c subroutine to sort data on strata variable, and determine obs index
c of strata limits, also sorts on survival times within strata
c on input s and c are the survival and censoring vars,
c d is a double precision array containing the covariate data
c and istr gives the column of d containing the strat variable
c nnob is the row dimension of d in the calling program.  Actually
c data matrix is no #obs x nv #vars.
c on output nstr is the # strata nostr(i) is the row # of the largest
c observation in strata i  (nostr must be at least nstr in lenght)
c t1 is double precision and it1 integer working vectors of length no
      double precision d(nnob,nv),t1(no),s(no),c(no)
      integer it1(no),nostr(1),iuti(no,2),tduti(no)
      if (istr.le.0) then
        nstr=1
        nostr(1)=no
      else
        do 214 i=1,no
          it1(i)=i
214      t1(i)=d(i,istr)
        call sortg(t1,it1,1,no)
        call dperm(s,it1,1,no,t1)
        call dperm(c,it1,1,no,t1)
        do 217 j=1,nv

```

```

        call dperm(d(1,j),it1,1,no,t1)
217      continue
      nstr=0
      do 220 i=2,no
      if (d(i,istr).gt.d(i-1,istr)) then
          nstr=nstr+1
          nostr(nstr)=i-1
      endif
220      continue
      nstr=nstr+1
      nostr(nstr)=no
      endif
c   sort on survival times within strata
      l1=1
      do 225 l1=1,nstr
      l2=nostr(l1)
      do 15 i=l1,l2
      it1(i)=i
15      continue
      call sortg(s,it1,l1,l2)
      call dperm(c,it1,l1,l2,t1)
      do 21 j=1,nv
      call dperm(d(1,j),it1,l1,l2,t1)
21      continue
      call ut(l1,l2,s,c,iuti,nostr(nstr+l1),tduti,no)
      call uft(d,s,c,l1,iuti,nostr(nstr+l1),tduti,no,nnob,nv)
      l1=l2+1
225      continue
      return
      end

      subroutine ut(n1,n2,t,c,iuti,nuti,tduti,no)
c subroutine to determine unique times & # obs at each time
c on input t is a vector of times, and c is a vector of failure
c indicators. only observations #'ed n1-n2 (min-max) in t are considered
c no is the row dimension of iuti in the calling program.
c on output nuti is the # of unique values in t (between t(n1) and t(n2))
c iuti(i,1) and iuti(i,2) are the lower and upper index (in t) of the
c ith unique time, and tduti is the # failures (c=1) at the ith
c unique time. t must be sorted in ascending order before calling ut.
c note that in this revised (8-15-90) version nuti is not the number
c of unique times, but rather the index such that the unique times
c are stored in indecies n1 to nuti
      double precision t(n2),c(n2),t1,t2
      integer iuti(no,2),tduti(1)
      l1=n1
      l2=n1-1
      t1=t(n1)
      iuti(n1,1)=n1
      t2=0
10      t2=t2+c(l1)
      l1=l1+1
      if (l1.gt.n2) go to 11
      if (t(l1).eq.t1) go to 10
11      l2=l2+1
      tduti(l2)=t2
      iuti(l2,2)=l1-1
      t2=0
      if (l1.le.n2) then
          t1=t(l1)
          go to 10
      endif
      nuti=l2

```

```

      do 20 i=n1+1,nuti
      iuti(i,1)=iuti(i-1,2)+1
20  continue
      return
      end

      subroutine uft(d,s,c,n1,iuti,nuti,tduti,no,nnob,nv)
c subroutine to take output from ut and compress so only has lines
c for failures. Also sorts data matrix d at each time so failures
c occur 1st
c nuti is the # unique times (output from ut) iuti gives the index range
c for the ith time interval and tduti the # failures at the ith unique
c time. d is the data matrix, nnob is the row dimension of d
c in the calling program, and nv the # cols in d being used.
c kc is the col of d for failure indicator
c On output d has been sorted within unique times so failures occur
c first. nuti is the # unique failure times, tduti is the # failures
c at each unique failure time, and iuti gives the index (row of d) for
c the obs >= the failure time and < the next failure time (the risk set
c increment) col 1 is the min and col 2 the max index.

      double precision d(nnob,nv),t1,c(1),s(1)
      integer iuti(no,2),tduti(nuti)
      do 10 k=n1,nuti
      if (tduti(k).gt.0) then
      nf=tduti(k)
      nn=iuti(k,2)-iuti(k,1)+1
      if (nn.gt.nf) then
      do 5 i=1,nf
      if (c(iuti(k,1)+i-1).le.0) then
6       if (c(iuti(k,1)+nn-1).ge..99) then
          t1=c(iuti(k,1)+i-1)
          c(iuti(k,1)+i-1)=c(iuti(k,1)+nn-1)
          c(iuti(k,1)+nn-1)=t1
          do 7 j=1,nv
          t1=d(iuti(k,1)+i-1,j)
          d(iuti(k,1)+i-1,j)=d(iuti(k,1)+nn-1,j)
          d(iuti(k,1)+nn-1,j)=t1
7       continue
          nn=nn-1
          if (nn.eq.nf) go to 8
        else
          nn=nn-1
          go to 6
        endif
      endif
5       continue
8       continue
      endif
10  continue
      nutd=n1-1
      do 20 k=n1,nuti
      if (tduti(k).gt.0) then
      nutd=nutd+1
      iuti(nutd,1)=iuti(k,1)
      tduti(nutd)=tduti(k)
      endif
20  continue
      do 25 k=n1+1,nutd
25  iuti(k-1,2)=iuti(k,1)-1
      if (nutd.ge.n1) iuti(nutd,2)=iuti(nuti,2)
      nuti=nutd
      return

```

```

    end

    subroutine tint(ntud,t,s,c,itud,no,iuti,nuti,n1)
c subroutine to determine rows of iuti such that observations fall
c in time intervals given by t
c ntud is the # of time intervals, t(i) is the lower limit on interval
c i+1, (that is, the intervals are of the form I(j)=[t(j-1),t(j)]).
c s is the vector of survival times and c the vector of failure
c indicators, no is the # of observations.
c n1 is the lower limit index for the current strata
c iuti gives the indexes (row #'s) of the data matrix d such that
c iuti(j,1) is the row # of a failure time and iuti(j,2) is the
c row # of the largest observation smaller than the next largest
c failure time (ie the risk set increment at the jth unique failure
c time). nuti is the # unique failure times, which is the # rows
c in iuti. (nuti and iuti probably created by uft)
c on output itud gives the row #'s of iuti that correspond to failure
c times in each of the time intervals. if there are no failures in
c a particular interval that row of itud will be -1 -1
    double precision s(no),c(no),t(ntud-1)
    integer itud(ntud,2),iuti(no,2)
    if (ntud.eq.1) then
        itud(1,1)=n1
        itud(1,2)=nuti
    else
        do 5 i=1,ntud
    5  itud(i,2)=0
        l1=1
        do 6 i=n1,nuti
    6  if (t(l1).gt.s(iuti(i,1))) then
            itud(l1,2)=itud(l1,2)+1
        else
            l1=l1+1
            if (l1.ge.ntud) go to 8
            go to 7
        endif
    7  continue
    8  l1=n1-1
        do 10 i=1,ntud-1
        if (itud(i,2).le.0) then
            itud(i,1)=-1
            itud(i,2)=-1
        else
            itud(i,1)=l1+1
            l1=l1+itud(i,2)
            itud(i,2)=l1
        endif
    10 continue
        if (l1.lt.nuti) then
            itud(ntud,1)=l1+1
            itud(ntud,2)=nuti
        else
            itud(ntud,1)=-1
            itud(ntud,2)=-1
        endif
        endif
    return
end

    subroutine tint4(ntud,s,c,t,no,nfp)
c subroutine to determine time intervals . The idea is to create ntud
c intervals with roughly equal #'s of failures. s (input) is the
c survival times and c (input) the failure indicator.

```

```

c Data must be sorted on s ntud is the # of intervals (input).
c On output nfp(i) is the # of failures in
c the ith interval. Ij is [t(j-1),t(j)). (t(0)=0 is not given)
c t(j) is chosen so there will be a failure at t(j)
c That is, t(j) is actually the smallest failure in the j+1 interval.
c On input no is the # obs (length of s and c).
    double precision s(no),c(no),t(ntud-1),t1
    integer nfp(ntud)
    nft=0
    do 10 i=1,no
10   nft=nft+c(i)
    l2=1
    t1=s(1)
    ntud1=ntud-1
    nfp(1)=0
    do 15 k=1,ntud1
c nfk is the target # failures for the kth interval
c continue advancing until total # >= nfk and find the
c next failure. that next failure is t(k)
    nfk=max(int(nft/float(ntud-k+1)+.5),1)
16   if (c(l2).eq.0) then
        l2=l2+1
        if (l2.gt.no) then
            ntud=k
c           write (6,100) ntud
c 100      format ('ntud reset in tint4 to', i5)
c           call intpr("ntud changed",12,ntud,1)
            go to 99
        else
            go to 16
        endif
    else
        if (s(l2).le.t1) then
            nfp(k)=nfp(k)+1
            l2=l2+1
            if (l2.gt.no) then
                ntud=k
c               write (6,100) ntud
c               call intpr("ntud changed",12,ntud,1)
                go to 99
            else
                go to 16
            endif
        else
            t1=s(l2)
            if (nfp(k).lt.nfk) then
                nfp(k)=nfp(k)+1
                l2=l2+1
                go to 16
            else
                t(k)=s(l2)
                nfp(k+1)=1
                l2=l2+1
                nft=nft-nfp(k)
            endif
        endif
    endif
15   continue
    nfp(ntud)=nft
c     call dblepr("knots",5,t,ntud-1)
c     call intpr(" ",1,nfp,ntud)
99   return
    end

```

```

        subroutine cholg(n,v,ch,nnvar,12)
c subroutine to calc cholesky decomposition of a pd matrix (v)
c on output ch will be upper triangular cholesky decomp
c v=t(ch)*ch
c nnvar and 12 are dimension of v and ch as specified in calling prog
c returns ch(1,1)=-1 if not pd
c ch and v can be the same matrix
c calls dpocog (linpack dpoco) which calls dpofag dscalg daxpyg ddotg dasumg
        double precision v(nnvar,n),ch(12,n),rcond,z(200)
        if (n.gt.200) go to 520
        do 10 i=1,n
        do 10 j=i,n
10      ch(i,j)=v(i,j)
        call dpocog(ch,12,n,rcond,z,info)
c         call dblepr("rcond",5,rcond,1)
        if (rcond.lt.1.d-15.or.info.gt.0) go to 520
        do 40 i=2,n
        do 40 j=1,i-1
40      ch(i,j)=0
        return
520    continue
        ch(1,1)=-1
        return
        end

c*** from netlib, Sat Oct 12 09:38:00 EDT 1991 ***
c from linpack
c these are standard linpack routines. Names changed so
c don't overwrite S routines of same names, since don't know if
c they are identical.
        subroutine dpocog(a,lda,n,rcond,z,info)
        integer lda,n,info
        double precision a(lda,1),z(1)
        double precision rcond
c
c dpoco factors a double precision symmetric positive definite
c matrix and estimates the condition of the matrix.
c
c if rcond is not needed, dpofa is slightly faster.
c to solve a*x = b , follow dpoco by dposl.
c to compute inverse(a)*c , follow dpoco by dposl.
c to compute determinant(a) , follow dpoco by dpodi.
c to compute inverse(a) , follow dpoco by dpodi.
c
c on entry
c
c     a      double precision(lda, n)
c             the symmetric matrix to be factored. only the
c             diagonal and upper triangle are used.
c
c     lda    integer
c             the leading dimension of the array a .
c
c     n      integer
c             the order of the matrix a .
c
c on return
c
c     a      an upper triangular matrix r so that a = trans(r)*r
c             where trans(r) is the transpose.
c             the strict lower triangle is unaltered.
c             if info .ne. 0 , the factorization is not complete.

```

```

c
c      rcond    double precision
c      an estimate of the reciprocal condition of a .
c      for the system a*x = b , relative perturbations
c      in a and b of size epsilon may cause
c      relative perturbations in x of size epsilon/rcond .
c      if rcond is so small that the logical expression
c          1.0 + rcond .eq. 1.0
c      is true, then a may be singular to working
c      precision. in particular, rcond is zero if
c      exact singularity is detected or the estimate
c      underflows. if info .ne. 0 , rcond is unchanged.
c
c      z      double precision(n)
c      a work vector whose contents are usually unimportant.
c      if a is close to a singular matrix, then z is
c      an approximate null vector in the sense that
c      norm(a*z) = rcond*norm(a)*norm(z) .
c      if info .ne. 0 , z is unchanged.
c
c      info   integer
c      = 0 for normal return.
c      = k signals an error condition. the leading minor
c      of order k is not positive definite.
c
c      linpack. this version dated 08/14/78 .
c      cleve moler, university of new mexico, argonne national lab.
c
c      subroutines and functions
c
c      linpack dpofa
c      blas daxpy,ddot,dscal,dasum
c      fortran dabs,dmax1,dreal,dsign
c
c      internal variables
c
c      double precision ddotg,ek,t,wk,wkm
c      double precision anorm,s,dasumg,sm,ynorm
c      integer i,j,jm1,k,kb,kp1
c
c
c      find norm of a using only upper half
c
c      do 30 j = 1, n
c          z(j) = dasumg(j,a(1,j),1)
c          jm1 = j - 1
c          if (jm1 .lt. 1) go to 20
c          do 10 i = 1, jm1
c              z(i) = z(i) + dabs(a(i,j))
c 10      continue
c 20      continue
c 30      continue
c          anorm = 0.0d0
c          do 40 j = 1, n
c              anorm = dmax1(anorm,z(j))
c 40      continue
c
c      factor
c
c      call dpofag(a,lda,n,info)
c      if (info .ne. 0) go to 180
c
c      rcond = 1/(norm(a)*(estimate of norm(inverse(a)))) .

```

```

c      estimate = norm(z)/norm(y) where a*z = y and a*y = e .
c      the components of e are chosen to cause maximum local
c      growth in the elements of w where trans(r)*w = e .
c      the vectors are frequently rescaled to avoid overflow.
c
c      solve trans(r)*w = e
c
c      ek = 1.0d0
c      do 50 j = 1, n
c          z(j) = 0.0d0
50    continue
c      do 110 k = 1, n
c          if (z(k) .ne. 0.0d0) ek = dsign(ek,-z(k))
c          if (dabs(ek-z(k)) .le. a(k,k)) go to 60
c              s = a(k,k)/dabs(ek-z(k))
c              call dscaln(n,s,z,1)
c              ek = s*ek
60    continue
c      wk = ek - z(k)
c      wkm = -ek - z(k)
c      s = dabs(wk)
c      sm = dabs(wkm)
c      wk = wk/a(k,k)
c      wkm = wkm/a(k,k)
c      kp1 = k + 1
c      if (kp1 .gt. n) go to 100
c          do 70 j = kp1, n
c              sm = sm + dabs(z(j)+wkm*a(k,j))
c              z(j) = z(j) + wk*a(k,j)
c              s = s + dabs(z(j))
70    continue
c      if (s .ge. sm) go to 90
c          t = wkm - wk
c          wk = wkm
c          do 80 j = kp1, n
c              z(j) = z(j) + t*a(k,j)
80    continue
90    continue
100   continue
c      z(k) = wk
110   continue
c      s = 1.0d0/dasumg(n,z,1)
c      call dscaln(n,s,z,1)
c
c      solve r*y = w
c
c      do 130 kb = 1, n
c          k = n + 1 - kb
c          if (dabs(z(k)) .le. a(k,k)) go to 120
c              s = a(k,k)/dabs(z(k))
c              call dscaln(n,s,z,1)
120   continue
c          z(k) = z(k)/a(k,k)
c          t = -z(k)
c          call daxpyg(k-1,t,a(1,k),1,z(1),1)
130   continue
c          s = 1.0d0/dasumg(n,z,1)
c          call dscaln(n,s,z,1)
c
c      ynorm = 1.0d0
c
c      solve trans(r)*v = y
c

```

```

do 150 k = 1, n
  z(k) = z(k) - ddotg(k-1,a(1,k),1,z(1),1)
  if (dabs(z(k)) .le. a(k,k)) go to 140
    s = a(k,k)/dabs(z(k))
    call dscalg(n,s,z,1)
    ynorm = s*ynorm
140  continue
  z(k) = z(k)/a(k,k)
150  continue
  s = 1.0d0/dasumg(n,z,1)
  call dscalg(n,s,z,1)
  ynorm = s*ynorm
c
c      solve r*z = v
c
  do 170 kb = 1, n
    k = n + 1 - kb
    if (dabs(z(k)) .le. a(k,k)) go to 160
      s = a(k,k)/dabs(z(k))
      call dscalg(n,s,z,1)
      ynorm = s*ynorm
160  continue
  z(k) = z(k)/a(k,k)
  t = -z(k)
  call daxpyg(k-1,t,a(1,k),1,z(1),1)
170  continue
c      make znorm = 1.0
  s = 1.0d0/dasumg(n,z,1)
  call dscalg(n,s,z,1)
  ynorm = s*ynorm
c
  if (anorm .ne. 0.0d0) rcond = ynorm/anorm
  if (anorm .eq. 0.0d0) rcond = 0.0d0
180 continue
  return
  end

subroutine dpofag(a,lda,n,info)
integer lda,n,info
double precision a(lda,1)
c
c      dpofa factors a double precision symmetric positive definite
c      matrix.
c
c      dpofa is usually called by dpoco, but it can be called
c      directly with a saving in time if rcond is not needed.
c      (time for dpoco) = (1 + 18/n)*(time for dpofa) .
c
c      on entry
c
c          a      double precision(lda, n)
c                  the symmetric matrix to be factored. only the
c                  diagonal and upper triangle are used.
c
c          lda     integer
c                  the leading dimension of the array a .
c
c          n      integer
c                  the order of the matrix a .
c
c      on return
c
c          a      an upper triangular matrix r so that a = trans(r)*r

```

```

c           where trans(r)  is the transpose.
c           the strict lower triangle is unaltered.
c           if info .ne. 0 , the factorization is not complete.
c
c   info    integer
c           = 0  for normal return.
c           = k  signals an error condition.  the leading minor
c                 of order  k  is not positive definite.
c
c   linpack.  this version dated 08/14/78 .
c   cleve moler, university of new mexico, argonne national lab.
c
c   subroutines and functions
c
c   blas ddot
c   fortran dsqrt
c
c   internal variables
c
c   double precision ddotg,t
c   double precision s
c   integer j,jm1,k
c begin block with ...exits to 40
c
c
do 30 j = 1, n
  info = j
  s = 0.0d0
  jm1 = j - 1
  if (jm1 .lt. 1) go to 20
  do 10 k = 1, jm1
    t = a(k,j) - ddotg(k-1,a(1,k),1,a(1,j),1)
    t = t/a(k,k)
    a(k,j) = t
    s = s + t*t
10      continue
20      continue
      s = a(j,j) - s
c .....exit
      if (s .le. 0.0d0) go to 40
      a(j,j) = dsqrt(s)
30      continue
      info = 0
40 continue
return
end

subroutine dscalg(n,da,dx,incx)
c
c   scales a vector by a constant.
c   uses unrolled loops for increment equal to one.
c   jack dongarra, linpack, 3/11/78.
c
c   double precision da,dx(1)
c   integer i,incx,m,mp1,n,nincx
c
c   if(n.le.0)return
c   if(incx.eq.1)go to 20
c
c       code for increment not equal to 1
c
nincx = n*incx
do 10 i = 1,nincx,incx

```

```

        dx(i) = da*dx(i)
10 continue
        return

c
c      code for increment equal to 1
c
c
c      clean-up loop
c

20 m = mod(n,5)
    if( m .eq. 0 ) go to 40
    do 30 i = 1,m
        dx(i) = da*dx(i)
30 continue
    if( n .lt. 5 ) return
40 mp1 = m + 1
    do 50 i = mp1,n,5
        dx(i) = da*dx(i)
        dx(i + 1) = da*dx(i + 1)
        dx(i + 2) = da*dx(i + 2)
        dx(i + 3) = da*dx(i + 3)
        dx(i + 4) = da*dx(i + 4)
50 continue
    return
end

double precision function dasumg(n,dx,incx)
c
c      takes the sum of the absolute values.
c      jack dongarra, linpack, 3/11/78.
c
c      double precision dx(1),dtemp
c      integer i,incx,m,mp1,n,nincx
c
c      dasumg = 0.0d0
c      dtemp = 0.0d0
c      if(n.le.0)return
c      if(incx.eq.1)go to 20
c
c      code for increment not equal to 1
c
nincx = n*incx
do 10 i = 1,nincx,incx
    dtemp = dtemp + dabs(dx(i))
10 continue
dasumg = dtemp
return

c
c      code for increment equal to 1
c
c
c      clean-up loop
c

20 m = mod(n,6)
    if( m .eq. 0 ) go to 40
    do 30 i = 1,m
        dtemp = dtemp + dabs(dx(i))
30 continue
    if( n .lt. 6 ) go to 60
40 mp1 = m + 1
    do 50 i = mp1,n,6
        dtemp = dtemp + dabs(dx(i)) + dabs(dx(i + 1)) + dabs(dx(i + 2))
*      + dabs(dx(i + 3)) + dabs(dx(i + 4)) + dabs(dx(i + 5))

```

```

50 continue
60 dasumg = dtemp
      return
      end

      subroutine daxpyg(n,da,dx,incx,dy,incy)
c
c      constant times a vector plus a vector.
c      uses unrolled loops for increments equal to one.
c      jack dongarra, linpack, 3/11/78.
c
      double precision dx(1),dy(1),da
      integer i,incx,incy,ix,iy,m,mp1,n
c
      if(n.le.0)return
      if (da .eq. 0.0d0) return
      if(incx.eq.1.and.incy.eq.1)go to 20
c
c      code for unequal increments or equal increments
c      not equal to 1
c
      ix = 1
      iy = 1
      if(incx.lt.0)ix = (-n+1)*incx + 1
      if(incy.lt.0)iy = (-n+1)*incy + 1
      do 10 i = 1,n
         dy(iy) = dy(iy) + da*dx(ix)
         ix = ix + incx
         iy = iy + incy
10  continue
      return
c
c      code for both increments equal to 1
c
c
c      clean-up loop
c
20 m = mod(n,4)
      if( m .eq. 0 ) go to 40
      do 30 i = 1,m
         dy(i) = dy(i) + da*dx(i)
30  continue
      if( n .lt. 4 ) return
40 mp1 = m + 1
      do 50 i = mp1,n,4
         dy(i) = dy(i) + da*dx(i)
         dy(i + 1) = dy(i + 1) + da*dx(i + 1)
         dy(i + 2) = dy(i + 2) + da*dx(i + 2)
         dy(i + 3) = dy(i + 3) + da*dx(i + 3)
50  continue
      return
      end

      double precision function ddotg(n,dx,incx,dy,incy)
c
c      forms the dot product of two vectors.
c      uses unrolled loops for increments equal to one.
c      jack dongarra, linpack, 3/11/78.
c
      double precision dx(1),dy(1),dtemp
      integer i,incx,incy,ix,iy,m,mp1,n
c
      ddotg = 0.0d0

```

```

dtemp = 0.0d0
if(n.le.0) return
if(incx.eq.1.and.incy.eq.1) go to 20
c
c      code for unequal increments or equal increments
c      not equal to 1
c
ix = 1
iy = 1
if(incx.lt.0) ix = (-n+1)*incx + 1
if(incy.lt.0) iy = (-n+1)*incy + 1
do 10 i = 1,n
    dtemp = dtemp + dx(ix)*dy(iy)
    ix = ix + incx
    iy = iy + incy
10 continue
ddotg = dtemp
return

c
c      code for both increments equal to 1
c
c
c      clean-up loop
c
20 m = mod(n,5)
if( m .eq. 0 ) go to 40
do 30 i = 1,m
    dtemp = dtemp + dx(i)*dy(i)
30 continue
if( n .lt. 5 ) go to 60
40 mp1 = m + 1
do 50 i = mp1,n,5
    dtemp = dtemp + dx(i)*dy(i) + dx(i + 1)*dy(i + 1) +
*   dx(i + 2)*dy(i + 2) + dx(i + 3)*dy(i + 3) + dx(i + 4)*dy(i + 4)
50 continue
60 ddotg = dtemp
return
end

subroutine solve(n,ch,b,x,nnvar)
double precision ch(nnvar,n),b(n),x(n)
c ch is upper triang chol decomp, b (input) is the rhs, x (output)
c is the solution
c
c first solve t(ch)*x=b
call slv1(n,ch,b,x,nnvar)
c then solve ch*x=xold
call slv2(n,ch,x,x,nnvar)
return
end

subroutine sortg(v,a,ii,jj)
c
c puts into a the permutation vector which sorts v into
c increasing order. only elements from ii to jj are considered.
c arrays iu(k) and il(k) permit sorting up to 2**k+1-1 elements
c v is returned sorted
c this is a modification of cacm algorithm #347 by r. c. singleton,
c which is a modified hoare quicksort.
c
c on input a has integers ii to jj in components ii to jj
c
dimension a(jj),v(1),iu(20),il(20)

```

```

integer t,tt
integer a
double precision v,vt,vtt
m=1
i=ii
j=jj
10 if (i.ge.j) go to 80
20 k=i
ij=(j+i)/2
t=a(ij)
vt=v(ij)
if (v(i).le.vt) go to 30
a(ij)=a(i)
a(i)=t
t=a(ij)
v(ij)=v(i)
v(i)=vt
vt=v(ij)
30 l=j
if (v(j).ge.vt) go to 50
a(ij)=a(j)
a(j)=t
t=a(ij)
v(ij)=v(j)
v(j)=vt
vt=v(ij)
if (v(i).le.vt) go to 50
a(ij)=a(i)
a(i)=t
t=a(ij)
v(ij)=v(i)
v(i)=vt
vt=v(ij)
go to 50
40 a(l)=a(k)
a(k)=tt
v(l)=v(k)
v(k)=vtt
50 l=l-1
if (v(l).gt.vt) go to 50
tt=a(l)
vtt=v(l)
60 k=k+1
if (v(k).lt.vt) go to 60
if (k.le.l) go to 40
if (l-i.lt.j-k) go to 70
il(m)=i
iu(m)=l
i=k
m=m+1
go to 90
70 il(m)=k
iu(m)=j
j=1
m=m+1
go to 90
80 m=m-1
if (m.eq.0) return
i=il(m)
j=iu(m)
90 if (j-i.gt.10) go to 20
if (i.eq.ii) go to 10
i=i-1

```

```

100  i=i+1
     if (i.eq.j) go to 80
     t=a(i+1)
     vt=v(i+1)
     if (v(i).le.vt) go to 100
     k=i
110  a(k+1)=a(k)
     v(k+1)=v(k)
     k=k-1
     if (vt.lt.v(k)) go to 110
     a(k+1)=t
     v(k+1)=vt
     go to 100
     end

     subroutine dperm(a,ia,nl,nu,work)
     double precision a(nu),work(nu)
     integer ia(nu)
     do 10 i=nl,nu
10    work(i)=a(i)
     do 11 i=nl,nu
11    a(i)=work(ia(i))
     return
     end

     subroutine adisw(n1,n2,n,v,v2,lb)
c calcs adjusted information.  v is the full information matrix
c (v is nxn), n1 and n2 are min and max index of the block of interest
c lb is actual row dimension of v and v2.  final result will be in
c elements 1:(n2-n1+1)*1:(n2-n1+1) of v2.  v2 must also be nxn.
     double precision v(lb,n),v2(lb,n)
     np=n2-n1+1
     if (np.ge.n) then
     do 10 i=1,n
     do 10 j=1,n
10    v2(i,j)=v(i,j)
     return
     endif
     if (n1.le.1) then
     call gsweep(n1,n2,n2+1,n,v,v2,lb)
     do 12 k=n2+2,n
12    call gsweep(n1,n2,k,n,v2,v2,lb)
     else
     call gsweep(n1,n2,1,n,v,v2,lb)
     do 14 k=2,n1-1
14    call gsweep(n1,n2,k,n,v2,v2,lb)
     do 16 k=n2+1,n
16    call gsweep(n1,n2,k,n,v2,v2,lb)
     endif
     do 20 i=1,np
     i2=i+n1-1
     v2(i,i)=v2(i2,i2)
     do 21 j=i+1,np
     j2=j+n1-1
     v2(j,i)=v2(i2,j2)
21    v2(i,j)=v2(i2,j2)
20    continue
     return
     end

     subroutine gsweep(n1,n2,l,n,v,v2,lb)
c note that this is intended for use with adjusted inf routine:
c it is assumed that l is not in [n1,n2]

```

```

c v is the matrix to be swept on the lth component, result put in v2.
c v and v2 can be the same. assumes symmetric matrices, only upper
c triangular part is used. assumes that sweeping is proceeding in
c order, so that only elements from l to n (+[n1-n2]) need to be swept.
c

        double precision v(lb,n),v2(lb,n)
        v2(1,1)=-1/v(1,1)
        do 10 i=l+1,n
10      v2(1,i)=v(1,i)*v2(1,1)
        do 11 i=1,l-1
11      v2(i,1)=v(i,1)*v2(1,1)
        do 12 i=l+1,n
        do 13 j=1,l-1
13      v2(j,i)=v(j,i)+v2(j,1)*v2(1,i)/v2(1,1)
        do 14 j=i,n
14      v2(i,j)=v(i,j)+v2(1,i)*v2(1,j)/v2(1,1)
12      continue
        if (l.ge.n1) then
        do 22 i=n1,n2
        do 22 j=i,n2
22      v2(i,j)=v(i,j)+v2(i,1)*v2(j,1)/v2(1,1)
        endif
        return
        end

        SUBROUTINE eigen(A,N,NP,D,E,iopt,iflg)
        implicit double precision (a - h, o - z)
c subroutine to calculate eigenvalues and eigenvectors
c based on Numerical Recipies routines
c on input a is a symmetric matrix, nxn, but dimensioned npxn
c in the calling program.
c a is destroyed, in this routine, but the call to tql2 writes
c the eigenvectors to a and the eigen values to d. The original
c matrix can then be calculated as ADA', where D is the diag matrix
c with the eigenvalues on the diagonal. Note that the eigenvalues
c are not sorted.
c iflg=1 means the iteration in tql2 failed to converge.
c this routine reduces a to tridiag form. tql2 actually calculates
c eigenvalues and eigenvectors.
c iopt<=0 only eigenvalues calculated
c iopt>0 eigenvalues & vectors calculated
        DIMENSION A(NP,N),D(1),E(1)
        IF(N.GT.1) THEN
        DO 18 I=N,2,-1
          L=I-1
          H=0.d0
          SCALE=0.d0
          IF(L.GT.1)THEN
          DO 11 K=1,L
            SCALE=SCALE+ABS(A(I,K))
11        CONTINUE
          IF(SCALE.EQ.0.d0)THEN
            E(I)=A(I,L)
          ELSE
          DO 12 K=1,L
            A(I,K)=A(I,K)/SCALE
            H=H+A(I,K)**2
12        CONTINUE
          F=A(I,L)
          G=-SIGN(SQRT(H),F)
          E(I)=SCALE*G
          H=H-F*G
          A(I,L)=F-G

```

```

F=0.d0
DO 15 J=1,L
    A(J,I)=A(I,J)/H
    G=0.d0
    DO 13 K=1,J
        G=G+A(J,K)*A(I,K)
13      CONTINUE
        IF(L.GT.J)THEN
            DO 14 K=J+1,L
                G=G+A(K,J)*A(I,K)
14      CONTINUE
        ENDIF
        E(J)=G/H
        F=F+E(J)*A(I,J)
15      CONTINUE
        HH=F/(H+H)
        DO 17 J=1,L
            F=A(I,J)
            G=E(J)-HH*F
            E(J)=G
            DO 16 K=1,J
                A(J,K)=A(J,K)-F*E(K)-G*A(I,K)
16      CONTINUE
17      CONTINUE
        ENDIF
    ELSE
        E(I)=A(I,L)
    ENDIF
    D(I)=H
18      CONTINUE
ENDIF
if (iopt.gt.0) then
    D(1)=0.d0
    E(1)=0.d0
    DO 23 I=1,N
        L=I-1
        IF(D(I).NE.0.d0)THEN
            DO 21 J=1,L
                G=0.d0
                DO 19 K=1,L
                    G=G+A(I,K)*A(K,J)
19      CONTINUE
                DO 20 K=1,L
                    A(K,J)=A(K,J)-G*A(K,I)
20      CONTINUE
21      CONTINUE
        ENDIF
        D(I)=A(I,I)
        A(I,I)=1.d0
        IF(L.GE.1)THEN
            DO 22 J=1,L
                A(I,J)=0.d0
                A(J,I)=0.d0
22      CONTINUE
        ENDIF
23      CONTINUE
else
    E(1)=0.d0
    DO 25 I=1,N
        D(I)=A(I,I)
25      CONTINUE
endif
call TQL2g(D,E,N,np,a,iopt,iflg)

```

```

RETURN
END

SUBROUTINE TQL2g(D,E,N,np,z,iopt,iflg)
c subroutine for finding eigenvalues of a tridiagonal matrix
c symmetric matrix reduced to tridiagonal form with tred2 (d and e
c are output from tred2). n is the length of d and e and portion of
c z that is used. np is actual dim of z in calling program.
c On output d contains the eigenvalues. iflg=1 if
c iteration failed to converge.
c if iopt>0 then on output z contains the eigenvectors
implicit double precision (a - h, o - z)
DIMENSION D(1),E(1),Z(NP,N)
iflg=0
IF (N.GT.1) THEN
    DO 11 I=2,N
        E(I-1)=E(I)
11    CONTINUE
        E(N)=0.d0
        DO 15 L=1,N
            ITER=0
1      DO 12 M=L,N-1
            DD=ABS(D(M))+ABS(D(M+1))
            IF (ABS(E(M))+DD.EQ.DD) GO TO 2
12    CONTINUE
        M=N
2      IF (M.NE.L)THEN
            IF(ITER.GE.30) then
                iflg=1
                go to 15
            endif
            ITER=ITER+1
            G=(D(L+1)-D(L))/(2.d0*E(L))
            R=SQRT(G**2+1.d0)
            G=D(M)-D(L)+E(L)/(G+SIGN(R,G))
            S=1.d0
            C=1.d0
            P=0.d0
            DO 14 I=M-1,L,-1
                F=S*E(I)
                B=C*E(I)
                IF (ABS(F).GE.ABS(G))THEN
                    C=G/F
                    R=SQRT(C**2+1.d0)
                    E(I+1)=F*R
                    S=1.d0/R
                    C=C*S
                ELSE
                    S=F/G
                    R=SQRT(S**2+1.d0)
                    E(I+1)=G*R
                    C=1.d0/R
                    S=S*C
                ENDIF
                G=D(I+1)-P
                R=(D(I)-G)*S+2.d0*C*B
                P=S*R
                D(I+1)=G+P
                G=C*R-B
                if (iopt.gt.0) then
                    DO 13 K=1,N
                        F=Z(K,I+1)
                        Z(K,I+1)=S*Z(K,I)+C*F
13            CONTINUE
                endif
            ENDIF
        ENDIF
    ENDIF
END

```

```

          Z(K,I)=C*Z(K,I)-S*F
13      CONTINUE
      endif
14      CONTINUE
      D(L)=D(L)-P
      E(L)=G
      E(M)=0.d0
      GO TO 1
      ENDIF
15      CONTINUE
      ENDIF
      RETURN
      END

      subroutine pdi(n,v,ch,nnvar,12)
c subroutine to invert a pd matrix. v=input matrix
c on output ch will be inverse
c returns ch(1,1)=-1 if not pd
      double precision v(nnvar,n),ch(12,n),tmp
      if (v(1,1).le.0) go to 520
      if (n.eq.1) then
        ch(1,1)=1/v(1,1)
        return
      endif
      call cholg(n,v,ch,nnvar,12)
      if (ch(1,1).le.0) go to 520
      n1=n-1
      ch(n,n)=1/ch(n,n)
      do 40 i=n1,1,-1
        ch(i,i)=1/ch(i,i)
        i1=i+1
      do 41 j=i1,n
41    ch(i,j)=ch(i,j)*ch(i,i)
40    continue
      do 45 j=n,2,-1
        j2=j-1
      do 46 k=j2,1,-1
        do 47 j1=j,n
47    ch(j1,k)=ch(j1,k)-ch(k,j)*ch(j1,j)
46    continue
45    continue
      do 50 i=1,n
      do 52 j=i,n
        tmp=0
      do 53 k=j,n
53    tmp=tmp+ch(k,i)*ch(k,j)
      ch(i,j)=tmp
52    continue
50    continue
      do 55 i=2,n
        i1=i-1
      do 56 j=1,i1
56    ch(i,j)=ch(j,i)
55    continue
      return
520  continue
c 520  write (6,101)
101  format('matrix not pd in pdi')
      ch(1,1)=-1
      return
      end

      subroutine pdi2(n,v,ch,nnvar,nv2)

```

```

c subroutine to factor a pd matrix, and invert the factorization
c . v=input matrix
c on output ch will be inverse square root of v
c returns ch(1,1)=-1 if not pd
    double precision v(nnvar,n),ch(nnvar,n)
    if (v(1,1).le.0) go to 520
    if (n.eq.1) then
        ch(1,1)=1/v(1,1)
        return
    endif
    call cholg(n,v,ch,nnvar,nv2)
    if (ch(1,1).le.0) go to 520
    n1=n-1
    ch(n,n)=1/ch(n,n)
    do 40 i=n1,1,-1
        ch(i,i)=1/ch(i,i)
        i1=i+1
        do 41 j=i1,n
        41 ch(i,j)=ch(i,j)*ch(i,i)
    40 continue
        do 45 j=n,2,-1
        j2=j-1
        do 46 k=j2,1,-1
        do 47 j1=j,n
        47 ch(j1,k)=ch(j1,k)-ch(k,j)*ch(j1,j)
    46 continue
    45 continue
        do 55 i=2,n
        i1=i-1
        do 56 j=1,i1
            ch(j,i)=ch(i,j)
    56 ch(i,j)=0
    55 continue
    return
520 continue
c write (6,101)
101 format('matrix not pd')
    ch(1,1)=-1
    return
end

subroutine slv1(n,ch,b,x,nnvar)
    double precision ch(nnvar,n),b(n),x(n)
c ch is upper triang chol decomp, b (input) is the rhs, x (output)
c is the solution
c this routine solves t(ch)*x=b
    x(1)=b(1)/ch(1,1)
    if (n.gt.1) then
        do 10 i=2,n
        l1=i-1
        x(i)=b(i)
        do 11 j=1,l1
        11 x(i)=x(i)-ch(j,i)*x(j)
        x(i)=x(i)/ch(i,i)
    10 continue
    endif
    return
end

subroutine slv2(n,ch,b,x,nnvar)
    double precision ch(nnvar,n),b(n),x(n)
c ch is upper triang chol decomp, b (input) is the rhs, x (output)
c is the solution

```

```

c  this routine solves ch*x=b
    x(n)=b(n)/ch(n,n)
    if (n.gt.1) then
        n1=n-1
        do 12 i=n1,1,-1
            l1=i+1
            x(i)=b(i)
            do 13 j=l1,n
                13 x(i)=x(i)-ch(i,j)*x(j)
                x(i)=x(i)/ch(i,i)
            12 continue
            endif
            return
        end
        subroutine extrct(v,np,lv,11,12,a,la)
        double precision v(lv,np),a(la,1)
c v input: np*np matrix
c 11 & 12: rows and columns from 1 to 11 and 12 to np are
c extracted from v and stored in a (a and v can be the same
c matrix.
        do 10 i=1,11
        do 10 j=1,11
    10 a(i,j)=v(i,j)
        k=11
        do 20 j=12,np
        k=k+1
        do 20 i=1,11
            a(i,k)=v(i,j)
            a(k,i)=v(j,i)
    20 continue
        k=12-11-1
        do 30 i=12,np
        do 30 j=12,np
    30 a(i-k,j-k)=v(i,j)
        return
    end

    subroutine matop5(vv,np,vc,n2,11,12,v5,n4,v6,v7,16)
    double precision vv(np,np),vc(n2,n2),v5(16,n4),v6(16,n4),
    &v7(n4,n4)
c 16 is the actual row dim of v5 and v6 in the calling program
c 11 to 12 are the spline terms, vv the unpenalized inf
c n4=12-11+1,n2=np-n4, vc is the ut chol decomp of v^star_beta,beta,
c where v^star is the pen inf, but without the pen applied to 11 to 12,
c and 'beta' means the parameters not in 11 to 12. v5,v6, and v7 are
c as dimensioned. On output v5 contains the appropriate terms for
c var_thetatheta|beta, & v6 the appropriate terms for v^star_thetatheta
c |beta
        do 10 i=1,11-1
        do 10 j=11,12
    10 v5(i,j-11+1)=vv(i,j)
        k=12-11+1
        do 20 i=12+1,np
        do 20 j=11,12
    20 v5(i-k,j-11+1)=vv(i,j)
        do 30 i=1,n4
            call solve(n2,vc,v5(1,i),v6(1,i),n2)
    30 continue
c calc v7=v^star_theta,theta|beta
        do 11 i=1,n4
        do 11 j=i,n4
            v7(i,j)=vv(i+11-1,j+11-1)
        do 13 k=1,n2

```

```

13 v7(i,j)=v7(i,j)-v5(k,i)*v6(k,j)
      v7(j,i)=v7(i,j)
11 continue
c set vc=v_beta,beta
      call extrct(vv,np,np,11-1,12+1,vc,n2)
c and set v5=(approp terms)-t(v6)*vc*v6
      do 22 i=1,n4
      do 22 j=i,n4
         v5(i,j)=2*v7(i,j)-vv(i+11-1,j+11-1)
         do 23 ii=1,n2
         do 23 jj=1,n2
23   v5(i,j)=v5(i,j)+v6(ii,i)*v6(jj,j)*vc(ii,jj)
         v5(j,i)=v5(i,j)
22   continue
         do 25 i=1,n4
         do 25 j=1,n4
25   v6(i,j)=v7(i,j)
      return
      end

      subroutine pencb(nk,wk,penm,mr)
c nk is # knots, wk is augmented knot seq.
c This subroutine calculates the penalty matrix for integrated squared
c second derivative penalty. The first col of penm will have the main
c diagonal, the 2nd col the next diag ... the 4th col the 3rd diag from
c the main.
c mr is the actual row dim of penm
      double precision wk(-2:nk+3),penm(mr,4),t1(4),tu(4),a(4),b(4)
      double precision t1,t2,t3,t4
      do 10 i=1,nk+4
      do 10 j=1,4
10   penm(i,j)=0
      t1=6/((wk(1)-wk(-2))*(wk(1)-wk(-1)))
      t1(2)=-6/((wk(1)-wk(-1))*(1/(wk(1)-wk(-2))+1/(wk(2)-wk(-1)))
      t1(3)=6/((wk(2)-wk(-1))*(wk(1)-wk(-1)))
      t1(4)=0
      do 20 l=1,nk+1
c t1 has the value of the 2nd deriv at the lower endpoint of the lth
c interval, tu at the upper (2nd derivs of basis fcns are linear on
c each interval
      t1=wk(l)-wk(l-1)
      t4=wk(l)+wk(l-1)
      t2=t1*t4/2
      t3=t1*(t4*t4-wk(l)*wk(l-1))/3
c note that t3 is (wk(l)^3-wk(l-1)^3)/3
      tu(1)=0
      tu(2)=6/((wk(l+1)-wk(l-2))*(wk(l+1)-wk(l-1)))
      tu(3)=-6/((wk(l+1)-wk(l-1))*(1/(wk(l+1)-wk(l-2))+
      &1/(wk(l+2)-wk(l-1)))
      tu(4)=6/((wk(l+2)-wk(l-1))*(wk(l+1)-wk(l-1)))
c calc slopes and intercepts on interval l:
      do 21 j=1,4
      b(j)=(tu(j)-t1(j))/(wk(l)-wk(l-1))
      a(j)=tu(j)-b(j)*wk(l)
21   continue
      do 22 j=1,4
      l1=l+j-1
      j2=j-1
      do 22 k=j,4
      penm(l1,k-j2)=penm(l1,k-j2)+2*(a(j)*a(k)*t1+(a(j)*b(k) +
      &a(k)*b(j))*t2+b(j)*b(k)*t3)
22   continue
      t1(1)=tu(2)

```

```

t1(2)=tu(3)
t1(3)=tu(4)
t1(4)=0
20 continue
return
end

subroutine penh(nb,penm,mr,md,h,nmin,ih,alpha)
c penm is the penalty matrix, nb the actual
c # basis functions used, h the neg sec deriv matrix
c nmin the entry in h where spline params begin. alpha*p
c is added to the submatrix bounded by h(nm,nm) and h(nm+nb-1,
c nm+nb-1)
c mr is the actual row dimension of penm, md is the # nonzero
c diagonals in the penalty matrix
    double precision penm(mr,md),h(ih,ih),alpha,u
    do 10 i=1,nb
    i1=nmin+i-1
    h(i1,i1)=h(i1,i1)+alpha*penm(i,1)
    if (i.lt.nb) then
        m2=min(nb,i+md-1)
        l=1
        do 12 j=i+1,m2
        j1=nmin+j-1
        l=l+1
        u=alpha*penm(i,l)
        h(i1,j1)=h(i1,j1)+u
        h(j1,i1)=h(j1,i1)+u
12    continue
    endif
10    continue
    return
end

subroutine gdf(ntud3,vv,v3,var,vvi,lb,df,alpha,v4,mindf,md,mr,
&iopt)
c change made 10-26-91 iopt returns 0 if ok, 1 if singular or other
c problems
c change made 5-5-91, fast alg no longer an option (was found to be
c unstable, iopt not really used anymore, but left in for compatibility
c with calling programs)
c change made 3-21-91, must input a starting value for alpha:
c ntud3 is the # parameters,
c vv is the penalty matrix (mr*md), v3 is the unpenalized -2nd deriv,
c var is the actual var matrix (both have row dim lb). var and v3 can be
c the same matrix (vv v3 and var are not changed)
c v4 is a working matrix with dim ntud3*ntud3.
c vvi a vector of length ntud3
c df is the target df (input) on output df=attained df,
c alpha is the smp (output)
c mindf is the minimum degree of freedom possible for this spline/pen
c combination. (note that this is double precision)
c md is the number of nonzero diagonals in the penalty matrix
c mr the actual row dim of pen mat
c calls cholg, solve, (and degf2) and penh.
c
    double precision vv(mr,md),vvi(ntud3),v3(lb,lb),df
    double precision al,au,ap,del,var(lb,lb),alpha
    double precision fl,fu,fp,v4(ntud3,ntud3),mindf,eps
    eps=1.d-3
    mit=100
    nit=0
    if (df.ge.ntud3) then

```

```

    alpha=0
    df=ntud3
    return
else if (df.le.mindf) then
    df=mindf+.05d0
endif
c calculate smoothing parameter to get appropriate degrees of freedom
c start nonpd algorithm:
200 iopt=0
    ap=alpha
    fu=df+1
202 call degf2(vv,v3,var,v4,vvi,ntud3,ap,lb,fp,md,mr)
if (v4(1,1).lt.0) then
    iopt=1
    return
endif
nit=nit+1
if (abs(fp-df).lt.eps) go to 250
if (nit.gt.mit) then
    iopt=1
    go to 250
endif
if (fp.gt.df) then
    al=ap
    fl=fp
    go to 205
else
    au=ap
    fu=fp
    ap=ap/4
    go to 202
endif
205 if (fu.lt.df) go to 208
au=al
del=4
206 au=au*del
call degf2(vv,v3,var,v4,vvi,ntud3,au,lb,fu,md,mr)
if (v4(1,1).lt.0) then
    iopt=1
    return
endif
nit=nit+1
if (abs(fu-df).lt.eps) then
    ap=au
    fp=fu
    go to 250
endif
if (nit.gt.mit) then
    iopt=1
    go to 250
endif
if (fu.gt.df) then
    al=au
    fl=fu
    go to 206
endif
208 ap=(al+au)/2
call degf2(vv,v3,var,v4,vvi,ntud3,ap,lb,fp,md,mr)
if (v4(1,1).lt.0) then
    iopt=1
    return
endif
nit=nit+1

```

```

if (abs(fp-df).lt.eps) go to 250
if (nit.gt.mit) then
  iopt=1
  go to 250
endif
if (fp.gt.df) then
  al=ap
  fl=fp
else
  au=ap
  fu=fp
endif
go to 208
250 alpha=ap
df=fp
110 format(4e18.10)
c      write (70,*) nit
      return
end

      subroutine degf2(vv,v3,var,v4,vvi,ntud3,a0,lb,t1,md,mr)
c  on output t1 is the trace (=df)
c  vv is pen v3 is inf v4 vvi are work var is the true var
      double precision vv(mr,md),v3(lb,lb),v4(ntud3,ntud3),a0
      double precision vvi(ntud3),t1,var(lb,lb)
      do 10 i=1,ntud3
      do 10 j=1,ntud3
      v4(i,j)=v3(i,j)
10    continue
      call penh(ntud3,vv,mr,md,v4,1,ntud3,a0)
      call cholg(ntud3,v4,v4,ntud3,ntud3)
      if (v4(1,1).lt.0) then
      call dblepr("not pd in degf",14,v4(1,1),1)
143   format ('matrix not pd in degf')
      return
      endif
      t1=0
      do 20 i=1,ntud3
      call solve(ntud3,v4,var(1,i),vvi,ntud3)
      t1=t1+vvi(i)
20    continue
      return
end

      subroutine penlik(nb,beta,penm,mr,md,lik,s,alpha)
c given a penalty matrix penm and parameter values beta,
c subtracts penalty terms from liklihood and score
c the beginning index of beta and s (the score) in the
c call should correspond to where the spline terms begin
c ie call penlik(...beta(7)...s(7)... ) if the spline is
c stored consecutively beginning with the 7th component
c nk is the number of knots, nb is the number
c of basis fcns actually in use, alpha is the value of the
c smoothing parameter.
      double precision penm(mr,md),beta(1),s(1),lik,alpha,u
      do 10 i=1,nb
      u=0
      m1=max(1,i-md+1)
      m2=min(nb,i+md-1)
      l=0
      do 12 j=i,m2
      l=l+1
      u=u+penm(i,l)*beta(j)
12    continue
      lik=lik-u

```

```

12 continue
  if (m1.lt.i) then
    l=i-m1+2
    do 11 j=m1,i-1
    l=l-1
    u=u+beta(j)*penm(j,l)
11  continue
  endif
  s(i)=s(i)-alpha*u
  lik=lik-alpha*u*beta(i)/2
10  continue
  return
end

subroutine sph(nov,s,c,x,tknot,beta,lik0,lik,sb,vv,
&iwork,dsub,isub,testr,eig,eps,z,exz,s0,s1,ts1,w,
&wp,ts1p,id,s1p,z1,z2,gn)
  double precision s(1),c(1),x(1),tknot(1),times(1),beta(1),lik0
  double precision lik(2),eps,sb(1),vv(1),dsub(1),testr(1),eig(1)
  double precision uplik
  double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C>>>>
  double precision wp(1),ts1p(1),s1p(1),z2(1)
  integer id(1),gn(1)
C<<<<
  integer nov(15),iwork(1),isub(1)
  common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
  &nk2,nk3,nk4,malph
  common /up/ uplik
c
c all covariates get included in the first
c nfx columns of x. nsx then gives the number that will have spline
c components. That is, splines will be linear+B-spline terms. So
c for each spline variable need nknot+2 B-spline terms.
c
c dsub must have length at least 3*nsx+4*nsx*(nknot+4)+2*(nknot+3)^2
c +np*np+max(no+5*np,(nknot+3)*((nknot+3)))+(if ksmopt>0) (2*(nknot+3)*nkk
c iwork must have length at least 3*no+max(no,np)+1+2*nstr+2*nstr (since ntud=1)
c where nstr is the # strata.
c
c nov parameters:
c   1=no, 2=nrx, 3=ncx, 4=istr (col # of x for strata, -1 if none,
c   5=maxiter, 6=knot option (1 use knots provided, 0 program calculates),
c   7=smoothing option (<0 use input smoothing params, 0 calc smoothing
c   param only in first iteration, >0 recalc after each iteration)
c   8=nfx total # covariates, 9=nsx, the # using splines
c   10=analysis option (<=0 estimates only, 1 est & var, >1 est, var & test),
c   12=nknot, 13=nest
c
c isub(7) to isub(6+nsx) must have the col #'s in x of the spline terms
c testr must have length nsx*6, where nsx is nov(9)
c length(eig)=2*nsx*(nknot+3): eig will contain test eigenvalues (output)
c x needs to have ncx+nsx*4 columns, to include all spline basis functions
c if smoothing parameters are input need to be in
c components nsx*2+1 to nsx*3 of dsub
c on output x(.,ncx+1) to x(.,ncx+2*nsx) will be overwritten with
c estimates of spline terms and variances
c dsub 3*nsx+1,... will have inverse 2nd deriv matrix
c
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of penm's
c isub(5) reserved for the beginning index (in dsub) of unpenalized inf
c isub(6+1) to isub(6+nsx) gives col #'s (in X) of spline covs.

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c isub(6+1+nsx) will be the beginning index of inter.
c
c dsub(1:nsx) are target df.
c dsub(nsx+1:2*nsx) will be attained df.
c dsub(2*nsx+1:3*nsx) will be smoothing parameters
c
c isub must have length at least 6+nsx+nsx*no
c
c     calls sph1, calvar, tstcv, cestc directly
c indirect calls: coxrg, coxft, coxg, strat, ut, uft, tint, tint4
c     cholg, solve sortg dperm sub8, sub8c, sub9c, eval8, act8, extrct,
c     matop5, adisw, intsb2, pencb, spln, penh, gdf, degf2, penlik, eigen,
c     qfg, pdi, pdi2
        no=nov(1)
        nrx=nov(2)
        ncx=nov(3)
        istr=nov(4)
        nknot=nov(12)
        mit=nov(5)
        knopt=nov(6)
        ksmopt=nov(7)
        nfx=nov(8)
        nsx=nov(9)
        nest=nov(13)
        nk2=nknot+2
        nk3=nknot+3
        nk4=nknot+4
        np=nfx+nsx*nk2
        nkk=max(np-nk2,nk2)
        kmd=(np-nk2)*(np-nk2)
        kmd=max(kmd,nk3*nk3+2*nk2*nkk)
c
c in subsequent routines dsub is divided into dsub & work.
c the dsub part must be 3*nsx (for dft,dfa,alpha)
c             + 4*(nknot+4)*nsx (penalty matrices)
c             + np*np (unpen inf)
c + either nk3*nk3 (ksmopt<=0) or kmd (ksmopt>0)
c the work part (indexed beginning with mdw) needs to be
c at least max(no,4*np+np*np in the call to coxrg
c in the call to tstcv the total length of dsub must be at least
c 3*nsx+np*np+(nk3*nk3)*3
c
if (ksmopt.gt.0) then
    mdw=3*nsx+np*np+4*nk4*nsx+kmd+1
else
    mdw=3*nsx+np*np+4*nk4*nsx+nk3*nk3+1
endif
call sph1 (s,c,x,istr,isub,tknot,times,
&beta,lik0,lik(1),sb,vv,iwork,dsub(mdw),dsub(1),eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1,z2,gn)
if (icnv.gt.0) then
    vv(1)=-icnv
    return
endif
lik(2)=uplik
if (nsx.gt.0) then
    mdw2=3*nsx+1+np*np
    mdw3=mdw2+nk3*nk3
    mdw4=mdw3+nk3*nk3
c note the call to tstcv overwrites some of the index info
c in iwork
    if (nov(10).gt.0) then
        call calvar(vv,np,np,dsub(isub(5)),np,vv,np,dsub(mdw))

```

```

c copy inverse 2nd deriv matrix to dsub(3*nsx+1),...
m2=3*nsx+1
do 21 i=0,np*np-1
21   dsub(m2+i)=dsub(mdw+i)
if (nov(10).gt.1) then
  call tscv(np,nk2,nfx,nsx,beta,vv,np,dsub(m2),np,
&      dsub(mdw2),dsub(mdw3),testr,iwork,dsub(mdw4),isub(7),eig)
  endif
endif
call cestc(x,nrx,ncx,no,nfx,nsx,isub(7),beta,vv,
& np,np,nknot,tknot,nov(12),nest)
endif
nov(12)=nknot
return
end

c routine sph1 for fitting proportional hazards spline models
c called by sph
c file contains sph1, sub8, sub8c, sub9c, eval8, act8
c all but sph1 are called by coxrg as part of fitting the model
c
c routines in this file call coxrg, strat, sortg, extrct, cholg, matop5,
c adisw, intsb2, pencb, spln, penh, gdf, penlik.
c
c
subroutine sph1(s,c,x,istr,isub,tknot,
&times,beta,lik0,lik,sb,vv,iwork,work,dsub,eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1,z2,gn)
double precision vv(1),sb(1),work(1),beta(1),times(1),s(no)
double precision dsub(1),tknot(-2:nk3,nsx)
double precision c(no),x(nrx,ncx),eps,lik,lik0
C*****
double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C*****
C>>>
double precision wp(1),ts1p(1),s1p(1),z2(no,np)
integer id(1),gn(1)
C<<<<
integer iwork(1),isub(1)
external eval8,act8,sub8
common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
c
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of penm's
c isub(5) reserved for the beginning index (in dsub) of unpenalized inf
c isub(6+1) to isub(6+nsx) gives col #'s (in X) of spline covs.
c isub(6+1+nsx) will be the beginning index of inter.
c
c dsub(1:nsx) are target df.
c dsub(nsx+1:2*nsx) will be attained df.
c dsub(2*nsx+1:3*nsx) will be smoothing parameters
c
c X must have ncx+4*nsx columns although when the call is made
c
c isub must have length at least 6+nsx+nsx*no
c
c first get knot locations:
c
c this repeats some code from coxrg, but need to determine this first
c so splines can be calculated before calling coxrg.
c
mnstr=1
miuti=mnstr+1
mtduiti=miuti+2*no

```

```

c miwk needs length max(no,np)
  miwk=mtduti+no
  mnosr=miwk+max(no,np)
  call strat(istr,s,c,x,iwork(mnosr),nstr,iwork(miwk),work,
  &no,ncx,nrx,iwork(miuti),iwork(mtduti))
c      write (6,*) nstr
c      call pr1(iwork(miuti),iwork(mtduti),no)
  iwork(mnstr)=nstr
  mnuti=mnosr+nstr
c proportional hazards, so length of itud is 2*nstr
  ntud=1
  mitud=mnuti+nstr
c      write (6,*) (iwork(i),i=mnosr,mnosr+nstr-1)
c      write (6,*) (iwork(i),i=mnuti,mnuti+nstr-1)
  malph=2*nsx+1
  if (ksmopt.ge.0) then
    do 749 i=0,nsx-1
  749   dsub(malph+i)=1
  endif
  isub(3)=7+nsx
  isub(4)=malph+nsx
  isub(5)=isub(4)+4*nk4*nsx
c      write (6,*) nknot,no
c calc knot locations and penalty matrices
  89  nknp=nknot
  l1=isub(4)
  if (nsx.gt.0) then
    do 23 k=1,nsx
    if (knopt.le.0) then
      do 5 i=1,no
      work(i)=x(i,isub(6+k))
  5   continue
      call sortg(work,iwork(miwk),1,no)
      call intsb2(nknot,work,tknot(0,k),no)
      do 6 i=1,2
      tknot(-i,k)=tknot(0,k)
  6   tknot(nknot+i+1,k)=tknot(nknot+1,k)
    endif
    call pencb(nknot,tknot(-2,k),dsub(l1),nk4)
    l1=l1+4*nk4
  23  continue
    if (nknot.ne.nknp) then
      nk2=nknot+2
      nk3=nknot+3
      nk4=nknot+4
      go to 89
    endif
c calc splines
  do 25 k=1,nsx
  l1=isub(3)+(k-1)*no-1
c last +1 in 12 is to allow for the strata variable
  12=ncx+4*(k-1)
  13=isub(6+k)
c      do 35 i = 1, no
c          do 30 j = 1, np
c              z2(i,k) = x(i,k)
  c 30   continue
c 35   continue
  do 26 i=1,no
  call spln(nknot,3,tknot(-2,k),x(i,13),1,isub(l1+i),work,1)
  do 27 ii=1,4
  27 x(i,l2+ii)=work(ii)
  26   continue

```

```

25 continue
endif
np=nfx+nsx*nk2
ntime=0
do 35 i = 1, no
    do 30 j = 1, np
        z2(i,j) = x(i,j)
30    continue
35 continue
do 28 i=1,np
28 beta(i)=0
call coxrg(s,c,x,no,nrx,ncx,istr,np,mit,1,ntime,times,beta,
&lik0,lik,sb,vv,iwork,work,act8,eval8,sub8,isub,dsub,eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1,gn)
return
end

subroutine eval8(g, gp, b, lci, i, l, s, c, xx, nrx2, ncx2, no2, np2, lc2, nact,
&iact, isub, dsub)
double precision gp(np), s(no), c(no), xx(nrx,ncx), dsub(1), b(np), g
integer isub(1), iact(1)
common /params/ no, nrx, ncx, nknot, mit, knopt, ksmopt, np, nfx, nsx,
&nk2, nk3, nk4, malph
do 9 ii=1,np
9 gp(ii)=0
g=0
do 10 ii=1,nfx
    gp(ii)=xx(i,ii)
10 g=g+gp(ii)*b(ii)
do 15 ii=1,nsx
    l1=ncx+4*(ii-1)
    l2=isub(3)+(ii-1)*no+i-1
    if (isub(l2).eq.0) go to 15
    l3=nfx+nk2*(ii-1)+isub(l2)-1
    l4=min(nk2, isub(l2)+3)-isub(l2)+1
    do 11 iq=1,l4
        gp(l3+iq)=xx(i,l1+iq)
11    g=g+gp(l3+iq)*b(l3+iq)
15 continue
return
end

subroutine eval8n(zc, lci, i, l, s, c, xx, nrx2, ncx2, no2, np2, lc2, nact,
&iact, isub, dsub, nfx1)
double precision s(no), c(no), xx(nrx,ncx), dsub(1)
integer isub(1), iact(1), zc(ncx,3)
common /params/ no, nrx, ncx, nknot, mit, knopt, ksmopt, np, nfx, nsx,
&nk2, nk3, nk4, malph
nfx1 = nfx
do 15 ii=1,nsx
    l1=ncx+4*(ii-1)
    l2=isub(3)+(ii-1)*no+i-1
    if (isub(l2).eq.0) go to 15
    l3=nfx+nk2*(ii-1)+isub(l2)-1
    l4=min(nk2, isub(l2)+3)-isub(l2)+1
    zc(ii,1) = l1
    zc(ii,2) = l3
    zc(ii,3) = l4
15 continue
return
end

```

```

subroutine act8(lci,11,12,iuti,no2,s,c,x,nrx2,ncx2,iact,nact,
&np2,lc2,isub,dsub)
  double precision s(no),c(no),x(nrx,ncx),dsub(1)
  integer iuti(no,2),iact(1),isub(1)
  common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
  do 10 i=1,np
10  iact(i)=i
  nact=np
  return
  end

subroutine sub9c(beta,lik,sb,vv,penm,v3,dft,dfa,alpha,v6,vvi)
  double precision beta(np),lik,sb(np),vv(np,np),v3(np,np)
  double precision penm(nk4,4,nsx),dft(nsx),dfa(nsx),alpha(nsx)
  double precision v6(1),mindf,vvi(np,np)
  common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
c first update values of smoothing parameters to give required df.
  mv8=1+nk2*nk2
  np2=np-nk2
  nkk=max(np2,nk2)
  mv9=mv8+nk2*nkk
  do 820 i=1,np
  do 820 j=1,np
820  vvi(i,j)=vv(i,j)
  do 821 j=2,nsx
    11=nfx+1+(j-1)*nk2
    call penh(nk2,penm(1,1,j),nk4,4,vvi,11,np,alpha(j))
821  continue
  do 819 j=1,nsx
    11=nfx+1+(j-1)*nk2
    12=11+nk2-1
    call extrct(vvi,np,np,11-1,12+1,v6,np2)
    call cholg(np2,v6,v3,np2,np2)
    call matop5(vv,np,v3,np2,11,12,v6(mv8),nk2,v6(mv9),v6,nkk)
    dfa(j)=dft(j)
c 3rd from last arg in gdf # dcols used
c in pen matrix, 4th from last mindf Here this is 0 since const and lin
c terms modelled separately
  mindf=0.d0
  call gdf(nk2,penm(1,1,j),v6(mv9),v6(mv8),v6,nkk,dfa(j),
&alpha(j),v3,mindf,4,nk4,iopp1)
c
c  if alg failed then return without applying penalty
c
  if (iopp1.gt.0) return
  write (6,100) dft(j),dfa(j),alpha(j)
  call penh(nk2,penm(1,1,j),nk4,4,vvi,11,np,alpha(j))
  if (j.lt.nsx) then
    call penh(nk2,penm(1,1,j+1),nk4,4,vvi,11+nk2,np,-alpha(j+1))
  endif
819  continue
c then copy vv to v3, vvi to vv:
  do 10 i=1,np
  do 10 j=1,np
    v3(i,j)=vv(i,j)
10  vv(i,j)=vvi(i,j)
c then add penalty terms to lik,score, (lik,sb)
  do 535 j=1,nsx
    11=nfx+1+(j-1)*nk2
    call penlik(nk2,beta(11),penm(1,1,j),nk4,4,lik,sb(11),alpha(j))

```

```

c      call penh(nk2,penm,nk4,4,vv,l1,np,alpha(j))
535  continue
100  format (3e15.8)
      return
      end

      subroutine sub8(np2,beta,lik,sb,vv,vvi,isub,dsub,nit)
double precision beta(np),lik,sb(np),vv(np,np),dsub(1)
double precision vvi(1),uplik
integer isub(1)
common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
      common /up/ uplik
c to provide working matrices, dsub must be at least
c 3*nsx+4*(n or ntud,depending on ntime)+np*np+2*(nknot+4)**2
      mv6=isub(5)+np*np
      uplik=lik
c      call dblepr('lik',3,lik,1)
      if (nsx.gt.0) then
          if (ksmopt.gt.0.and.nsx.gt.1.and.nit.lt.9.and.nit.gt.0) then
              call sub9c(beta,lik,sb,vv,dsub(isub(4)),dsub(isub(5)),
&dsub(1),dsub(1+nsx),dsub(malph),dsub(mv6),vvi)
          else
              call sub8c(beta,lik,sb,vv,dsub(isub(4)),dsub(isub(5)),
&dsub(1),dsub(1+nsx),dsub(malph),dsub(mv6),vvi,nit)
          c      write (6,*) lik
          endif
          endif
cc      call dblepr('dsub',4,dsub,3)
c      call dblepr('plik',4,lik,1)
110  format (e20.12)
      return
      end

      subroutine sub8c(beta,lik,sb,vv,penm,v3,dft,dfa,
&alpha,v6,v7,nit)
double precision beta(np),lik,sb(np),vv(np,np),v3(np,np)
double precision penm(nk4,4,nsx),dft(nsx),dfa(nsx),alpha(nsx)
double precision v6(nk2,nk2),v7(nk2,nk2),mindf
common /params/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph
c first update values of smoothing parameters to give required df.
c note that adisw does not change vv.
c ksmopt<0 means use input value of alpha
c ksmopt=0 means calc using crude alg on first iter only
      if ((nit.eq.1.and.ksmopt.eq.0).or.(ksmopt.gt.0.and.nit.lt.9.
&and.nit.gt.0)) then
          do 819 j=1,nsx
              l1=nfx+1+(j-1)*nk2
              l2=l1+nk2-1
              call adisw(l1,l2,np,vv,v3,np)
              dfa(j)=dft(j)
c 3rd from last arg in gdf # dcols used
c in pen matrix, 4th from last mindf Here this is 0 since const and lin
c terms modelled separately
              mindf=0.d0
              call gdf(nk2,penm(1,1,j),v3,v3,v6,np,dfa(j),alpha(j),v7,mindf,
&4,nk4,iopp1)
c
c  if alg failed then return without applying penalty
c
              if (iopp1.gt.0) return
              write (6,100) dft(j),dfa(j),alpha(j)

```

```

819 continue
endif
c then copy vv to v3:
do 10 i=1,np
do 10 j=1,np
10 v3(i,j)=vv(i,j)
c then add penalty terms to lik,score, and inf (lik,sb,vv)
do 535 j=1,nsx
11=nfx+1+(j-1)*nk2
call penlik(nk2,beta(11),penm(1,1,j),nk4,4,lik,sb(11),alpha(j))
call penh(nk2,penm(1,1,j),nk4,4,vv,11,np,alpha(j))
535 continue
100 format (3e15.8)
return
end

subroutine tstcv(np,nk2,nfx,nsx,beta,vv,lvv,v4,lv4,v6,v7,test,
&iwork,v8,isx,eig)
c for spline given as linear+(nknot+2) cubic B-spline terms
c calcs Wald tests of no association and linearity
c nfx is the total number of vars, nsx the number of spline terms.
c isx(j) gives which entry in beta corresponds to the linear part of the
c jth spline term.
c on input v4 is the inverse penalized 2nd derivative matrix, and vv the
c estimated var-cov matrix of the paramters estimates.
c in this routine iwork must be at least 2*(nk2+1)
c in length
c v6,v7,v8 are work space,
c on output eig will contain the eigenvalues and test the test
c results
c calls pdi2, sortg, eigen from ~/src/util.a & qfg from ~/src/dist.a
    double precision vv(lvv,np),beta(np),v4(lv4,np)
    double precision v6(nk2+1,nk2+1),v7(nk2+1,nk2+1),test(6,nsx)
    double precision v8(nk2+1,nk2+1),qfg,trace(7),eig(nk2+1,2*nsx)
    integer iwork(1),isx(nsx)
c stats:
    nk3=nk2+1
    do 50 k=1,nsx
        kk=isx(k)
        11=nfx+(k-1)*nk2
c overall test for var i:
c first calc stat:
    v6(1,1)=v4(kk,kk)
    v8(1,1)=beta(kk)
    do 15 i=1,nk2
        ii=i+1
        v6(1,ii)=v4(kk,11+i)
        v6(ii,1)=v6(1,ii)
        v8(ii,1)=beta(11+i)
        do 15 j=1,nk2
            v6(ii,j+1)=v4(i+11,j+11)
15 continue
    call pdi2(nk3,v6,v7,nk3,nk3)
    test(1,k)=0
    do 27 ii=1,nk3
        iwork(ii)=1
        test(2,k)=0
        do 28 jj=1,ii
28    test(2,k)=test(2,k)+v7(jj,ii)*v8(jj,1)
        test(1,k)=test(1,k)+test(2,k)*test(2,k)
27 continue
c then get eigenvalues to calc dist
    v6(1,1)=vv(kk,kk)

```

```

do 17 i=1,nk2
ii=i+1
v6(1,ii)=vv(kk,l1+i)
v6(ii,1)=v6(1,ii)
do 17 j=1,nk2
17 v6(ii,j+1)=vv(i+l1,j+l1)
do 18 i=1,nk3
do 18 j=i,nk3
v8(i,j)=0
do 19 ii=1,i
do 19 jj=1,j
19 v8(i,j)=v8(i,j)+v7(ii,i)*v6(ii,jj)*v7(jj,j)
v8(j,i)=v8(i,j)
18 continue
call eigen(v8,nk3,nk3,eig(1,2*k-1),v6,1,iflg)
call sortg(eig(1,2*k-1),iwork,1,nk3)
do 37 ii=1,nk3
iwork(ii)=1
37 v8(ii,1)=0
test(2,k)=1-qfg(eig(1,2*k-1),v8,iwork,nk3,0.d0,test(1,k),5000,
&1.d-5,iwork(nk3+1),trace,efault)
test(3,k)=0
do 35 i=1,nk3
35 test(3,k)=test(3,k)+eig(i,2*k-1)
c test for linearity :
c calc stat:
do 215 i=1,nk2
do 215 j=1,nk2
215 v6(i,j)=v4(i+l1,j+l1)
c call matwr(v6,nk2,nk3,88)
call pdi2(nk2,v6,v7,nk3,nk3)
test(4,k)=0
do 427 ii=1,nk2
iwork(ii)=1
test(5,k)=0
do 428 jj=1,ii
428 test(5,k)=test(5,k)+v7(jj,ii)*beta(l1+jj)
test(4,k)=test(4,k)+test(5,k)*test(5,k)
427 continue
c get eigenvalues:
do 317 i=1,nk2
do 317 j=1,nk2
317 v6(i,j)=vv(i+l1,j+l1)
c call matwr(v6,nk2,nk3,89)
do 318 i=1,nk2
do 318 j=i,nk2
v8(i,j)=0
do 319 ii=1,i
do 319 jj=1,j
319 v8(i,j)=v8(i,j)+v7(ii,i)*v6(ii,jj)*v7(jj,j)
v8(j,i)=v8(i,j)
318 continue
c call matwr(v8,nk2,nk3,90)
call eigen(v8,nk2,nk3,eig(1,2*k),v6,1,iflg)
call sortg(eig(1,2*k),iwork,1,nk2)
do 327 ii=1,nk2
iwork(ii)=1
327 v8(ii,1)=0
test(5,k)=1-qfg(eig(1,2*k),v8,iwork,nk2,0.d0,test(4,k),5000,
&1.d-5,iwork(nk3+1),trace,efault)
test(6,k)=0
do 335 i=1,nk2
335 test(6,k)=test(6,k)+eig(i,2*k)

```

```

50 continue
      return
      end

      subroutine cestc(x,nrx,ncx,no,nfx,nsx,isx,
     &beta,var,np,lvar,nknot,tknot,nknto,nest)
c routine for calculating function estimates and variances
c from output of sph/wsph.
c x is covariate values including spline terms.  for info on ordering
c etc see sph.f
      double precision x(nrx,1),est,v,c(9),w(4),bl,w2(4),q1,q2,q3
      double precision beta(np),var(lvar,np),tknot(-2:(nknto+3),nsx)
      integer inter,isx(1),ind(9)
      if (nest.gt.no) nest=no
      nk2=nknot+2
      do 5 k=1,nsx
         kk=isx(k)
         l1=nfx+(k-1)*nk2
c
c   baseline value based on knots:
c
c       bl=(tknot((nknot+1)/2,k)+tknot((nknot+2)/2,k))/2
c
c   or on means of covariates:
c
      bl=0
      do 7 ii=1,no
    7 bl=bl+x(ii,kk)
      bl=bl/no
c
      call spln(nknot,3,tknot(-2,k),bl,1,intbl,w,1)
      m2=min(intbl+3,nk2)-intbl+1
      do 201 jj=1,m2
         ind(jj)=intbl-1+jj+11
    201 c(jj)=-w(jj)
         q1=bl
         q2=q1
         do 205 ii=1,no
            if (x(ii,kk).ge.tknot(0,k)) q1=min(x(ii,kk),q1)
    205 if (x(ii,kk).le.tknot(nknot+1,k)) q2=max(x(ii,kk),q2)
         q3=(q2-q1)*1.d-8
         q1=q1+q3
         q2=q2-q3
         do 50 i=1,nest
            q3=q1+(i-1)*(q2-q1)/(nest-1)
            call spln(nknot,3,tknot(-2,k),q3,1,inter,w2,1)
            if (inter.le.0) then
               x(i,kk)=-1
               x(i,ncx+2*k-1)=-1
               x(i,ncx+2*k)=-1
               go to 50
            endif
            12=11+inter-1
            13=min(inter+3,nk2)-inter+1
            do 202 jj=1,13
               ind(jj+m2)=12+jj
    202 c(jj+m2)=w2(jj)
               lu=13+m2+1
               ind(lu)=kk
               c(lu)=q3-bl
               est=0
               v=0
               do 10 ii=1,lu

```

```

        est=est+beta(ind(ii))*c(ii)
        do 11 jj=1,lu
11      v=v+var(ind(ii),ind(jj))*c(ii)*c(jj)
10      continue
       x(i,kk)=q3
       x(i,ncx+2*k-1)=est
       x(i,ncx+2*k)=v
50      continue
5      continue
       return
       end

c
c This software comes with absolutely no guarantees. You have permission to
c use it for any noncommercial purpose, and to modify it as needed.
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c

      subroutine sphip(nov,s,c,x,tknot,beta,lik0,lik,sb,vv,
     &iwork,dsub,isub,testr,eig,eps,z,exz,s0,s1,ts1,w
     &wp,ts1p,id,s1p,z1)
      double precision s(1),c(1),x(1),tknot(1),times(1),beta(1)
      double precision lik(2),eps,sb(1),vv(1),dsub(1),testr(1),eig(1)
      double precision uplik,scx,ucx,scy,ucy,lik0
C*****
      double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C*****
C>>>>
      double precision wp(1),ts1p(1),s1p(1)
      integer id(1)
C<<<<
      integer nov(15),iwork(1),isub(1)
      common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
     &np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
      common /up/ uplik
c
c includes nfx fixed covs and a 2-cov interaction modelled using
c tensor product B-splines
c on input isub(7) to isub(6+nfx) should contain the column numbers
c in x for the fixed covs. isub(1) and isub(2) should give the column
c numbers of the 1st and 2nd terms for the spline interaction
c (isx and isy)
c
c nov parameters:
c   1=no, 2=nrx(=no), 3=ncx, 4=istr (col # of x for strata, -1 if none,
c   5=maxiter, 6=knot option (1 use knots provided, 0 program calculates),
c   7=smoothing option (<0 use input smoothing params, 0 calc smoothing
c   param only in first iteration, >0 recalc after each iteration)
c   8=nfx (nfx+nsx in sph, nfx is #col in linear.cov, nsx in spline.cov),
c   9=nsx, 10=analysis option (<=0 estimates only, 1 est & var, >1 est,
c   var & test),
c     sphip: 11,12=nknx,nkny, 13,14=nestx,nesty, 15=rescale opt (0=rescale
c   covariates, 0=do not)
c
c dsub must have length max(3+nex*ney*4+np*np,3+16*nkx4*nky4+nb*nb+np*np+
c max(max(no,4*np*np*np),np*np+nb*nb))
c iwork must have length at least 3*no+max(no,np)+1+2*nstr+2*nstr (since ntud=1)
c where nstr is the # strata.
c isub must have length 6+nfx+2*no
c testr must have length 12, (order will be stat,pv,df for overall
c test, ave effect of x, ave effect of y, no interaction test)
c length(eig)=4*(nkx4*nky4-1): eig will contain test eigenvalues (output)
c on input ncx should indicate the number of columns of x with input
c data (which should be consecutive in X. columns ncx+1 to ncx+4 will

```

```

c be used for x bspline terms, and columns ncx+5 to ncx+8 for the y
c bspline terms. Note that x therefore must have at least ncx+8 columns
c if smoothing parameter is input needs to be in component 1 of dsub
c target df for total 2 var interaction term needs to be in dsub(2)
c and attained df on output will be in dsub(3)
c On output estimates will be in
c dsub(4), ... dsub(3+nex*ney*4), with the first nex*ney comps giving
c the x coords, then the y coords, then the estimates, then the variances
c total # parameters is nfx+(nknx+4)*(nkny+4)-1
c inverse 2nd deriv matrix will be in dsub(4+nex*ney)...
c
c total length of tknot must be nknx+6+nkny+6
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of penm's
c isub(5) reserved for the beginning index (in dsub) of unpenalized inf
c isub(6+1) to isub(6+nfx) gives col #'s (in X) of fixed covs.
c isub(6+1+nfx) will be the beginning index of inter.
c
c
c note that the parameter order loops over y within x: ie
c ordered b(x1,y1),b(x1,y2),...,b(x1,y[nky+4]),b(x2,y1),...
c
c     calls sphil, calvar, testi, cesti directly
c
c indirect calls: coxrg, coxft, coxg, strat, ut, uft, tint, tint4
c cholg, solve sortg dperm subi, subic, evali, acti,
c adisw, intsb2, pencxc, spln, penh2, gdf2, degf3, penlk2, eigen,
c qfg, pdi, pdi2, penhg, penlkg, csplcf, int1, int2, st1, st2, intb1
c
no=nov(1)
nrx=nov(2)
ncx=nov(3)
istr=nov(4)
isx=isub(1)
isy=isub(2)
nknx=nov(11)
nkny=nov(12)
mit=nov(5)
knopt=nov(6)
ksmopt=nov(7)
nfx=nov(8)
nkx2=nknx+2
nkx3=nknx+3
nkx4=nknx+4
nky2=nkny+2
nky3=nkny+3
nky4=nkny+4
np=nfx+nky4*nkx4-1
nb=np-nfx
c
c in subsequent routines dsub is divided into dsub & work.
c the dsub part must be 3 (for alpha,dft,dfa)
c           + 4*(nkx4)*4*(nky4) (penalty matrices)
c           + np*np (unpen inf)
c           + nb*nb (work matrix in dsub)
c the work part (indexed beginning with mdw) needs to be
c at least max(no,4*np+np*np) in the call to coxrg
c & at least np*np+nb*nb for the call to testi
c
c also, to hold the output, the two parts together must be at least
c 3+nestx*nesty*4+np*np
c
mdw=3+np*np+4*nkx4*4*nky4+nb*nb

```

```

if (nov(15).le.0) then
  ucx=0
  ucy=0
  isx2=(sx-1)*nr
  isy2=(sy-1)*nr
  do 444 i=1,no
    ucx=ucx+x(i+isx2)
  444  ucy=ucy+x(i+isy2)
    ucx=ucx/no
    ucy=ucy/no
    scx=0
    scy=0
    do 443 i=1,no
      scx=scx+(x(i+isx2)-ucx)**2
  443  scy=scy+(x(i+isy2)-ucy)**2
      scx=sqrt(scx/(no-1))
      scy=sqrt(scy/(no-1))
      do 442 i=1,no
        x(i+isx2)=(x(i+isx2)-ucx)/scx
  442  x(i+isy2)=(x(i+isy2)-ucy)/scy
else
  ucx=0
  ucy=0
  scx=1
  scy=1
endif
call sphil (s,c,x,istr,isub,tknot,times,
&beta,lik0,lik(1),sb,vv,iwork,dsub(mdw),dsub(1),eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
if (icnv.gt.0) then
  vv(1)=-icnv
  return
endif
lik(2)=uplik
nb=np-nfx
m6=4
m7=m6+nb*nb
m8=mdw+np*np
if (nov(10).gt.0) then
  call calvar(vv,np,np,dsub(isub(5)),np,vv,np,dsub(mdw))
endif
if (nov(10).gt.1) then
  call testi(np,nb,nkx4,nky4,nfx,beta,vv,np,dsub(mdw),np,
  &dsub(m6),dsub(m7),testr,iwork,dsub(m8),eig,nknx,nkny,tknot)
endif
c move the inverse 2nd deriv matrix from cal var to start at
c dsub(4+nex*ney):
  m2=4+4*nov(13)*nov(14)
  if (m2.gt.mdw) then
    do 20 i=np*np-1,0,-1
  20  dsub(m2+i)=dsub(mdw+i)
  else if (m2.lt.mdw) then
    do 22 i=0,np*np-1
  22  dsub(m2+i)=dsub(mdw+i)
  endif
  call cesti(nknx,nkny,tknot,np,beta,vv,np,nov(13),nov(14),dsub(4),
  &nfx,ucx,ucy,scx,scy)
  return
end
c fits 2 dim tensor product bspline in prop hazards model
c file contains sphil acti evali subi subic gdf2 degf3
c acti evali subi subic are called by coxrg as part of fitting model
c

```

```

c in addition, routines in this file call strat, sortg, coxrg,
c cholq, adisw, solve, intsb2, pencxc, spln, penh2, penlk2
c
c subroutine sph1(s,c,x,istr,isub,tknot,
c &times,beta,lik0,lik,sb,vv,iwork,work,dsub,eps,icnv,
c &z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
c double precision vv(1),sb(1),work(1),beta(1),times(1),s(no)
c double precision dsub(1),tknot(-2:nkx3,2)
c double precision c(no),x(nrx,ncx),eps,lik,lik0
C*****C*****C*****C*****C*****C*****C*****C*****C*****
c double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C*****C*****C*****C*****C*****C*****C*****C*****C*****
C>>>>
c double precision wp(1),ts1p(1),s1p(1)
c integer id(1)
C<<<<
c integer iwork(1),isub(1)
c external evali,acti,subi
c common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
c &np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
c total length of tknot must be nknx+6+nkny+6
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of penm's
c isub(5) reserved for the beginning index (in dsub) of unpenalized inf
c isub(6+1) to isub(6+nfx) gives col #'s (in X) of fixed covs.
c isub(6+1+nfx) will be the beginning index of inter.
c
c X must have ncx+8 columns although when the call is made
c
c isub must have length at least 7+nfx+2*no
c
c note that the parameter order loops over y within x: ie
c ordered b(x1,y1),b(x1,y2),...,b(x1,y[nky+4]),b(x2,y1),...
c
c first get knot locations:
c
c this repeats some code from coxrg, but need to determine this first
c so splines can be calculated before calling coxrg.
c
c mnstr=1
c miuti=mnstr+1
c mtduti=miuti+2*no
c miwk needs length max(no,np)
c miwk=mtduti+no
c mnostr=miwk+max(no,np)
c call strat(istr,s,c,x,iwork(mnostr),nstr,iwork(miwk),work,
c &no,ncx,nrx,iwork(miuti),iwork(mtduti))
c      write (6,*) nstr
c      iwork(mnstr)=nstr
c      mmuti=mnostr+nstr
c proportional hazards, so length of itud is 2*nstr
c      ntud=1
c      mitud=mmuti+nstr
c      write (6,*) (iwork(i),i=mnostr,mnostr+nstr-1)
c      write (6,*) (iwork(i),i=mmuti,mmuti+nstr-1)
c      isub(3)=7+nfx
c      isub(4)=4
c      isub(5)=isub(4)+nkx4*nky4*16
c      write (6,*) no,nknx,nkny,isx,isy
c calc knot locations and penalty matrices
89  nknpx=nknx
      nknpy=nkny
      continue

```

```

if (knopt.le.0) then
  do 5 i=1,no
  work(i)=x(i,isx)
5  continue
  call sortg(work,iwork(miwk),1,no)
  call intsb2(nknx,work,tknot(0,1),no)
  do 7 i=1,2
  tknot(-i,1)=tknot(0,1)
7  tknot(nknx+i+1,1)=tknot(nknx+1,1)
  do 6 i=1,no
  work(i)=x(i,isy)
6  continue
  call sortg(work,iwork(miwk),1,no)
  call intsb2(nkny,work,tknot(0,2),no)
  do 8 i=1,2
  tknot(-i,2)=tknot(0,2)
8  tknot(nkny+i+1,2)=tknot(nkny+1,2)
endif
c      write (6,*) ((tknot(i,j),i=-2,nky3),j=1,2)
c in the next call the last 2 arguments require 36*nkx4 &36*nky4 room
call pencxc(nknx,tknot(-2,1),nkny,tknot(-2,2),dsub(isub(4)),
&dsub(isub(5)),dsub(isub(5)+12*nkx4))
if (nknx.ne.nknpx.or.nknpv.ne.nkny) then
  nkx2=nknx+2
  nkx3=nknx+3
  nkx4=nknx+4
  nky2=nkny+2
  nky3=nkny+3
  nky4=nkny+4
  go to 89
endif
102 format (4e15.6)
c calc splines
c last +1 in 12 is to allow for the strata variable
  l1=isub(3)-1
  do 26 i=1,no
  call spln(nknx,3,tknot(-2,1),x(i,isx),1,isub(l1+i),work,1)
  do 27 ii=1,4
27  x(i,ncx+ii)=work(ii)
26  continue
  l1=isub(3)+no-1
  do 28 i=1,no
  call spln(nkny,3,tknot(-2,2),x(i,isy),1,isub(l1+i),work,1)
  do 29 ii=1,4
29  x(i,ncx+4+ii)=work(ii)
28  continue
c      do 500 i=1,no
c 500  write (45,101) isub(isub(3)+i-1),isub(isub(3)+no+i-1),
c      &(x(i,j),j=1,ncx+8)
c 101  format(2i4,12e12.4)
  np=nfx+nkx4*nky4-1
c      write (6,*) np,mit
c initialize smoothing parameter for first iteration
  if (ksmopt.ge.0) dsub(1)=1.d0
  ntime=0
  do 38 i=1,np
38  beta(i)=0
  call coxrg(s,c,x,no,nrx,ncx,istr,np,mit,1,ntime,times,beta,
&lik0,lik,sb,vv,iwork,work,acti,evali,subi,isub,dsub,eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
  return
end

```

```

subroutine evali(g,gp,b,lci,i,l,s,c,xx,nrx2,ncx2,no2,np2,lc2,nact,
&iact,isub,dsub)
  double precision gp(np),s(no),c(no),xx(nrx,ncx),dsub(1),b(np),g
  integer isub(1),iact(1)
  common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
&np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
  do 9 ii=1,np
9  gp(ii)=0
  g=0
  do 10 ii=1,nfx
    gp(ii)=xx(i,isub(6+ii))
    g=g+gp(ii)*b(ii)
10  continue
c  first get which interval x and y coords are in:
  lx=isub(isub(3)+i-1)
  ly=isub(isub(3)+no+i-1)
  if (lx.le.0.or.ly.le.0) return
  ly2=ncx+4
  do 15 ii=lx,1x+3
    lx2=ii-lx+1+ncx
c 11 is one before first index of y terms for the iith x term:
  l1=nfx+nky4*(ii-1)+ly-1
  if (ii.eq.nkx4.and.ly+3.eq.nky4) then
    l2=3
  else
    l2=4
  endif
  write (44,*) i,1x,ly,ii,l1,1x2,ly2
  do 11 iq=1,l2
    gp(l1+iq)=xx(i,1x2)*xx(i,ly2+iq)
    g=g+gp(l1+iq)*b(l1+iq)
11  continue
15  continue
  return
end

subroutine acti(lci,l1,l2,iuti,no2,s,c,x,nrx2,ncx2,iact,nact,
&np2,lc2,isub,dsub)
  double precision s(no),c(no),x(nrx,ncx),dsub(1)
  integer iuti(no,2),iact(1),isub(1)
  common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
&np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
  do 10 i=1,np
10  iact(i)=i
  nact=np
  return
end

subroutine subi(np2,beta,lik,sb,vv,svi,isub,dsub,nit)
  double precision beta(np),lik,sb(np),vv(np,np),dsub(1)
  double precision svvi(1),uplik
  integer isub(1)
  common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
&np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
  common /up/ uplik
c to provide working matrices, dsub must be at least ?? (following
c line is wrong)
c 3*nsx+4*(n or ntud,depending on ntime)+np*np+2*(nknot+4)**2
  nb=np-nfx
  mv6=isub(5)+np*np
  mv7=mv6+nb*nb
c  write (6,*) lik
  uplik=lik

```

```

c      call dblepr("lik",3,lik,1)
c      call subic(beta,lik,sb,vv,dsub(isub(4)),dsub(isub(5)),
c      &dsub(2),dsub(3),dsub(1),dsub(mv6),vvi,nit,nb)
c      write (6,*) lik
c      call dblepr("plik",4,lik,1)
110  format (e20.12)
      return
      end

      subroutine subic(beta,lik,sb,vv,penm,v3,dft,dfa,
&alpha,v6,v7,nit,nb)
      double precision beta(np),lik,sb(np),vv(np,np),v3(np,np)
      double precision penm(nky4,4,nkx4,4),dft,dfa,alpha
      double precision v6(nb,nb),v7(nb,nb),mindf
      common /parami/ no,nrx,ncx,isx,isy,nknx,nkny,mit,knopt,ksmopt,
      &np,nfx,nkx2,nkx3,nkx4,nky2,nky3,nky4
c ksmopt<0 means use input value of alpha
c ksmopt=0 means calc using crude alg on first iter only
      if ((nit.eq.1.and.ksmopt.eq.0).or.(ksmopt.gt.0.and.nit.lt.9.
      &and.nit.gt.0)) then
c first update values of smoothing parameters to give required df.
c note that adisw does not change vv.
      call adisw(nfx+1,np,np,vv,v3,np)
      dfa=dft
c 3rd from last arg in gdf # dcols used
c in pen matrix, 4th from last mindf Here I guess this is 2 since as
c alpha->infty, should get lin(x)+lin(y)
      mindf=2.d0
      call gdf2(nkx4,nky4,nb,penm,v3,v3,v6,np,dfa,alpha,v7,mindf,
      &nkx4,4,nky4,4,iopp1)
c if algorithm failed in gdf return without penalizing.
      if (iopp1.gt.0) return
c      write (6,100) dft,dfa,alpha
819  continue
      endif
c then copy vv to v3:
      do 10 i=1,np
      do 10 j=1,np
10    v3(i,j)=vv(i,j)
c then add penalty terms to lik,score, and inf (lik,sb,vv)
      11=nfx+1
      call penlk2(nkx4,nky4,nb,beta(11),penm,nkx4,4,nky4,4,lik,
      &sb(11),alpha)
      call penh2(nkx4,nky4,nb,penm,nkx4,nky4,4,4,vv,11,np,alpha)
100   format (3e15.8)
      return
      end

      subroutine gdf2(nb,ny,nbt,vv,v3,var,vvi,lb,df,alpha,v4,mindf,
      &mrxy,mdy,iopt)
c change 3-21-91: initialize alpha before calling
c iopt not currently used
c on output,iopt =0 means algorithm worked ok
c iopt=1 means singular or other problems
c the fast option (iopt>0) only gives valid results if var and v3 are the same.
c nbt is the total # parameters,
c nb the number of x basis functions, ny the number of y basis
c functions
c vv is the penalty matrix , v3 is the unpenalized -2nd deriv,
c var is the actual var matrix (both have row dim lb). var and v3 can be
c the same matrix (vv v3 and var are not changed)
c vvi and v4 are working matrices with dim ntud3*ntud3.
c df is the target df (input) on output df=attained df,

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c alpha is the smp (output)
c mindf is the minimum degree of freedom possible for this spline/pen
c combination. (note that this is double precision)
c md is the number of nonzero diagonals in the penalty matrix
c mr the actual row dim of pen mat
c calls cholg, solve, (and degf3) and penh2.
    double precision vv(mry,mdy,mrx,mdx),vvi(nbt,nbt),v3(lb,lb)
    double precision al,au,ap,del,var(lb,lb),alpha,df
    double precision fl,fu,fp,v4(nbt,nbt),mindf,eps
    eps=1.d-3
    mit=100
    nit=0
    if (df.ge.nbt) then
        alpha=0
        df=nbt
        return
    else if (df.le.mindf) then
        df=mindf+.05d0
    endif
c
c calculate smoothing parameter to get appropriate degrees of freedom
c start nonpd algorithm:
200   iopt=0
    ap=alpha
    fu=df+1
202   call degf3(nbx,nby,nbt,vv,v3,var,v4,vvi,ap,lb,fp,
    &mrx,mdx,mry,mdy)
    if (v4(1,1).lt.0) then
        iopt=1
        return
    endif
    nit=nit+1
    if (abs(fp-df).lt.eps) go to 250
    if (nit.gt.mit) then
        iopt=1
        go to 250
    endif
    if (fp.gt.df) then
        al=ap
        fl=fp
        go to 205
    else
        au=ap
        fu=fp
        ap=ap/4
        go to 202
    endif
205   if (fu.lt.df) go to 208
    au=al
    del=4
206   au=au*del
    call degf3(nbx,nby,nbt,vv,v3,var,v4,vvi,au,lb,fu,
    &mrx,mdx,mry,mdy)
    if (v4(1,1).lt.0) then
        iopt=1
        return
    endif
    nit=nit+1
    if (abs(fu-df).lt.eps) then
        ap=au
        fp=fu
        go to 250
    endif

```

```

if (nit.gt.mit) then
    iopt=1
    go to 250
endif
if (fu.gt.df) then
    al=au
    fl=fu
    go to 206
endif
208 ap=(al+au)/2
call degf3(nbx,nby,nbt,vv,v3,var,v4,vvi,ap,lb,fp,
&mrx,mdx,mry,mdy)
if (v4(1,1).lt.0) then
    iopt=1
    return
endif
if (abs(fp-df).lt.eps) go to 250
nit=nit+1
if (nit.gt.mit) then
    iopt=1
    go to 250
endif
if (fp.gt.df) then
    al=ap
    fl=fp
else
    au=ap
    fu=fp
endif
go to 208
250 alpha=ap
df=fp
110 format(4e18.10)
c      write (70,*) nit
return
end

subroutine degf3(nbx,nby,nbt,vv,v3,var,v4,vvi,a0,lb,t1,
&mrx,mdx,mry,mdy)
c on output t1 is the trace (=df)
c vv is pen v3 is inf v4 vvi are work var is the true var
double precision vv(mry,mdy,mrx,mdx),v3(lb,lb),v4(nbt,nbt),a0
double precision vvi(nbt),t1,var(lb,lb)
do 10 i=1,nbt
do 10 j=1,nbt
v4(i,j)=v3(i,j)
10 continue
call penh2(nbx,nby,nbt,vv,mrx,mry,mdx,mdy,v4,1,nbt,a0)
call cholg(nbt,v4,v4,nbt,nbt)
if (v4(1,1).lt.0) then
call dblepr("not pd in degf",14,v4(1,1),1)
143 format ('matrix not pd in degf')
endif
t1=0
do 20 i=1,nbt
call solve(nbt,v4,var(1,i),vvi,nbt)
t1=t1+vvi(i)
20 continue
return
end

subroutine testi(np,nb,nx,ny,nfx,beta,vv,lvv,v4,lv4,v6,v7,test,
&iwork,v8,eig,nknx,nkny,tknot)

```

```

c For a tensor product b-spline computes test (Wald) that
c all coefs=0
c nb is the number of basis fcns for spline terms assumed to be
c components nfx+1 to nfx+nb of beta and var.
c on input v4 is the inverse penalized 2nd derivative matrix, and vv the
c estimated var-cov matrix of the parameters estimates.
c length(iwork)=2*nk4
c v6,v7,v8 are work space,
c on output eig will contain the eigenvalues and test the test
c results
c someday will add tests of no interaction, and tests for average main
c effects
c calls pdi2, sortg, eigen from ~/src/util.a & qfg from ~/src/dist.a
c also intb1
    double precision vv(lvv,np),beta(np),v4(lv4,np)
    double precision v6(nb,nb),v7(nb,nb),test(3,4)
    double precision v8(nb,nb),eig(nb,4),tknot(-2:nknx+3,2)
    integer iwork(1)
c first overall test for an effect:
c first calc stat:
    do 15 i=1,nb
    do 15 j=1,nb
        v6(i,j)=v4(i+nfx,j+nfx)
15   continue
    call st1(nb,beta(nfx+1),v6,nb,v7,test(1,1))
c then get eigenvalues to calc dist
    do 17 i=1,nb
    do 17 j=1,nb
17   v6(i,j)=vv(i+nfx,j+nfx)
    call st2(nb,v6,v7,v8,nb,test,eig,iwork)
c test for main effects of x (weighted sum over y):
c test statistic:
c (any line numbers 2xx)
c col 1 of v8 will be new params:
c get weights(integrate over y):
    call intb1(nkny,tknot(-2,2),v8(1,2))
    ny2=nb-ny*(nx-1)
    do 201 i=1,nx
        v8(i,1)=0
        lx=nfx+(i-1)*ny
        nyu=ny
        if (i.eq.nx) nyu=ny2
        do 202 k=1,nyu
202   v8(i,1)=v8(i,1)+v8(k,2)*beta(lx+k)
        do 203 j=i,nx
            nyu2=ny
            if (j.eq.nx) nyu2=ny2
            v7(i,j)=0
            lx2=nfx+(j-1)*nx
            do 204 ii=1,nyu
                do 204 jj=1,nyu2
204   v7(i,j)=v7(i,j)+v8(ii,2)*v8(jj,2)*v4(ii+lx,jj+lx2)
            v7(j,i)=v7(i,j)
203   continue
201   continue
c now take contrasts:
    do 221 i=1,nx-1
        v8(i,1)=v8(i,1)-v8(i+1,1)
        do 221 j=i,nx-1
            v6(i,j)=v7(i,j)-v7(i+1,j)-v7(i,j+1)+v7(i+1,j+1)
            v6(j,i)=v6(i,j)
221   continue
    call st1(nx-1,v8(1,1),v6,nb,v7,test(1,2))

```

```

c don't overwrite v7, because you need the same matrix in the call to
c st2
c get var-cov matrix:
do 211 i=1,nx
lx=nfx+(i-1)*ny
nyu=ny
if (i.eq.nx) nyu=ny2
do 213 j=i,nx
nyu2=ny
if (j.eq.nx) nyu2=ny2
v6(i,j)=0
lx2=nfx+(j-1)*nx
do 214 ii=1,nyu
do 214 jj=1,nyu2
214 v6(i,j)=v6(i,j)+v8(ii,2)*v8(jj,2)*vv(ii+lx,jj+lx2)
v6(j,i)=v6(i,j)
213 continue
211 continue
c note that the params and weights in v8 are no longer needed:
do 222 i=1,nx-1
do 222 j=i,nx-1
v8(i,j)=v6(i,j)-v6(i+1,j)-v6(i,j+1)+v6(i+1,j+1)
v8(j,i)=v8(i,j)
222 continue
call st2(nx-1,v8,v7,v6,nb,test(1,2),eig(1,2),iwork)
c average y-effect
c (3xx)
c first get weights integrating over x
call intb1(nknx,tknot(-2,1),v8(1,2))
nx2=nb-nx*(ny-1)
do 301 i=1,ny
v8(i,1)=0
ly=nfx+i
nxu=nx
if (i.eq.ny) nxu=nx2
do 302 k=0,nxu-1
302 v8(i,1)=v8(i,1)+v8(k+1,2)*beta(ly+k*ny)
do 303 j=i,ny
nxu2=nx
if (j.eq.ny) nxu2=nx2
v7(i,j)=0
ly2=nfx+j
do 304 ii=0,nxu-1
do 304 jj=0,nxu2-1
304 v7(i,j)=v7(i,j)+v8(ii+1,2)*v8(jj+1,2)*v4(ii*ny+ly,jj*ny+ly2)
v7(j,i)=v7(i,j)
303 continue
301 continue
do 321 i=1,ny-1
v8(i,1)=v8(i,1)-v8(i+1,1)
do 321 j=i,ny-1
v6(i,j)=v7(i,j)-v7(i+1,j)-v7(i,j+1)+v7(i+1,j+1)
v6(j,i)=v6(i,j)
321 continue
call st1(ny-1,v8(1,1),v6,nb,v7,test(1,3))
c remember, v7 needed later
c get var-cov matrix:
do 311 i=1,ny
ly=nfx+i
nxu=nx
if (i.eq.ny) nxu=nx2
do 313 j=i,ny
nxu2=nx

```

```

if (j.eq.ny) nxu2=nx2
v6(i,j)=0
ly2=nfx+j
do 314 ii=0,nxu-1
do 314 jj=0,nxu2-1
314 v6(i,j)=v6(i,j)+v8(ii+1,2)*v8(jj+1,2)*vv(ii*ny+ly,jj*ny+ly2)
v6(j,i)=v6(i,j)
313 continue
311 continue
do 322 i=1,ny-1
do 322 j=i,ny-1
v8(i,j)=v6(i,j)-v6(i+1,j)-v6(i,j+1)+v6(i+1,j+1)
v8(j,i)=v8(i,j)
322 continue
call st2(ny-1,v8,v7,v6,nb,test(1,3),eig(1,3),iwork)
c test for interaction
c (4xx)
c 1. y contrasts:
c note: this all assumes that nb=nx*ny-1 (everything but the const
c in splines)
ny2=ny-1
l=0
do 401 i=1,nx
lx=nfx+(i-1)*ny
nyu=ny-1
if (i.eq.nx) nyu=ny2-1
do 402 j=1,nyu
l=l+1
v8(l,1)=beta(lx+j)-beta(lx+j+1)
l2=l-1
do 403 i2=i,nx
lx2=nfx+(i2-1)*ny
nyu2=ny-1
if (i2.eq.nx) nyu2=ny2-1
jk=1
if (i2.eq.i) jk=j
do 403 j2=jk,nyu2
l2=l2+1
v7(l,l2)=v4(lx+j, lx2+j2)-v4(lx+j+1, lx2+j2)-v4(lx+j, lx2+j2+1) +
&v4(lx+j+1, lx2+j2+1)
v7(l2,1)=v7(l,l2)
403 continue
l2=l2+1
v7(l,l2)=v4(lx+j, nfx+nb)-v4(lx+j+1, nfx+nb)
v7(l2,1)=v7(l,l2)
402 continue
401 continue
l=l+1
v8(l,1)=beta(nfx+nb)
v7(l,1)=v4(nfx+nb, nfx+nb)
c x-contrasts of y contrasts:
l=0
do 405 i=1,nx-1
do 406 j=1,ny2
l=l+1
v8(l,1)=v8(l,1)-v8(l+ny2,1)
l2=l-1
do 407 ii=i,nx-1
jk=1
if (ii.eq.i) jk=j
do 407 jj=jk,ny2
l2=l2+1
v6(l,l2)=v7(l,12)-v7(l+ny2,12)-v7(l,l2+ny2)+v7(l+ny2,l2+ny2)

```

```

v6(12,1)=v6(1,12)
407 continue
406 continue
405 continue
call st1(l,v8,v6,nb,v7,test(1,4))
c same thing for var-cov matrix:
ny2=ny-1
l=0
do 411 i=1,nx
lx=nfx+(i-1)*ny
nyu=ny-1
if (i.eq.nx) nyu=ny2-1
do 412 j=1,nyu
l=l+1
l2=l-1
do 413 i2=i,nx
lx2=nfx+(i2-1)*ny
nyu2=ny-1
if (i2.eq.nx) nyu2=ny2-1
jk=1
if (i2.eq.i) jk=j
do 413 j2=jk,nyu2
l2=l2+1
v8(l,12)=vv(lx+j,1x2+j2)-vv(lx+j+1,1x2+j2)-vv(lx+j,1x2+j2+1)+&vv(lx+j+1,1x2+j2+1)
v8(12,1)=v8(l,12)
413 continue
l2=l2+1
v8(l,12)=vv(lx+j,nfx+nb)-vv(lx+j+1,nfx+nb)
v8(12,1)=v8(l,12)
412 continue
411 continue
l=l+1
v8(l,1)=vv(nfx+nb,nfx+nb)
c x-contrasts of y contrasts:
l=0
do 415 i=1,nx-1
do 416 j=1,ny2
l=l+1
l2=l-1
do 417 ii=i,nx-1
jk=1
if (ii.eq.i) jk=j
do 417 jj=jk,ny2
l2=l2+1
v6(l,12)=v8(l,12)-v8(l+ny2,12)-v8(l,12+ny2)+v8(l+ny2,12+ny2)
v6(12,1)=v6(l,12)
417 continue
416 continue
415 continue
call st2(l,v6,v7,v8,nb,test(1,4),eig(1,4),iwork)
100 format(5e15.8)
return
end

subroutine st1(nn,beta,v6,nb,v7,test)
c calcs value of test statistic given beta and inv 2nd deriv matrix
double precision v6(nb,nb),v7(nb,nb),test(2),beta(1)
call pdi2(nn,v6,v7,nb,nb)
test(1)=0
do 27 ii=1,nn
test(2)=0
do 28 jj=1,ii

```

```

28 test(2)=test(2)+v7(jj,ii)*beta(jj)
      test(1)=test(1)+test(2)*test(2)
27 continue
      return
      end

      subroutine st2(nn,v6,v7,v8,nb,test,eig,iwork)
      double precision test(3),eig(1),v6(nb,nb),v7(nb,nb),v8(nb,nb)
      double precision qfg,trace(7)
      integer iwork(1)
c calculates eigenvalues and p-values, given chol decomp q-form (v7)
c and var-cov matrix (v6)
      do 18 i=1,nn
      do 18 j=i,nn
      v8(i,j)=0
      do 19 ii=1,i
      do 19 jj=1,j
19   v8(i,j)=v8(i,j)+v7(ii,i)*v6(ii,jj)*v7(jj,j)
      v8(j,i)=v8(i,j)
18   continue
      call eigen(v8,nn,nb,eig,v6,1,iflg)
      call sortg(eig,iwork,1,nn)
      do 50 ii=1,nn
      iwork(ii)=1
50   v6(ii,1)=0
      test(2)=1-qfg(eig,v6,iwork,nn,0.d0,test(1),
     &5000,1.d-5,iwork(nn+1),trace,efault)
      test(3)=0
      do 35 i=1,nb
35   test(3)=test(3)+eig(i)
      return
      end

      subroutine cesti(nknx,nkny,tknot,np,beta,var,lvar,nex,ney,
     &est,nfx,ucx,ucy,scx,scy)
c designed to be used with output from sphi
c estimates are calculated at all combinations of equally spaced point
c s between xmin, xmax and ymin ymax.
c assumes cubic splines
c On output est contains the x and y ordinates (cols 1 and 2) and
c the lhr estimate (col 3) and its variance (col 4)
      double precision tknot(-2:nknx+3,2),beta(np),var(lvar,np)
      double precision est(nex*ney,4),a(4),b(4),xmin,xmax,ymin,ymax
      double precision dwork(16),ex,ucx,ucy,scx,scy
      integer iwork(16)
      l=0
      nky4=nkny+4
      nkx4=nknx+4
      xmin=tknot(0,1)
      ymin=tknot(0,2)
      xmax=tknot(nknx+1,1)
      ymax=tknot(nkny+1,2)
      do 10 i=1,nex
      ex=xmin+(i-1)*(xmax-xmin)/(nex-1)
      if (ex.gt.xmax) ex=xmax
      call spln(nknx,3,tknot(-2,1),ex,1,intx,a,1)
      do 15 j=1,ney
      l=l+1
      est(l,1)=ex
      est(l,2)=ymin+(j-1)*(ymax-ymin)/(ney-1)
      if (est(l,2).gt.ymax) est(l,2)=ymax
      call spln(nkny,3,tknot(-2,2),est(l,2),1,inty,b,1)
      ll=0

```

```

do 21 ii=intx,intx+3
if (ii.eq.nkx4.and.inty.eq.nkny+1) then
12=2
else
12=3
endif
do 22 jj=inty,inty+12
11=11+1
iwork(11)=nfx+(ii-1)*nky4+jj
dwork(11)=a(ii-intx+1)*b(jj-inty+1)
22 continue
21 continue
est(1,3)=0
est(1,4)=0
do 25 ii=1,11
est(1,3)=est(1,3)+beta(iwork(ii))*dwork(ii)
do 25 jj=1,11
est(1,4)=est(1,4)+dwork(ii)*dwork(jj)*var(iwork(ii),iwork(jj))
25 continue
est(1,1)=scx*est(1,1)+ucx
est(1,2)=scy*est(1,2)+ucy
15 continue
10 continue
return
end

subroutine pencxc(nkx,tx,nky,ty,pen,wkx,wky)
calc pen matrix for 2 dim laplacian penalty
c note that this loops over y within x: ie parameters are
c ordered b(x1,y1),b(x1,y2),...,b(x1,y[nky+4]),b(x2,y1),...
c wky and wkx are workspace of length as given below
    double precision tx(-2:(nkx+4)),ty(-2:(nky+4)),
&pen(nky+4,4,nkx+4,4),wkx(3,nkx+4,4),wky(3,nky+4,4),a1(4),a2(4),
&a3(4),a4(4),tint(7)
    do 4 i=1,3
    do 4 j=1,nkx+4
    do 4 k=1,4
4   wkx(i,j,k)=0
    do 5 i=1,3
    do 5 j=1,nky+4
    do 5 k=1,4
5   wky(i,j,k)=0
    do 10 i=1,nkx+1
c i is looping over intervals, The ith interval contributes to the
c (i,i),(i,i+1),(i,i+2),(i,i+3),(i+1,i+1),...,(i+3,i+3) pairs, since
c B(j) is >0 on I(j), I(j-1), I(j-2), I(j-3)
        call csplcf(nkx,tx,i,i,a1)
        call csplcf(nkx,tx,i+1,i,a2)
        call csplcf(nkx,tx,i+2,i,a3)
        call csplcf(nkx,tx,i+3,i,a4)
        call int1(tint,tx(i),tx(i-1))
        call int2(tint,a1,a1,wkx(1,i,1))
        call int2(tint,a1,a2,wkx(1,i,2))
        call int2(tint,a1,a3,wkx(1,i,3))
        call int2(tint,a1,a4,wkx(1,i,4))
        call int2(tint,a2,a2,wkx(1,i+1,1))
        call int2(tint,a2,a3,wkx(1,i+1,2))
        call int2(tint,a2,a4,wkx(1,i+1,3))
        call int2(tint,a3,a3,wkx(1,i+2,1))
        call int2(tint,a3,a4,wkx(1,i+2,2))
        call int2(tint,a4,a4,wkx(1,i+3,1))
10  continue
c      write (65,100) (((wkx(i,j,k),i=1,3),j=1,nkx+4),k=1,4)

```

```

do 20 i=1,nky+1
call csplcf(nky,ty,i,i,a1)
call csplcf(nky,ty,i+1,i,a2)
call csplcf(nky,ty,i+2,i,a3)
call csplcf(nky,ty,i+3,i,a4)
call int1(tint,ty(i),ty(i-1))
call int2(tint,a1,a1,wky(1,i,1))
call int2(tint,a1,a2,wky(1,i,2))
call int2(tint,a1,a3,wky(1,i,3))
call int2(tint,a1,a4,wky(1,i,4))
call int2(tint,a2,a2,wky(1,i+1,1))
call int2(tint,a2,a3,wky(1,i+1,2))
call int2(tint,a2,a4,wky(1,i+1,3))
call int2(tint,a3,a3,wky(1,i+2,1))
call int2(tint,a3,a4,wky(1,i+2,2))
call int2(tint,a4,a4,wky(1,i+3,1))
20 continue
c      write (66,100) (((wkx(i,j,k),i=1,3),j=1,nkx+4),k=1,4)
100 format (e14.6)
      do 25 i=1,nky+4
      do 25 j=1,4
      do 25 k=1,nkx+4
      do 25 l=1,4
25 pen(i,j,k,l)=0
      do 30 ix=1,nkx+4
      ixu=min(ix+3,nkx+4)-ix+1
      do 32 ixj=1,ixu
      do 35 iy=1,nky+4
      iyu=min(iy+3,nky+4)-iy+1
      do 37 iyj=1,iyu
      pen(iy,iyj,ix,ixj)=wkx(1,ix,ixj)*wky(3,iy,iyj)+  

&2*wkx(2,ix,ixj)*wky(2,iy,iyj)+wkx(3,ix,ixj)*wky(1,iy,iyj)
37 continue
35 continue
32 continue
30 continue
      return
      end

      subroutine int2(tint,a1,a2,b)
      double precision tint(7),a1(4),a2(4),b(3)
c      a1 are the coefficients of B1, a2 the coefficients of B2, on the
current interval. tint is the output from int1
c      first calculate the contribution to B1*B2 term:
      b(1)=b(1)+a1(4)*a2(4)*tint(7)+(a1(3)*a2(4)+a1(4)*a2(3))*tint(6)+  

&(a1(2)*a2(4)+a1(3)*a2(3)+a1(4)*a2(2))*tint(5)+  

&(a1(1)*a2(4)+a1(2)*a2(3)+a1(3)*a2(2)+a1(4)*a2(1))*tint(4)+  

&(a1(1)*a2(3)+a1(2)*a2(2)+a1(3)*a2(1))*tint(3)+  

&(a1(1)*a2(2)+a1(2)*a2(1))*tint(2)+a1(1)*a2(1)*tint(1)
c      contribution to B1'*B2' term:
      b(2)=b(2)+a1(4)*a2(4)*tint(5)*9+(6*a1(3)*a2(4)+6*a1(4)*a2(3))*  

&tint(4)+(3*a1(2)*a2(4)+4*a1(3)*a2(3)+3*a1(4)*a2(2))*tint(3)+  

&(2*a1(2)*a2(3)+2*a1(3)*a2(2))*tint(2)+a1(2)*a2(2)*tint(1)
c      contribution to B1'''B2'' term:
      b(3)=b(3)+a1(4)*a2(4)*tint(3)*36+(12*a1(3)*a2(4)+12*a1(4)*a2(3))*  

&tint(2)+4*a1(3)*a2(3)*tint(1)
      return
      end

      subroutine int1(tint,txu,txl)
c      txu is the knot at the upper limit of the interval
c      txl the knot at the lower limit. Computes the integral of t^j, for
c      j=0,...,6 (stored in tint)

```

```

double precision tint(7),txu,txl
tint(1)=txu-txl
tint(2)=tint(1)*(txu+txl)
tint(3)=tint(2)*(txu+txl)-txu*txl*tint(1)
tint(4)=tint(3)*(txu+txl)-txu*txl*tint(2)
tint(5)=tint(4)*(txu+txl)-txu*txl*tint(3)
tint(6)=tint(5)*(txu+txl)-txu*txl*tint(4)
tint(7)=tint(6)*(txu+txl)-txu*txl*tint(5)
do 20 j=2,7
20 tint(j)=tint(j)/j
return
end

subroutine penh2(nbx,nby,nbt,penm,mrx,mry,mdx,mdy,h,nmin,
&ih,alpha)
c penm is the penalty matrix, nbx and nby are the actual
c # x & y basis functions used, nbt is the total number used.
c in the configuration this is being written for, nbx would be #
c x-knots + 4, (sim for nby), but nbt would be nbx*nby-1 to avoid
c colinearity with the constant term (ie written for cubics with
c everything but the overall constant included in the tensor
c product term)
c h the neg sec deriv matrix
c nmin the entry in h where spline params begin. alpha*p
c is added to the submatrix bounded by h(nm,nm)and h(nm+nbt-1,
c nm+nbt-1)
c mrx and mry are the actual row dimension of x and y parts of penm,
c and mdx and mdy are the # nonzero diags in each poriton
    double precision penm(mry,mdy,mrx,mdx),h(ih,ih),alpha
    nmax=nmin+nbt-1
    do 10 i=1,nbx
    l1=nmin+(i-1)*nby
    call penhg(nby,penm(1,1,i,1),mry,mdy,h,l1,l1,nmax,ih,alpha)
    l1
    l2=l1
    m2=min(nbx,i+mdx-1)
    do 12 j=i+1,m2
    l1+l1
    l2=l2+nby
12    call penhg(nby,penm(1,1,i,1),mry,mdy,h,l1,l2,nmax,ih,alpha)
10    continue
    do 16 i=nmin,nmax-1
    do 16 j=i+1,nmax
16    h(j,i)=h(i,j)
    return
end

subroutine penlk2(nbx,nby,nbt,beta,penm,mrx,mdx,mry,mdy,
&lik,s,alpha)
c given a penalty matrix penm and parameter values beta,
c subtracts penalty terms from liklihood and score
c the beginning index of beta and s (the score) in the
c call should correspond to where the spline terms begin
c ie call penlik(...beta(7)...s(7)... ) if the spline is
c stored consecutively beginning with the 7th component
c nk is the number of knots, nb is the number
c of basis fcns actually in use, alpha is the value of the
c smoothing parameter.
    double precision penm(mrx,mdx,mry,mdy),beta(1),s(1),lik,alpha
    nb2=nby-(nbx*nby-nbt)
    l1=1
    do 10 i=1,nbx
    m1=max(1,i-mdx+1)

```

```

m2=min(nbx,i+mdx-1)
c      nb=nby
l=0
11=1+(i-1)*nby
12=11
do 12 j=i,m2
l=l+1
c      if (j.eq.nbx) nb=nb2
c 11 is the beginning row index, 12 the beginning column index
nbr=nby
nbc=nby
if (i.eq.nbx) nbr=nb2
if (j.eq.nbx) nbc=nb2
call penlkg(nbr,nbc,beta(12),beta(11),penm(1,1,i,1),mry,mdy,lik,
&s(11),alpha)
12=12+nby
12 continue
if (m1.lt.i) then
  l=i-m1+2
  do 11 j=m1,i-1
12=1+(j-1)*nby
  l=l-1
nbr=nby
nbc=nby
if (i.eq.nbx) nbr=nb2
if (j.eq.nbx) nbc=nb2
call penlkg(nbr,nbc,beta(12),beta(11),penm(1,1,j,1),mry,mdy,lik,
&s(11),alpha)
11 continue
endif
10 continue
return
end

subroutine penhg(nb,penm,mr,md,h,nm1,nm2,max,ih,alpha)
c penm is the penalty matrix, nb the actual
c # basis functions used, h the neg sec deriv matrix
c nm1 the entry in h where spline params begin. alpha*p
c is added to the submatrix bounded by h(nm1,nm2) and h(nm1+nb-1,
c nm2+nb-1)
c mr is the actual row dimension of penm, md is the # nonzero
c diagonals in the penalty matrix
c penm must be symmetric stored in the appropriate banded format
double precision penm(mr,md),h(ih,ih),alpha,u
do 10 i=1,nb
  i1=nm1+i-1
  i2=nm2+i-1
  if (i1.le.max.and.i2.le.max) then
    h(i1,i2)=h(i1,i2)+alpha*penm(i,1)
  endif
  if (i.lt.nb) then
    m2=min(nb,i+md-1)
    l=1
    do 12 j=i+1,m2
      j1=nm2+j-1
      j2=nm1+j-1
      l=l+1
      u=alpha*penm(i,1)
      if (i1.le.max.and.j1.le.max) then
        h(i1,j1)=h(i1,j1)+u
      endif
      if (j2.le.max.and.i2.le.max) then
        h(j2,i2)=h(j2,i2)+u
      endif
    12 continue
  endif
10 continue
return
end

```

```

        endif
12      continue
        endif
10      continue
        return
        end

        subroutine penlkg(nb1,nb2,beta,beta2,penm,mr,md,lik,s,alpha)
c given a penalty matrix penm and parameter values beta,
c subtracts penalty terms from liklihood and score
c the beginning index of beta and s (the score) in the
c call should correspond to where the spline terms begin
c ie call penlik(...beta(7)...s(7)...) if the spline is
c stored consecutively beginning with the 7th component
c nk is the number of knots, nb1 is the number or rows of penm
c nb2 the # cols (in unbande form) acutally used
c alpha is the value of the smoothing parameter.
c beta2 added to allow off center contributions. That is, the beta
c parameters are used for the score and the complicated part of the
c liklihood, but multiplied by beta2 to get the likelihood contribution
c ie like contri = beta2' penm beta
        double precision penm(mr,md),beta(1),s(1),lik,alpha,u,beta2(1)
        do 10 i=1,nb1
          u=0
          m1=max(1,i-md+1)
          m2=min(nb2,i+md-1)
          l=0
          do 12 j=i,m2
            l=l+1
            u=u+penm(i,l)*beta(j)
12      continue
        if (m1.lt.i) then
          l=i-m1+2
          do 11 j=m1,i-1
            l=l-1
            u=u+beta(j)*penm(j,l)
11      continue
        endif
        s(i)=s(i)-alpha*u
        lik=lik-alpha*u*beta2(i)/2
10      continue
        return
        end

        subroutine csplcf(nknot,t,j,k,a)
        double precision t(-2:(nknot+3)),a(0:3),a1,b1,c1,a2,b2,c2,t1,t2
c t is the augmented knot sequence, nknot the number of knots
c on interval I(k) the cubic spline basis function B(j) can be
c represented as B(j)(t)=a(3)t^3+a(2)t^2+a(1)t+a(0)
c this routine calculates a(3),a(2),a(1),a(0)
c     call dblepr("knots",5,t(-2),nknot+6)
        if (k.gt.j.or.k.lt.j-3) then
          a(0)=0
          a(1)=0
          a(2)=0
          a(3)=0
        else if (k.eq.j-3) then
          b1=1/(t(j-3)-t(j-4))
          a1=-t(j-4)*b1
          b2=1/(t(j-2)-t(j-4))
          a2=-t(j-4)*b2
          c1=b1*b2
          b1=b1*a2+b2*a1

```

```

a1=a1*a2
b2=1/(t(j-1)-t(j-4))
a2=-t(j-4)*b2
a(3)=c1*b2
a(2)=c1*a2+b1*b2
a(1)=b1*a2+a1*b2
a(0)=a1*a2
else if (k.eq.j-2) then
b1=-1/(t(j-2)-t(j-3))
a1=-t(j-2)*b1
b2=-b1
a2=t(j-3)*b1
t1=1/(t(j-2)-t(j-4))
t2=-1/(t(j-1)-t(j-3))
c1=b1*t1+b2*t2
b1=-b1*t(j-4)+t1*a1-b2*t(j-1)*t2+t2*a2
c
b1=-(t(j-4)+t(j-2))*b1*t1-(t(j-1)+t(j-3))*t2*b2
a1=-t(j-4)*t1*a1-t(j-1)*t2*a2
c2=-t2*b2
b2=-t2*a2+b2*t2*t(j-3)
a2=a2*t2*t(j-3)
t1=1/(t(j-1)-t(j-4))
t2=-1/(t(j)-t(j-3))
a(3)=t1*c1+t2*c2
a(2)=t1*b1-t1*c1*t(j-4)+t2*b2-t2*c2*t(j)
a(1)=t1*a1-t1*b1*t(j-4)+t2*a2-t2*b2*t(j)
a(0)=-a1*t1*t(j-4)-t2*t(j)*a2
else if (k.eq.j-1) then
b1=-1/(t(j-1)-t(j-2))
a1=-t(j-1)*b1
b2=-b1
a2=-b2*t(j-2)
t1=1/(t(j-1)-t(j-3))
t2=-1/(t(j)-t(j-2))
c2=t2*b2+t1*b1
b2=-t(j)*b2*t2+a2*t2-t1*t(j-3)*b1+a1*t1
a2=-t(j)*t2*a2-t(j-3)*t1*a1
c1=-t1*b1
b1=t1*t(j-1)*b1-t1*a1
a1=t1*t(j-1)*a1
t1=1/(t(j-1)-t(j-4))
t2=-1/(t(j)-t(j-3))
a(3)=t1*c1+t2*c2
a(2)=t1*b1-t1*c1*t(j-4)+t2*b2-t2*c2*t(j)
a(1)=t1*a1-t1*b1*t(j-4)+t2*a2-t2*b2*t(j)
a(0)=-a1*t1*t(j-4)-t2*t(j)*a2
else if (k.eq.j) then
b1=-1/(t(j)-t(j-1))
a1=-t(j)*b1
t1=-1/(t(j)-t(j-2))
c1=t1*b1
b1=t1*a1-b1*t1*t(j)
a1=-t1*t(j)*a1
t1=-1/(t(j)-t(j-3))
a(3)=t1*c1
a(2)=-c1*t1*t(j)+t1*b1
a(1)=-t1*t(j)*b1+t1*a1
a(0)=-t1*t(j)*a1
endif
return
end
subroutine intb1(nk,tk,o)
c subroutine to calculate the integral of cubic spline

```

```

c basis functions
c nk is the number of knots
    double precision tk(-2:nk+3),o(nk+4),a(4)
    do 10 i=1,nk+4
        o(i)=0
        kb1=max(1,i-3)
        kb2=min(i,nk+1)
c note: ith basis function defined on intervals kb1 to kb2
    do 11 j=kb1,kb2
        call csplcf(nk,tk,i,j,a)
c note: a(4) is the coef of the cubic term ... a(1) the coef of the const
    do 12 k=1,4
12   o(i)=o(i)+a(k)*(tk(j)**k-tk(j-1)**k)/k
11   continue
10   continue
    return
    end

    subroutine stvc(nov,s,c,x,tknot,times,beta,lik0,lik,sb,vv,
&iwork,dsub,isub,testr,eig,v2i,eps,z,exz,s0,s1,ts1,w,
&wp,ts1p,id,s1p,z1)
    double precision s(1),c(1),x(1),tknot(1),times(1),beta(1),lik0
    double precision lik(2),sb(1),vv(1),dsub(1),testr(1),eig(1)
    double precision uplik,v2i(1)
C*****
    double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C*****
C>>>
    double precision wp(1),ts1p(1),s1p(1)
    integer id(1)
C<<<
    integer nov(15),iwork(1),isub(1)
    common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
    &nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
    common /up/ uplik
c
c nov parameters:
c   1=no, 2=nrx(=no), 3=ncx, 4=istr (col # of x for strata, -1 if none,
c   5=maxiter, 6=knot option (1 use knots provided, 0 program calculates),
c   7=smoothing option (<0 use input smoothing params, 0 calc smoothing
c   param only in first iteration, >0 recalc after each iteration)
c   8=nfx (nfx+nsx in sph, nfx is #col in linear.cov, nsx in spline.cov),
c   9=nsx, 10=analysis option (<=0 estimates only, 1 est & var, >1 est,
c   var & test), 12=order of spline, 13=nest, 14=nknot, 15=nitime (nstimes
c   the number of points where covariates can switch values, >no for all
c   failure times)
c
c dsub must have length at least <<see below>>
c iwork must have length at least 3*no+max(no,np)+1+2*nstr+2*ntud*nstr
c where nstr is the # strata & ntud is either ntime+1, or 1 if
c ntime>no. Note that nsx is in nov(9)
c isub must have length 8+nsx+nfx+(n or ntud)
c testr must have length nsx*6
c eig: length nk4*2*nsx, will have eigenvalues for tests (output)
c if smoothing parameters are input need to be in components
c   nsx*2+1 to nsx*3 of dsub
c nov(12) is the order of the spline: 3 for cubic with squared second
c derivative penalty, 2 for quad & 1st deriv penalty, & 0 for constant
c with 1st diff penalty
c nov(13) is the number of point estimates when ntime>no
c estimates stored in dsub, which must be at least nest*(2*nsx+1) or
c (2*ntime+2)*(2*nsx+1) beginning index will be dsub(isub(5))
c isub(5) is either 3*nsx+iord1*no+1 or 3*nsx+iord1*(ntime+1)+1

```

```

c v2i will have the inverse 2nd deriv matrix
c
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of splines
c isub(5) reserved for the beginning index (in dsub) of penm
c isub(6) reserved for the beginning index (in dsub) of unpenalized inf
c isub(7) reserved for number of knots
c isub(8) needs to be row dim of matrix of splines
c isub(8+1) to isub(8+nfx) gives col #'s of fixed cova
c isub(8+1+nfx) to isub(8+nfx+nsx) gives col #'s of spline cova
c isub(8+1+nfx+nsx) will be the beginning index of inter.
c
c ntime is as in coxrg: #switch points, not # intervals
c
c dsub(1:nsx) are target df.
c dsub(nsx+1:2*nsx) will be attained df.
c dsub(2*nsx+1:3*nsx) will be attained df.
c
c     calls stvc1, calvar, testw, cestb, cest directly
c     indirect calls: coxrg, coxft, coxg, strat, ut, uft, tint, tint4
c     cholg, solve sortg dperm sub5, sub5c, sub6c, eval5, act5, eval6, act6,
c     cks, gspl1, gspl2, pencb, penql, penco, extrct,
c     matop5, adisw, spln, penh, gdf, degf2, penlik, eigen,
c     qfg, pdi, pdi2
c
c
no=nov(1)
nrx=nov(2)
ncx=nov(3)
istr=nov(4)
nknot=nov(14)
ntime=nov(15)
mit=nov(5)
knopt=nov(6)
ksmopt=nov(7)
nfx=nov(8)
nsx=nov(9)
iord=nov(12)
nest=nov(13)
malph=2*nsx+1
if (iord.ne.0.and.iord.ne.2.and.iord.ne.3) iord=0
iord1=iord+1
itk=-iord+1
ncp=iord1
if (iord.eq.0) then
ntime=nknot
ncp=2
endif
nbas=nknot+iord1
nk2=nknot+2
nk3=nknot+3
nk4=nknot+4
np=nfx+nsx*nbas
nkk=max(nbas,np-nbas)
kmd=(np-nbas)*(np-nbas)
kmd=max(kmd,nbas*nbas+2*nbas*nkk)
c if ntime>no & ksmopt>0 then length(dsub)=3*nsx (df,smp)
c +iord1*no (B-spline terms) (isub(4))
c +ncp*nk4 (pen matrix) (isub(5))
c +np*np (unpen inf) (isub(6))
c +max((np-nbas)*(np-nbas),(nbas*nbas)+2*(nbas*max(nbas,np-nbas)))
c +np*np+no+5*np (work)
c if ntime>no and ksmopt<=0 then length(dsub)=3*nsx (df,smp)
c +iord1*no (B-spline terms) (isub(4))

```

```

c +ncp*nk4 (pen matrix) (isub(5))
c +np*np (unpen inf) (isub(6))
c +nbas*nbas+
c +max(no,np*np+4*np) (work)
c if ntime<no & ksmopt>0 then length(dsub)=3*nsx (df,smp)
c +iord1*(ntime+1) (B-spline terms) (isub(4))
c +ncp*nk4 (pen matrix) (isub(5))
c +np*np (unpen inf) (isub(6))
c +max((np-nbas)*(np-nbas),(nbas*nbas)+2*(nbas*max(nbas,np-nbas)))
c +max(no,np*np+4*np) (work)
c if ntime<no and ksmopt<=0 then length(dsub)=3*nsx (df,smp)
c +iord1*(ntime+1) (B-spline terms) (isub(4))
c +ncp*nk4 (pen matrix) (isub(5))
c +np*np (unpen inf) (isub(6))
c +nbas*nbas+
c +max(no,np*np+4*np) (work)
c also, to accomodate est, must be at least 3*nsx+iord1*no+
c nest*(2*nsx+1) (ntime>no) or
c (2*ntime+2)*(2*nsx+1) (ntime<no)
    if (ntime.gt.no) then
        if (ksmopt.gt.0) then
            mdw=3*nsx+iord1*no+np*np+nk4*ncp+kmd+1
        else
            mdw=3*nsx+iord1*no+np*np+nk4*ncp+nbas*nbas+1
        endif
    else
        if (ksmopt.gt.0) then
            mdw=3*nsx+iord1*(ntime+1)+np*np+nk4*ncp+kmd+1
        else
            mdw=3*nsx+iord1*(ntime+1)+np*np+nk4*ncp+nbas*nbas+1
        endif
    endif
    call stvc1 (s,c,x,istr,isub,tknot,ntime,times,
&beta,lik0,lik(1),sb,vv,iwork,dsub(mdw),dsub(1),eps,icnv,
&z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
    if (icnv.gt.0) then
        vv(1)=-icnv
        return
    endif
    lik(2)=uplik
    nov(15)=ntime
    nov(14)=nknot
c after call to calvar,
c done with dsub from isub(5) on out:
    mdw2=isub(5)
    mdw3=mdw2+nbas*nbas
    mdw4=mdw3+nbas*nbas
c (note:index inform in iwork partially overwritten in testw call)
    if (nov(10).gt.0) then
        call calvar(vv,np,np,dsub(isub(6)),np,vv,np,v2i)
    endif
    if (nov(10).gt.1) then
        call testw(np,nbas,nfx,nsx,beta,vv,np,v2i,np,
&dsub(mdw2),dsub(mdw3),testr,iwork,dsub(mdw4),eig)
    endif
    if (ntime.gt.no) then
        call cest(nknot,tknot,np,beta,vv,np,tknot(3),tknot(nknot+4),
&nest,dsub(isub(5)),nfx,nsx,iord)
    else
        call cestb(ntime,times,tknot(3),tknot(nknot+4),isub(isub(3)),
&dsub(isub(4)),dsub(isub(5)),nfx,nsx,beta,vv,np,np,nknot,iord)
    endif
    return

```

```

    end
c
c for fitting time varying coefficient models
c file contains stvc1 gsp12 gsp11 eval5 eval6 evalc act5 act6 sub6c
c sub5 sub5c
c
c in addition, routines in this file call sortg, dperm, tint4, cks,
c   strat, tint, pencb, penql, penco, coxrg, spln, adisw, extrct, cholg,
c   matop5, gdf, penh, penlik
    subroutine stvc1(s,c,x,istr,isub,tknot,ntime,
    &times,beta,lik0,lik,sb,vv,iwork,work,dsub,eps,icnv,
    &z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
    double precision vv(1),sb(1),work(1),beta(1),times(1),s(no)
    double precision dsub(1)
    double precision c(no),x(nrx,ncx),eps,lik,lik0,tknot(-2:nk3)
C*****
    double precision z(1),exz(1),s0(1),s1(1),ts1(1),w(1),z1(1)
C*****
C>>>>
    double precision wp(1),ts1p(1),s1p(1)
    integer id(1)
C<<<<
    integer iwork(1),isub(1)
    external eval5,eval6,act5,act6,sub5
    common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
    &nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
c isub(3) reserved for the beginning index (in isub) of inter
c isub(4) reserved for the beginning index (in dsub) of splines
c isub(5) reserved for the beginning index (in dsub) of penm
c isub(6) reserved for the beginning index (in dsub) of unpenalized inf
c isub(7) reserved for number of knots
c isub(8) needs to be row dim of matrix of splines
c isub(8+1) to isub(8+nfx) gives col #'s of fixed covs
c isub(8+1+nfx) to isub(8+nfx+nsx) gives col #'s of spline covs
c isub(8+1+nfx+nsx) will be the beginning index of inter.
c
c ntime is as in coxrg: #switch points, not # intervals
c
c dsub(1:nsx) are target df.
c dsub(nsx+1:2*nsx) will be attained df.
c
c first get knot locations:
c
    if (knopt.le.0.or.ntime.lt.no) then
        do 5 i=1,no
      5   iwork(i)=i
        call sortg(s,iwork,1,no)
        call dperm(c,iwork,1,no,work)
        do 6 j=1,ncx
      6   call dperm(x(1,j),iwork,1,no,work)
    endif
    knkp=nknot
    if (knopt.le.0) then
        ntud=nknot+1
        call tint4(ntud,s,c,tknot(1),no,iwork)
    if (iord.eq.0) then
        ntime=ntud-1
        do 717 kk=1,ntime
          times(kk)=tknot(kk)
    717    endif
        if (knkp.ne.ntud-1) then
            nknot=ntud-1
            nbas=nknot+iord1

```

```

nk2=nknot+2
nk3=nknot+3
nk4=nknot+4
np=nfx+nsx*nbas
endif
call cks(s(1),s(no),tknot(itk),nknot,iord)
endif
c      write (6,*) (tknot(i),i=-2,(nknot+3))
c
c If ntime>no then spline functions will change at all failure times,
c otherwise, only at ntime times.
c
c this repeats some code from coxrg, but need to determine this first
c so splines can be calculated before calling coxrg.
c
if (ntime.lt.no) then
ntud=ntime+1
if (times(ntime).le.0) then
    call tint4(ntud,s,c,times,no,iwork)
c note: if ntud gets altered in tint4 (because not enough failures
c then need to reset ntime to match.
    ntime=ntud-1
endif
c      write (37,*) (times(i),i=1,ntime)
else
ntud=1
endif
mnstr=1
miuti=mnstr+1
mtduti=miuti+2*no
c miwk needs length max(no,np)
miwk=mtduti+no
mnostr=miwk+max(no,np)
call strat(istr,s,c,x,iwork(mnostr),nstr,iwork(miwk),work,
&no,ncx,nrx,iwork(miuti),iwork(mtduti))
c      write (6,*) nstr
iwork(mnstr)=nstr
mnuti=mnostr+nstr
mitud=mnuti+nstr
c      write (6,*) (iwork(i),i=mnostr,mnostr+nstr-1)
c      write (6,*) (iwork(i),i=mnuti,mnuti+nstr-1)
c      write (6,*) ntime
if (ntime.lt.no) then
    l1=1
    do 225 l1=1,nstr
    l2=(l1-1)*2*ntud
    call tint(ntud,times,s,c,iwork(mitud+l2),no,iwork(miuti),
& iwork(mnuti+l1-1),l1)
    l1=iwork(mnostr+l1-1)+1
225      continue
endif
isub(3)=9+nfx+nsx
if (ksmopt.ge.0) then
do 789 i=1,nsx
789      dsub(i+2*nsx)=.1
endif
isub(4)=3*nsx+1
if (ntime.gt.no) then
    isub(5)=isub(4)+iord1*no
    isub(8)=no
    call gspl1(tknot,s,nstr,iwork(miuti),iwork(mnuti),
& iwork(mnostr),isub(isub(3)),dsub(isub(4)))
else

```

```

        isub(5)=isub(4)+iord1*ntud
        isub(8)=ntud
        call gspl2(tknot,s,nstr,iwork(miuti),iwork(mnuti),
& iwork(mtduti),iwork(mnostr),ntud,iwork(mitud),isub(isub(3)),
& dsub(isub(4)))
        endif
c isub(6) is the beginning index of the unpenalized inf matrix.
        isub(6)=ncp*nk4+isub(5)
        if (iord.eq.3) then
            call pencb(nknot,tknot(itk),dsub(isub(5)),nk4)
        else if (iord.eq.2) then
            call penql(nknot,tknot(itk),dsub(isub(5)),nk4,iord)
        else
            call penco(nknot,tknot(itk),dsub(isub(5)),nk4)
        endif
c      write (6,*) np
c      write (6,*) (isub(i),i=1,isub(3))
        do 28 i=1,np
28    beta(i)=0
        if (ntime.gt.no) then
            call coxrg(s,c,x,no,nrx,ncx,istr,np,mit,1,ntime,times,beta,
& lik0,lik,sb,vv,iwork,work,act5,eval5,sub5,isub,dsub,eps,icnv,
& z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
        else
            call coxrg(s,c,x,no,nrx,ncx,istr,np,mit,1,ntime,times,beta,
& lik0,lik,sb,vv,iwork,work,act6,eval6,sub5,isub,dsub,eps,icnv,
& z,exz,s0,s1,ts1,w,wp,ts1p,id,s1p,z1)
        endif
        return
        end

        subroutine gspl2(tk,s,nstr,iuti,nuti,tduti,nostr,ntud,
&itud,inter,tt)
        double precision s(no),tk(-2:nk3),tt(ntud,iord1),b(4),smin,smax
        integer iuti(no,2),nostr(nstr),nuti(nstr),inter(ntud),tduti(no)
        integer itud(ntud,2,nstr)
        common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
        do 10 i=1,ntud
10    find midrange of failures in time interval
        iflg=0
        do 15 j=1,nstr
15    l1=itud(i,1,j)
        l2=itud(i,2,j)
        if (l1.gt.0) then
            if (iflg.eq.0) then
                iflg=1
                smin=s(iuti(l1,1))
                smax=s(iuti(l2,1))
            else
                smin=min(smin,s(iuti(l1,1)))
                smax=max(smax,s(iuti(l2,1)))
            endif
        endif
15    continue
        smin=(smin+smax)/2
        call spln(nknot,iord,tk(itk),smin,1,inter(i),b,1)
        do 725 k=1,iord1
725    tt(i,k)=b(k)
10    continue
        return
        end

```

```

subroutine gspl1(tk,s,nstr,iuti,nuti,nostr,inter,tt)
double precision s(no),tk(-2:(nk3)),tt(no,iord1),b(4)
integer iuti(no,2),nostr(nstr),nuti(nstr),inter(no)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
nstr=1
do 724 ll=1,nstr
do 723 j=nstr,nuti(ll)
l2=iuti(j,1)
call spln(nknot,iord,tk(itk),s(l2),1,inter(l2),b,1)
do 725 k=1,iord1
725 tt(l2,k)=b(k)
723 continue
nstr=nostr(ll)
724 continue
return
end

subroutine eval6(g,gp,b,lci,i,l,s,c,xx,nrx2,ncx2,no2,np2,lc2,nact,
&iact,isub,dsub)
double precision gp(np),s(no),c(no),xx(nrx,ncx),dsub(1),b(np),g
integer isub(1),iact(1)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
call evalc(g,gp,b,isub(9),isub(9+nfx),nact,iact,
&i,lci,xx,dsub(isub(4)),isub(8))
return
end

subroutine eval5(g,gp,b,lci,i,l,s,c,xx,nrx2,ncx2,no2,np2,lc2,nact,
&iact,isub,dsub)
double precision gp(np),s(no),c(no),xx(nrx,ncx),dsub(1),b(np),g
integer isub(1),iact(1)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
call evalc(g,gp,b,isub(9),isub(9+nfx),nact,iact,
&i,l,xx,dsub(isub(4)),isub(8))
return
end

subroutine evalc(g,gp,b,ifx,lsx,nact,iact,i,l,x,tt,ntt)
c in this routine, i is the row of x to use, and l the row of tt. ntt
c is the row dim of tt
double precision gp(1),tt(ntt,4),x(nrx,ncx),b(np),g
integer iact(1),lsx(nsx),ifx(nfx)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
do 20 ii=1,np
20 gp(ii)=0
g=0
do 12 j=1,nfx
    gp(j)=x(i,ifx(j))
    g=g+gp(j)*b(j)
12 continue
do 13 j=1,nsx
    isp=iact(nfx+(j-1)*iord1+1)-1
    do 14 jj=1,iord1
        gp(isp+jj)=x(i,lsx(j))*tt(l,jj)
        g=g+gp(isp+jj)*b(isp+jj)
14 continue
13 continue
100 format(6e12.4)
return

```

```

    end

    subroutine act6(lci,11,12,iuti,no2,s,c,x,nrx2,ncx2,iact,nact,
    &np2,lc2,isub,dsub)
    double precision s(no),c(no),x(nrx,ncx),dsub(1)
    integer iuti(no,2),iact(1),isub(1)
    common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
    &nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
    isp=isub(isub(3)+lci-1)
    do 10 i=1,nfx
10  iact(i)=i
    nact=nfx
    do 11 j=1,nsx
    isp2=nfx+nbas*(j-1)+isp
    do 12 jj=isp2,isp2+iord
    nact=nact+1
12  iact(nact)=jj
11  continue
    return
    end

    subroutine act5(lci,11,12,iuti,no2,s,c,x,nrx2,ncx2,iact,nact,
    &np2,lc2,isub,dsub)
    double precision s(no),c(no),x(nrx,ncx),dsub(1)
    integer iuti(no,2),iact(1),isub(1)
    common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
    &nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
    isp=isub(isub(3)+iuti(11,1)-1)
    do 10 i=1,nfx
10  iact(i)=i
    nact=nfx
    do 11 j=1,nsx
    isp2=nfx+nbas*(j-1)+isp
    do 12 jj=isp2,isp2+iord
    nact=nact+1
12  iact(nact)=jj
11  continue
    return
    end

    subroutine sub6c(beta,lik,sb,vv,penm,v3,dft,dfa,
    &alpha,v6,v7,nit)
    double precision beta(np),lik,sb(np),vv(np,np),v3(np,np)
    double precision penm(nk4,ncp),dft(nsx),dfa(nsx),alpha(nsx)
    double precision v6(1),mindf,v7(1)
    common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
    &nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
c v6 and v7 must have length at least nk4*nk4
c first update values of smoothing parameters to give required df.
    if ((nit.eq.1.and.ksmopt.eq.0).or.(ksmopt.gt.0.and.nit.lt.9.and.
    &nit.gt.0)) then
      do 819 j=1,nsx
      l1=nfx+1+(j-1)*nbas
      l2=l1+nbas-1
      call adisw(l1,l2,np,vv,v3,np)
      dfa(j)=dft(j)
c 3rd from last arg in gdf # dcols used
c in pen matrix, 4th from last mindf (2 for cubic spline with 2nd deriv
c penalty)
      mindf=2.d0
      if (iord.lt.3) mindf=1.d0
      call gdf(nbasis,penm,v3,v3,v6,np,dfa(j),alpha(j),v7,mindf,
    &ncp,nk4,iopp1)

```

```

c
c   if algorithm failed, return without adding penalty
c
c       if (iopp1.gt.0) return
c       write (6,100) dft(j),dfa(j),alpha(j)
819  continue
      endif
c then copy vv to v3:
      do 10 i=1,np
      do 10 j=1,np
      v3(i,j)=vv(i,j)
10   continue
c then add penalty terms to lik,score, and inf (lik,sb,vv)
      do 535 j=1,nsx
      l1=nfx+1+(j-1)*nbas
      call penlik(nbabs,beta(l1),penm,nk4,ncp,lik,sb(l1),alpha(j))
      call penh(nbabs,penm,nk4,ncp,vv,l1,np,alpha(j))
535  continue
100  format (3e15.8)
      return
      end

subroutine sub5(np2,beta,lik,sb,vv,vvi,isub,dsub,nit)
double precision beta(np),lik,sb(np),vv(np,np),dsub(1)
double precision vvi(np,np),uplik
integer isub(1)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
common /up/ uplik
c to provide working matrices, dsub must be at least
c 3*nsx+4*(n or ntud,depending on ntime)+np*np+2*(nknot+4)**2
      mv6=isub(6)+np*np
c      write (6,*) lik
      uplik=lik
c      call dblepr("lik",3,lik,1)
      if (ksmopt.gt.0.and.nsx.gt.1.and.nit.lt.9.and.nit.gt.0) then
      call sub5c(beta,lik,sb,vv,dsub(isub(5)),dsub(isub(6)),
&dsub(1),dsub(1+nsx),dsub(malph),dsub(mv6),vvi)
      else
      call sub6c(beta,lik,sb,vv,dsub(isub(5)),dsub(isub(6)),
&dsub(1),dsub(1+nsx),dsub(malph),dsub(mv6),vvi,nit)
      endif
c      write (6,*) lik
c      call dblepr("plik",4,lik,1)
110  format (e20.12)
      return
      end

subroutine sub5c(beta,lik,sb,vv,penm,v3,dft,dfa,alpha,v6,vvi)
double precision beta(np),lik,sb(np),vv(np,np),v3(np,np)
double precision penm(nk4,ncp),dft(nsx),dfa(nsx),alpha(nsx)
double precision v6(1),mindf,vvi(np,np)
common /paramt/ no,nrx,ncx,nknot,mit,knopt,ksmopt,np,nfx,nsx,
&nk2,nk3,nk4,malph,iord,iord1,itk,ncp,nbas
c v6 must have length at least nbas*nbas+2*(max(nbabs,np-nbas))*nbas
c first update values of smoothing parameters to give required df.
      mv8=1+nbabs*nbas
      nkk=max(nbabs,np-nbas)
      mv9=mv8+nbabs*nkk
      do 820 i=1,np
      do 820 j=1,np
820  vvi(i,j)=vv(i,j)
      do 821 j=2,nsx

```

```

    l1=nfx+1+(j-1)*nbas
    call penh(nbasis,penm,nk4,ncp,vvi,l1,np,alpha(j))
821  continue
    do 819 j=1,nsx
    l1=nfx+1+(j-1)*nbas
    l2=l1+nbas-1
    call extrct(vvi,np,np,l1-1,l2+1,v6,np-nbas)
    call cholg(np-nbas,v6,v3,np-nbas,np-nbas)
    call matop5(vv,np,v3,np-nbas,l1,l2,v6(mv8),nbasis,v6(mv9),v6,nkk)
    dfa(j)=dft(j)

c 3rd from last arg in gdf # dcols used
c in pen matrix, 4th from last mindf (2 for cubic spline with 2nd deriv
c penalty)
    mindf=2.d0
    if (iord.lt.3) mindf=1.d0
    call gdf(nbasis,penm,v6(mv9),v6(mv8),v6,nkk,dfa(j),alpha(j),
    &v3,mindf,ncp,nk4,iopp1)

c
c   if algorithm failed, return without adding penalty
c
    if (iopp1.gt.0) return
    write (6,100) dft(j),dfa(j),alpha(j)
    call penh(nbasis,penm,nk4,ncp,vvi,l1,np,alpha(j))
    if (j.lt.nsx) then
    call penh(nbasis,penm,nk4,ncp,vvi,l1+nbasis,np,-alpha(j+1))
    endif
819  continue
c then copy vv to v3:
    do 10 i=1,np
    do 10 j=1,np
    v3(i,j)=vv(i,j)
10  vv(i,j)=vvi(i,j)
c then add penalty terms to lik,score, and inf (lik,sb,vv)
    do 535 j=1,nsx
    l1=nfx+1+(j-1)*nbas
    call penlik(nbasis,beta(l1),penm,nk4,ncp,lik,sb(l1),alpha(j))
c      call penh(nbasis,penm,nk4,ncp,vv,l1,np,alpha(j))
535  continue
100  format (3e15.8)
    return
    end

    subroutine testw(np,nk4,nfx,nsx,beta,vv,lvv,v4,lv4,v6,v7,test,
    &iwork,v8,eig)
c For a spline with standard parametrization (all B-spline basis
c fcns included), tests (Wald) that all coefs=0 and that all coefs are
c equal. nk4 is the number of basis fcns per spline term. nfx the
c number of nonspline coeffs (come first in beta) and nsx the number
c of spline terms (parameters assumed to be consecutive in beta)
c so np=nfx+nsx*nk4
c on input v4 is the inverse penalized 2nd derivative matrix, and vv the
c estimated var-cov matrix of the parameters estimates.
c length(iwork)=2*nk4
c v6,v7,v8 are work space,
c on output eig will contain the eigenvalues and test the test
c results
c calls pdi2, sortg, eigen qfg from ~/src/util.a
    double precision vv(lvv,np),beta(np),v4(lv4,np)
    double precision v6(nk4,nk4),v7(nk4,nk4),test(6,nsx)
    double precision v8(nk4,nk4),qfg,trace(7),eig(nk4,2*nsx)
    integer iwork(1)
    do 50 k=1,nsx
    l1=nfx+(k-1)*nk4

```

```

c overall test for var i:
c first calc stat:
    do 15 i=1,nk4
    do 15 j=1,nk4
    v6(i,j)=v4(i+11,j+11)
15  continue
    call pdi2(nk4,v6,v7,nk4,nk4)
    test(1,k)=0
    do 27 ii=1,nk4
    test(2,k)=0
    do 28 jj=1,ii
28  test(2,k)=test(2,k)+v7(jj,ii)*beta(l1+jj)
    test(1,k)=test(1,k)+test(2,k)*test(2,k)
27  continue
c then get eigenvalues to calc dist
    do 17 i=1,nk4
    do 17 j=1,nk4
17  v6(i,j)=vv(i+11,j+11)
    do 18 i=1,nk4
    do 18 j=i,nk4
    v8(i,j)=0
    do 19 ii=1,i
    do 19 jj=1,j
19  v8(i,j)=v8(i,j)+v7(ii,i)*v6(ii,jj)*v7(jj,j)
    v8(j,i)=v8(i,j)
18  continue
    call eigen(v8,nk4,nk4,eig(1,2*k-1),v6,1,iflg)
    call sortg(eig(1,2*k-1),iwork,1,nk4)
    do 250 ii=1,nk4
    iwork(ii)=1
250 v6(ii,1)=0
    test(2,k)=1-qfg(eig(1,2*k-1),v6,iwork,nk4,0.d0,test(1,k),
&5000,1.d-5,iwork(nk4+1),trace,efault)
    test(3,k)=0
    do 35 i=1,nk4
35  test(3,k)=test(3,k)+eig(i,2*k-1)
c test for proportional hazards:
c note that eigenvalues the same as the first nk4-1 eigs used
c before. (actually not always)
    nk3=nk4-1
    do 215 i=1,nk4
    do 215 j=1,nk4
215 v6(i,j)=v4(i+11,j+11)
    do 217 i=1,nk3
    do 217 j=1,nk4
217 v6(i,j)=v6(i,j)-v6(i+1,j)
    do 218 j=1,nk3
    v8(j,1)=beta(l1+j)-beta(l1+j+1)
    do 218 i=1,nk3
218 v6(i,j)=v6(i,j)-v6(i,j+1)
    call pdi2(nk3,v6,v7,nk4,nk4)
    test(4,k)=0
    do 227 ii=1,nk3
    test(5,k)=0
    do 228 jj=1,ii
228 test(5,k)=test(5,k)+v7(jj,ii)*v8(jj,1)
    test(4,k)=test(4,k)+test(5,k)*test(5,k)
227 continue
c get eigenvalues, since actually not the same if more than 1 covariate
c modelled with splines
    do 237 i=1,nk4
    do 237 j=1,nk4
237 v6(i,j)=vv(i+11,j+11)

```

```

      do 235 i=1,nk3
      do 235 j=1,nk4
235  v6(i,j)=v6(i,j)-v6(i+1,j)
      do 236 j=1,nk3
      do 236 i=1,nk3
236  v6(i,j)=v6(i,j)-v6(i,j+1)
      do 238 i=1,nk3
      do 238 j=i,nk3
      v8(i,j)=0
      do 239 ii=1,i
      do 239 jj=1,j
239  v8(i,j)=v8(i,j)+v7(ii,i)*v6(ii,jj)*v7(jj,j)
      v8(j,i)=v8(i,j)
238  continue
      call eigen(v8,nk3,nk4,eig(1,2*k),v6,1,iflg)
      call sortg(eig(1,2*k),iwork,1,nk3)
      do 240 i=1,nk3
240  v6(i,1)=0
      test(5,k)=1-qfg(eig(1,2*k),v6,iwork,nk3,0.d0,test(4,k),5000,
     &1.d-5,iwork(nk4+1),trace,efault)
      test(6,k)=0
      do 335 i=1,nk4
335  test(6,k)=test(6,k)+eig(i,2*k)
      50 continue
100  format(5e15.8)
      return
      end

      subroutine cest(nknot,tknot,np,beta,var,lvar,smin,smax,nest,est,
     &nfx,nsx,iord)
c splines contain all B-spline terms
c estimates are calculated at nest equally spaced points from
c smin to smax
c c iord is the order of the spline (3 for cubic)
c On output est contains the time ordinates (col 1) and
c the 1hr estimate (col 2*j) and its variance (col 2*j+1)
      double precision tknot(-2:nknot+4),beta(np),var(lvar,np)
      double precision est(nest,1+2*nsx),b(4),smin,smax
      ior=iord+1
      nbas=nknot+ior
      do 10 i=1,nest
      est(i,1)=smin+(i-1)*(smax-smin)/(nest-1)
      call spln(nknot,iord,tknot(1-ior),est(i,1),1,int,b,1)
      do 12 j=1,nsx
      l1=nfx+(j-1)*nbas+int-1
      j1=j*2
      j2=j1+1
      est(i,j1)=0
      est(i,j2)=0
      do 17 i1=1,ior
      est(i,j1)=est(i,j1)+b(i1)*beta(l1+i1)
      do 17 k1=1,ior
      est(i,j2)=est(i,j2)+b(i1)*b(k1)*var(l1+i1,l1+k1)
17    continue
12    continue
10    continue
      return
      end

      subroutine cestb(ntime,times,smin,smax,inter,tt,est,nfx,nsx,
     &beta,var,lvar,np,nknot,iord)
c splines contain all B-spline terms (intervals in inter, B-splines in
c tt) in consecutive entries in beta and var (coefs and variances)

```

```

c iord is the order of the spline (3 for cubic)
c smin and smax are min and max values for plot, other step fcn
c boundries given by times. nfx is number fixed and nsx # spline
c terms. On output est contains the time ordinates (col 1) and
c the lhr estimate (col 2*j) and its variance (col 2*j+1)
    double precision times(ntime),smin,smax,tt(ntime+1,4)
    double precision est(2*(ntime+1),1+2*nsx),beta(np),var(lvar,np)
    integer inter(1)
    ior=iord+1
    nbas=nknot+ior
    est(1,1)=smin
    do 5 j=1,nsx
      l1=nfx+(j-1)*nbas+inter(1)-1
      j1=2*j
      j2=2*j+1
      est(1,j1)=0
      est(1,j2)=0
    do 17 i1=1,ior
      est(1,j1)=est(1,j1)+tt(1,i1)*beta(l1+i1)
      do 17 k1=1,ior
        est(1,j2)=est(1,j2)+tt(1,i1)*tt(1,k1)*var(l1+i1,l1+k1)
17  continue
5   continue
    do 10 i=2,ntime+1
      ii1=2*(i-1)
      est(ii1,1)=times(i-1)
      do 11 j=1,nsx
        est(ii1,2*j)=est(ii1-1,2*j)
11   est(ii1,2*j+1)=est(ii1-1,2*j+1)
      ii1=ii1+1
      est(ii1,1)=est(ii1-1,1)
      do 6 j=1,nsx
        l1=nfx+(j-1)*nbas+inter(i)-1
        j1=2*j
        j2=2*j+1
        est(ii1,j1)=0
        est(ii1,j2)=0
      do 18 i1=1,ior
        est(ii1,j1)=est(ii1,j1)+tt(i,i1)*beta(l1+i1)
        do 18 k1=1,ior
          est(ii1,j2)=est(ii1,j2)+tt(i,i1)*tt(i,k1)*var(l1+i1,l1+k1)
18  continue
6   continue
10  continue
      ii1=2*(ntime+1)
      est(ii1,1)=smax
      do 21 j=1,nsx
        est(ii1,2*j)=est(ii1-1,2*j)
21   est(ii1,2*j+1)=est(ii1-1,2*j+1)
      return
    end

      subroutine intsb2(ntud,s,t,no)
c subroutine to determine knots . Data must be sorted on s
c ntud is the # of knots (input).
c Ij is (t(j-1),t(j)]. t(0) is set to s(1) and t(ntud+1) to s(no)
      double precision s(no),t(0:(ntud+1))
      nft=no
      nfk=max(int(nft/float(ntud+1)+.5),1)
      t(0)=s(1)
      l1=1
      k=0
15  if (l1.ge.no) then

```

```

t(k+1)=s(no)
if (t(k+1).eq.t(k)) then
  ntud=k-1
else
  ntud=k
endif
c   write (6,100) ntud
100  format ('# nkots reset in intsub to',i5)
      return
endif
if (t(k).lt.s(l1+1)) then
  k=k+1
nft=no-l1
  nfk=max(int(nft/float(ntud-k+2)+.5),1)
l1=l1+nfk
t(k)=s(l1)
  if (k.eq.ntud) then
    t(k+1)=s(no)
    if (t(k+1).eq.t(k)) ntud=k-1
    return
  else
    go to 15
  endif
  else
    l1=l1+1
    go to 15
  endif
end

subroutine spln(nk,m,wk,y,n,inter,x,lx)
c nk # knots, m=highest degree polynomial in spline(0 for
c const, 3 for cubic), wk(1) to wk(nk)=interior knots (ordered smallest
c to largest), wk(-m+1) to wk(nk+m) must be augmented knot sequence
c on output x will be an (nx[m+1]) matrix giving the values of the
c active basis functions at each point. inter(ii) gives
c which interval y(ii) falls in (y is the vector of input points)
c knot intervals are of the form Ij=(t(j-1),tj].
c n is the number of points where values are calculated
c lx is actual row dim of x.
c NOTE: if m=0 then only uses interior knots 1 to nk. If
c m>0 then knots bounding the range of y also need to be included.
c if m>0 and y<wk(0) or y>wk(nk+1) then returns all 0's for inter
c and x(ii,..)
      double precision y(n),x(lx,m+1),wk(-m+1:nk+m),bj0,bj1
      double precision hj0,hj1
      integer inter(n)
      nb=nk+m+1
      l=1
      do 99 ii=1,n
      if (m.gt.0.and.(y(ii).lt.wk(0).or.y(ii).gt.wk(nk+1))) then
        inter(ii)=0
        do 5 jj=1,m+1
5       x(ii,jj)=0
        l=1
        go to 99
      endif
      do 10 j=1,m+1
10     x(ii,j)=0
22     if (y(ii).gt.wk(l)) then
        l=l+1
        go to 22
      endif
      inter(ii)=l

```

```

x(ii,1)=1
if (m.gt.0) then
  do 25 k=1,m
    max=l+k
    lc=1+k
    x(ii,lc)=x(ii,lc-1)*(y(ii)-wk(max-k-1))/(wk(max-1)-wk(max-k-1))
    if (k.gt.1) then
      do 26 k1=max-1,l+1,-1
      k3=k1-1+1
      bj1=x(ii,k3)
      hj1=(wk(k1)-y(ii))/(wk(k1)-wk(k1-k))
      bj0=x(ii,k3-1)
      hj0=(y(ii)-wk(k1-k-1))/(wk(k1-1)-wk(k1-k-1))
      x(ii,k3)=hj0*bj0+hj1*bj1
26    continue
    endif
    x(ii,1)=x(ii,1)*(wk(1)-y(ii))/(wk(1)-wk(1-k))
25    continue
  endif
  if (ii.lt.n) then
    if (y(ii+1).lt.y(ii)) l=1
  endif
99  continue
100 format (6e13.6)
      return
    end

    subroutine cks(smin,smax,tknot,nknot,iosp)
c subroutine to Complete the Knot Sequence.
c smin and smax are smallest and largest values of var modelled by spline
c tknot has interior knots in positions 1 to nknot
c iosp is the power of the largest polynomial in the spline
c (ie 3 for cubic)
      double precision smin,smax,tknot((1-iosp):(nknot+iosp))
      ntud=nknot+1
      if (smin.lt.tknot(1)) then
        tknot(0)=smin
      else
        tknot(0)=tknot(1)-1
      endif
      if (smax.gt.tknot(nknot)) then
        tknot(ntud)=smax
      else
        tknot(ntud)=tknot(nknot)+1
      endif
      do 717 jj=1,iosp-1
        tknot(ntud+jj)=tknot(ntud)
        tknot(-jj)=tknot(0)
717  continue
      return
    end

    subroutine penco(nk,wk,penm,mr)
c nk is # knots, wk is augmented knot seq.
c This subroutine calculates the penalty matrix for integrated squared
c first derivative penalty, for quadratic spline.
c The first col of penm will have the main
c diagonal, the 2nd col the next diag ... the 3rd col the 2nd diag from
c the main.
c mr is the actual row dim of penm
c iosp is the order of the spline. must be 2 for this routine
      double precision wk(1:nk),penm(mr,2)
      nb=nk+1

```

```

penm(1,1)=1
penm(nb,2)=0
penm(nb,1)=1
penm(1,2)=-1
do 20 i=2,nk
penm(i,1)=2
penm(i,2)=-1
20 continue
return
end

subroutine penql(nk,wk,penm,mr,iosp)
c nk is # knots, wk is augmented knot seq.
c This subroutine calculates the penalty matrix for integrated squared
c first derivative penalty, for quadratic spline.
c The first col of penm will have the main
c diagonal, the 2nd col the next diag ... the 3rd col the 2nd diag from
c the main.
c mr is the actual row dim of penm
c iosp is the order of the spline. must be 2 for this routine
double precision wk(-(iosp-1):nk+iosp),penm(mr,(iosp+1)),tl(4),
&tu(4),a(4),b(4)
double precision t1,t2,t3,t4
nb=nk+iosp+1
nw=iosp+1
do 10 i=1,nb
do 10 j=1,nw
10 penm(i,j)=0
t1=-2/(wk(1)-wk(-1))
t1(2)=-t1(1)
t1(3)=0
do 20 l=1,nk+1
c tl has the value of the 2nd deriv at the lower endpoint of the lth
c interval, tu at the upper (2nd derivs of basis fcns are linear on
c each interval
t1=wk(l)-wk(l-1)
t4=wk(l)+wk(l-1)
t2=t1*t4/2
t3=t1*(t4*t4-wk(l)*wk(l-1))/3
c note that t3 is (wk(l)^3-wk(l-1)^3)/3
tu(1)=0
tu(2)=-4/(wk(l+1)-wk(l-1))
if (l.eq.nk+1) tu(2)=tu(2)/2
tu(3)=-tu(2)
c calc slopes and intercepts on interval 1:
do 21 j=1,nw
b(j)=(tu(j)-t1(j))/(wk(l)-wk(l-1))
a(j)=tu(j)-b(j)*wk(l)
21 continue
do 22 j=1,nw
l1=l+j-1
j2=j-1
do 22 k=j,nw
penm(l1,k-j2)=penm(l1,k-j2)+2*(a(j)*a(k)*t1+(a(j)*b(k) +
&a(k)*b(j))*t2+b(j)*b(k)*t3)
22 continue
t1(1)=tu(2)
t1(2)=tu(3)
t1(3)=0
20 continue
return
end

```

```

subroutine dwch(e,c,ndf,k,stat,pv,iwk)
c subroutine to call function qfg (in dwchis)
c k is the # terms
c e is the eigenvalues (weights in weighted sum)
c c the noncentrality parameters, ndf the number of
c degrees of freedom. Note the sigma term in qfg is assumed to
c be 0 to simplify this call
c stat is the value of the quad form, calcs prob>stat and puts the
c result in pv.
c for interp of ifault and trace see qfg
    double precision e(k),c(k),stat,pv,trace(7),acc,qfg,sig
    integer ndf(k),iwk(k),efault
    acc=1.d-5
    sig=0.d0
    lim=10000
    pv=1-qfg(e,c,ndf,k,sig,stat,lim,acc,iwk,trace,efault)
    return
end

    double precision FUNCTION QFg(ALB,ANC,N,IRR,SIGMA,CC,LIM1,
&ACC,ITH,TRACE,IFAULT)
c*** from statlib, Wed Dec 12 08:23:41 EST 1990 ***
C
C      ALGORITHM AS 155 APPL. STATIST. (1980) VOL.29, NO.3
C
C      Distribution of a linear combination of non-central chi-squared
C      random variables.
C
c      calcs dist of sum(alb(j)*X(j))+sigma*X(0) where X(j) is chi-square
c      with n(j) df and noncent par anc(j), and X(0) is std normal
c      (hopefully) converted to dble precision (no real reason to do this
c      since don't comput results to that accuracy anyway)
c      args: alb input values of weights
c      anc input values of noncentrality parameters
c      n input integer values of df
c      irr input # chi-square terms (lenght of alb,anc,n)
c      sigma, see formula above
c      cc point at which df is evaluated
c      lim1 max # terms in integral
c      acc requested accuracy
c      ith is an integer working vector.
c      output: trace(1) abs val sum trace (2) total # int terms (3) number of
c      integrations (4) int interval i main integration (5) truncation point
c      in main integration (6) std dev of conv factor (7) number cycles to
c      locate integ params; ifault=0 ok, =1 req acc not obtained, =2 round
c      off error may be sig (this is computed based on single precision so
c      the dble precision calcs here are hopefully ok), =3 invalid params
c      =4 unable to locate int params.
        implicit double precision (a-h,o-z)
        INTEGER IRR,LIM1,IFAULT
        double precision SIGMA,CC,ACC
        double precision TRACE(7),ALB(IRR),ANC(IRR)
        INTEGER N(IRR),ITH(IRR)
        INTEGER J,NJ,NT,NTM
        double precision ACC1,ALMX,XLIM,XNT,XNTM
        double precision UTX,TAUSQ,SD,AINTV,AINTV1,X,UP,UN,D1,D2,ALJ,ANCJ
        DOUBLE PRECISION AINTL,ERSM
        double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,ZERO,HALF,
&ONE,TWO,FOUR,
1       SIXTN,FOURP5,PT07,PT2,QUART,TEN,PT33,PT67,PT75,ONEPT5,THREE,
2       ONEPT1
        INTEGER ICOUNT,IR,LIM
        LOGICAL NDTSR,FAIL

```

```

COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1 ICOUNT,IR,NDTSRT,FAIL,LIM
DATA ZERO /0.0d0/,HALF/0.5d0/,ONE/1.0d0/,TWO/2.0d0/,FOUR/4.0d0/,
&SIXTM/16.0d0/,FOURP5/4.5d0/,PT07/0.07d0/,PT2/0.2d0/,
&QUART/0.25d0/,TEN/10.0d0/,PT33/0.33d0/,PT67/0.67d0/,PT75/0.75d0/,
&ONEPT5/1.5d0/,THREE/3.0d0/,ONEPT1/1.1d0/
C
C      IROUND(X) = INT(X + SIGN(HALF,X))
C
C      Setting constants in COMMON.    ALN28 = ln(2) / 8.
C
C      PI = 3.14159265358979
C      ALN28 = 0.08664339758
C
C      C = CC
C      IR = IRR
C      LIM = LIM1
C      DO 10 J = 1,7
C      TRACE(J) = ZERO
10   CONTINUE
C      IFAULT = 0
C      ICOUNT = 0
C      AINTL = ZERO
C      ERSR = ZERO
C      QFG = -ONE
C      ACC1 = ACC
C      NDTSR = .TRUE.
C      FAIL = .FALSE.
C
C      Find mean, sd, max & min of ALB.
C      Check that parameter values are valid.
C
C      XLIM = LIM
C      SIGSQ = SIGMA**2
C      SD = SIGSQ
C      ALMAX = ZERO
C      ALMIN = ZERO
C      AMEAN = ZERO
C      J = 1
20   IF (.NOT.(J.LE.IR)) GO TO 60
      NJ = N(J)
      ALJ = ALB(J)
      ANCJ = ANC(J)
      IF (.NOT.(NJ.LT.0.OR.ANCJ.LT.ZERO)) GO TO 30
      IFAULT = 3
      GO TO 260
30   SD = SD + ALJ**2*(2*NJ + FOUR*ANCJ)
      AMEAN = AMEAN + ALJ*(NJ + ANCJ)
      IF (.NOT.(ALMAX.LT.ALJ)) GO TO 40
      ALMAX = ALJ
      GO TO 50
40   IF (.NOT.(ALMIN.GT.ALJ)) GO TO 50
      ALMIN = ALJ
50   J = J + 1
      GO TO 20
60   IF (.NOT.(SD.EQ.ZERO)) GO TO 80
      IF (.NOT.(C.GT.ZERO)) GO TO 70
      QFG = ONE
      GO TO 260
70   QFG = ZERO
      GO TO 260
80   IF (.NOT.(ALMIN.EQ.ZERO.AND.ALMAX.EQ.ZERO.AND.SIGMA.EQ.ZERO))
1   GO TO 90

```

```

IFAULT = 3
GO TO 260
90 SD = SQRT(SD)
IF (.NOT.(ALMAX.LT.-ALMIN)) GO TO 100
ALMX = -ALMIN
GO TO 110
100 ALMX = ALMAX
C
C Starting values for FINDU * CTFF.
C
110 UTX = SIXTN/SD
UP = FOURP5/SD
UN = -UP
C
C Truncation point with no convergence factor.
C
CALL FINDU (N,ALB,ANC,UTX,HALF*ACC1)
C
C Does convergence factor help ?
C
IF (.NOT.(C.NE.ZERO.AND.ALMX.GT.PT07*SD)) GO TO 130
TAUSQ = QUART*ACC1/CFE(N,ALB,ANC,ITH,C)
IF (.NOT.(FAIL)) GO TO 120
FAIL = .FALSE.
GO TO 130
120 IF (.NOT.(TRUNCN(N,ALB,ANC,UTX,TAUSQ).LT.PT2*ACC1)) GO TO 130
SIGSQ = SIGSQ + TAUSQ
CALL FINDU (N,ALB,ANC,UTX,QUART*ACC1)
TRACE(6) = SQRT(TAUSQ)
130 TRACE(5) = UTX
ACC1 = HALF*ACC1
C
C Find 'range' of distribution, quit if outside of this.
C
140 D1 = CTFF(N,ALB,ANC,ACC1,UP) - C
IF (.NOT.(D1.LT.ZERO)) GO TO 150
QFG = ONE
GO TO 260
150 D2 = C - CTFF(N,ALB,ANC,ACC1,UN)
IF (.NOT.(D2.LT.ZERO)) GO TO 160
QFG = ZERO
GO TO 260
C
C Find integration interval.
C
160 IF (.NOT.(D1.GT.D2)) GO TO 170
AINTV = D1
GO TO 180
170 AINTV = D2
180 AINTV = TWO*PI/AINTV
C
C Calculate number of terms required for main & auxiliary
C integrations.
C
XNT = UTX/AINTV
XNTM = THREE/SQRT(ACC1)
IF (.NOT.(XNT.GT.XNTM*ONEPT5)) GO TO 220
IF (.NOT.(XNTM.GT.XLIM)) GO TO 190
IFault = 1
GO TO 260
C
C Parameters for auxiliary integration.
C

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```

190  NTM = IROUND(XNTM)
     AINTV1 = UTX/XNTM
     X = TWO*PI/AINTV1
     IF (.NOT.(X.LE.ABS(C))) GO TO 200
     GO TO 220
C
C   Calculate convergence factor.
C
200  TAUSQ = CFE(N,ALB,ANC,ITH,C - X) + CFE(N,ALB,ANC,ITH,C + X)
     TAUSQ = PT33*ACC1/(ONEPT1*TAUSQ)
     IF (.NOT.(FAIL)) GO TO 210
     GO TO 220
210  ACC1 = PT67*ACC1
C
C   Auxiliary integration.
C
     CALL INTEGR (N,ALB,ANC,NTM,AINTV1,TAUSQ,.FALSE.)
     XLIM = XLIM - XNTM
     SIGSQ = SIGSQ + TAUSQ
     TRACE(3) = TRACE(3) + 1
     TRACE(2) = TRACE(2) + NTM + 1
C
C   Find truncation point with new convergence factor.
C
     CALL FINDU (N,ALB,ANC,UTX,QUART*ACC1)
     ACC1 = PT75*ACC1
     GO TO 140
C
C   Main integration.
C
220  TRACE(4) = AINTV
     IF (.NOT.(XNT.GT.XLIM)) GO TO 230
     IFAULT = 1
     GO TO 260
230  NT = IROUND(XNT)
     CALL INTEGR (N,ALB,ANC,NT,AINTV,ZERO,.TRUE.)
     TRACE(3) = TRACE(3) + 1
     TRACE(2) = TRACE(2) + NT + 1
     QFg = HALF - AINTL
     TRACE(1) = ERSM
     UP = ERSM
C
C   Test whether round-off error could be significant.
C   Allow for radix 8 or 16 machines.
C
     X = UP + ACC/TEN
     J = 1
240  IF (.NOT.(J.LE.8)) GO TO 260
     IF (.NOT.(J*X.EQ.J*UP)) GO TO 250
     IFAULT = 2
250  J = J*2
     GO TO 240
260  TRACE(7) = ICOUNT
     RETURN
     END
C
C   SUBROUTINE GCNTR
C
C   Count number of calls to ERRBD, TRUNCN & CFE.
C
     implicit double precision (a-h,o-z)
     DOUBLE PRECISION AINTL,ERSM
     double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C

```

```

INTEGER ICOUNT,IR,LIM
LOGICAL NDTSR,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1  ICOUNT,IR,NDTSRT,FAIL,LIM
C
  ICOUNT = ICOUNT + 1
  IF (.NOT.(ICOUNT.GT.LIM)) GO TO 20
c   WRITE (6,10)
c 10  FORMAT (' qfg: cannot locate integration parameters')
c   STOP
  call intpr("emerg stop in qfg",17,icount,1)
  call intpr("maybe in inf loop",17,icount,1)
20  RETURN
  END
C
  double precision FUNCTION ALOG1(X,FIRST)
C
C  If FIRST then return ln(1 + x) else ln(1 + x) - x.
C
  implicit double precision (a-h,o-z)
  double precision X
  LOGICAL FIRST
  double precision S,S1,TERM,Y,AK,PT1,ONE,TWO,THREE
  DATA PT1/0.1d0/,ONE/1.0d0/,TWO/2.0d0/,THREE/3.0d0/
C
  F1(I) = S + TERM/AK
C
  IF (.NOT.(ABS(X).GT.PT1)) GO TO 20
  IF (.NOT.(FIRST)) GO TO 10
  ALOG1 = LOG(ONE + X)
  GO TO 70
10  ALOG1 = LOG(ONE + X) - X
  GO TO 70
20  Y = X/(TWO + X)
  TERM = TWO*Y**3
  AK = THREE
  IF (.NOT.(FIRST)) GO TO 30
  S = TWO
  GO TO 40
30  S = -X
40  S = S*Y
  Y = Y**2
  S1 = F1(0)
50  IF (.NOT.(S1.NE.S)) GO TO 60
  AK = AK + TWO
  TERM = TERM*Y
  S = S1
  S1 = F1(0)
  GO TO 50
60  ALOG1 = S
70  RETURN
  END
C
  double precision FUNCTION EXP1(X)
  implicit double precision (a-h,o-z)
  double precision X
  double precision ZERO,NEG50
  DATA ZERO/0.0d0/,NEG50/-50.0d0/
C
  IF (.NOT.(X.LT.NEG50)) GO TO 10
  EXP1 = ZERO
  GO TO 20
10  EXP1 = EXP(X)

```

```

20 RETURN
END
C
SUBROUTINE ORDER (ALB,ITH)
C
C Find order of absolute values of ALB.
C
implicit double precision (a-h,o-z)
double precision ALB(*)
INTEGER ITH(*)
INTEGER J,K,K1,ITHK
double precision ALJ
DOUBLE PRECISION AINTL,ERSM
double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C
INTEGER ICOUNT,IR,LIM
LOGICAL NDTSR,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1 ICOUNT,IR,NDTSRT,FAIL,LIM
C
J = 1
10 IF (.NOT.(J.LE.IR)) GO TO 60
ALJ = ABS(ALB(J))
K = J - 1
20 IF (.NOT.(K.GT.0)) GO TO 40
ITHK = ITH(K)
K1 = K + 1
IF (.NOT.(ALJ.GT.ABS(ALB(ITHK)))) GO TO 50
ITH(K1) = ITHK
GO TO 30
30 K = K - 1
GO TO 20
40 K = 0
K1 = 1
50 ITH(K1) = J
J = J + 1
GO TO 10
60 NDTSR = .FALSE.
RETURN
END
C
double precision FUNCTION ERRBD(N,ALB,ANC,UU,CX)
C
C Find bound on tail probability using mgf.
C Cut-off point returned to CX.
C
implicit double precision (a-h,o-z)
double precision U,UU,CX
INTEGER N(*)
double precision ALB(*),ANC(*)
double precision SUM1,ALJ,ANCJ,X,Y,CONST
INTEGER J,NJ
DOUBLE PRECISION AINTL,ERSM
double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,HALF,ONE,TWO
INTEGER ICOUNT,IR,LIM
LOGICAL NDTSR,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1 ICOUNT,IR,NDTSRT,FAIL,LIM
DATA HALF/0.5d0/,ONE/1.0d0/,TWO/2.0d0/
C
CALL GCNTR
U = UU
CONST = U*SIGSQ
SUM1 = U*CONST

```

```

U = TWO*U
J = IR
10 IF (.NOT.(J.GT.0)) GO TO 20
NJ = N(J)
ALJ = ALB(J)
ANCJ = ANC(J)
X = U*ALJ
Y = ONE - X
CONST = CONST + ALJ*(ANCJ/Y + NJ)/Y
SUM1 = SUM1 + ANCJ*(X/Y)**2
SUM1 = SUM1 + NJ*(X**2/Y + ALOG1(-X,.FALSE.))
J = J - 1
GO TO 10
20 ERRBD = EXP1(-HALF*SUM1)
CX = CONST
RETURN
END

C
C      double precision FUNCTION CTFF(N,ALB,ANC,ACCX,UPN)
C
C      Find CTFF so that P(QF > CTFF) < ACCX if UPN > 0;
C      P(QF < CTFF) < ACCX otherwise.
C
C      implicit double precision (a-h,o-z)
double precision ACCX,UPN
INTEGER N(*)
double precision ALB(*),ANC(*)
double precision U1,U2,U,RB,CONST,C1,C2
DOUBLE PRECISION AINTL,ERSM
double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,ZERO,ONE,TWO
INTEGER ICOUNT,IR,LIM
LOGICAL NDTSRT,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1 ICOUNT,IR,NDTSRT,FAIL,LIM
DATA ZERO/0.0d0/,ONE/1.0d0/,TWO/2.0d0/
C
      F1(I) = U2/(ONE + U2*RB)
      F2(I) = (C1 - AMEAN)/(C2 - AMEAN)
C
      U2 = UPN
      U1 = ZERO
      C1 = AMEAN
      IF (.NOT.(U2.GT.ZERO)) GO TO 10
      RB = ALMAX
      GO TO 20
10     RB = ALMIN
20     RB = TWO*RB
      U = F1(0)
30     IF (.NOT.(ERRBD(N,ALB,ANC,U,C2).GT.AC CX)) GO TO 40
      U1 = U2
      C1 = C2
      U2 = TWO*U2
      U = F1(0)
      GO TO 30
40     U = F2(0)
50     IF (.NOT.(U.LT.0.9)) GO TO 80
      U = (U1 + U2)/TWO
      IF (.NOT.(ERRBD(N,ALB,ANC,U/(ONE + U*RB),CONST).GT.AC CX)) GO TO 60
      U1 = U
      C1 = CONST
      GO TO 70
60     U2 = U
      C2 = CONST

```

```

70   U = F2(0)
     GO TO 50
80   CTFF = C2
     UPN = U2
     RETURN
     END

C
C      double precision FUNCTION TRUNCN(N,ALB,ANC,UU,TAUSQ)
C
C      Bound integration error due to truncation at U.
C
implicit double precision (a-h,o-z)
double precision U,UU,TAUSQ
INTEGER N(*)
double precision ALB(*),ANC(*)
double precision SUM1,SUM2,PROD1,PROD2,PROD3,ALJ,ANCJ,X,Y,
&ERR1,ERR2
INTEGER J,NJ,NS
DOUBLE PRECISION AINTL,ERSM
double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,ZERO,QUART,
&HALF,ONE,TWO
INTEGER ICOUNT,IR,LIM
LOGICAL NDTSRT,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1  ICOUNT,IR,NDTSRT,FAIL,LIM
DATA ZERO/0.0d0/,QUART/0.25d0/,HALF/0.5d0/,ONE/1.0d0/,TWO/2.0d0/
C
CALL GCNTR
U = UU
SUM1 = ZERO
PROD2 = ZERO
PROD3 = ZERO
NS = 0
SUM2 = (SIGSQ + TAUSQ)*U**2
PROD1 = TWO*SUM2
U = TWO*U
J = 1
10  IF (.NOT.(J.LE.IR)) GO TO 40
    ALJ = ALB(J)
    ANCJ = ANC(J)
    NJ = N(J)
    X = (U*ALJ)**2
    SUM1 = SUM1 + ANCJ*X/(ONE + X)
    IF (.NOT.(X.GT.ONE)) GO TO 20
    PROD2 = PROD2 + NJ*LOG(X)
    PROD3 = PROD3 + NJ*ALOG1(X,.TRUE.)
    NS = NS + NJ
    GO TO 30
20  PROD1 = PROD1 + NJ*ALOG1(X,.TRUE.)
30  J = J + 1
    GO TO 10
40  SUM1 = HALF*SUM1
    PROD2 = PROD1 + PROD2
    PROD3 = PROD1 + PROD3
    X = EXP1(-SUM1 - QUART*PROD2)/PI
    Y = EXP1(-SUM1 - QUART*PROD3)/PI
    IF (.NOT.(NS.EQ.0)) GO TO 50
    ERR1 = ONE
    GO TO 60
50  ERR1 = X*TWO/NS
60  IF (.NOT.(PROD3.GT.ONE)) GO TO 70
    ERR2 = 2.5d0*Y
    GO TO 80

```

```

70    ERR2 = ONE
80    IF (.NOT.(ERR2.LT.ERR1)) GO TO 90
     ERR1 = ERR2
90    X = HALF*SUM2
     IF (.NOT.(X.LE.Y)) GO TO 100
     ERR2 = ONE
     GO TO 110
100   ERR2 = Y/X
110   IF (.NOT.(ERR1.LT.ERR2)) GO TO 120
     TRUNCN = ERR1
     GO TO 130
120   TRUNCN = ERR2
130   RETURN
     END

C
C      SUBROUTINE FINDU (N,ALB,ANC,UTX,ACCX)
C
C      Find U such that TRUNCN(U) < ACCX & TRUNCN(U / 1.2) > ACCX.
C
C      implicit double precision (a-h,o-z)
C      double precision UTX,ACCX
C      INTEGER N(*)
C      double precision ALB(*),ANC(*)
C      double precision U,UT
C      double precision DIVIS(4)
C      INTEGER I
C      DOUBLE PRECISION AINTL,ERSM
C      double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,FOUR
C      INTEGER ICOUNT,IR,LIM
C      LOGICAL NDTSRT,FAIL
C      COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1      ICOUNT,IR,NDTSRT,FAIL,LIM
      DATA DIVIS/2.0d0,1.4d0,1.2d0,1.1d0/, FOUR/4.0d0/
C
C      UT = UTX
C      U = UT/FOUR
C      IF (.NOT.(TRUNCN(N,ALB,ANC,U,ZERO).GT.AC CX)) GO TO 20
C      U = UT
10     IF (.NOT.(TRUNCN(N,ALB,ANC,U,ZERO).GT.AC CX)) GO TO 40
C      UT = UT*FOUR
C      U = UT
C      GO TO 10
20     UT = U
C      U = U/FOUR
30     IF (.NOT.(TRUNCN(N,ALB,ANC,U,ZERO).LE.AC CX)) GO TO 40
C      UT = U
C      U = U/FOUR
C      GO TO 30
40     DO 50 I = 1,4
C      U = UT/DIVIS(I)
C      IF (.NOT.(TRUNCN(N,ALB,ANC,U,ZERO).LE.AC CX)) GO TO 50
C      UT = U
50     CONTINUE
C      UTX = UT
C      RETURN
     END

C
C      SUBROUTINE INTEGR (N,ALB,ANC,NTERM,AINTRV,TAUSQ,MAIN)
C
C      Carry out integration with NTERM terms, at interval AINTRV.
C      If not MAIN then multiply integrand by 1 - exp(-0.5 * TAUSQ *
C      U**2).
C

```

```

implicit double precision (a-h,o-z)
INTEGER NTERM
double precision AINTRV,TAUSQ
LOGICAL MAIN
INTEGER N(*)
double precision ALB(*),ANC(*)
double precision AINPI,U,SUM1,SUM2,SUM3,X,Y,Z
INTEGER K,J,NJ
DOUBLE PRECISION AINTL,ERSM
double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,QUART,
&HALF,ONE,TWO
INTEGER ICOUNT,IR,LIM
LOGICAL NDTSRT,FAIL
COMMON /QFCOM/ AINTL,ERSM,PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,
1 ICOUNT,IR,NDTSRT,FAIL,LIM
DATA QUART/0.25d0/,HALF/0.5d0/,ONE/1.0d0/,TWO/2.0d0/
C
      AINPI = AINTRV/PI
      K = NTERM
10   IF (.NOT.(K.GE.0)) GO TO 50
      U = (K + HALF)*AINTRV
      SUM1 = -TWO*U*C
      SUM2 = ABS(SUM1)
      SUM3 = -HALF*SIGSQ*U**2
      J = IR
20   IF (.NOT.(J.GT.0)) GO TO 30
      NJ = N(J)
      X = TWO*ALB(J)*U
      Y = X**2
      SUM3 = SUM3 - QUART*NJ* ALOG1(Y,.TRUE.)
      Y = ANC(J)*X/(ONE + Y)
      Z = NJ*ATAN(X) + Y
      SUM1 = SUM1 + Z
      SUM2 = SUM2 + ABS(Z)
      SUM3 = SUM3 - HALF*X*Y
      J = J - 1
      GO TO 20
30   X = AINPI*EXP1(SUM3)/U
      IF (.NOT.(.NOT.MAIN)) GO TO 40
      X = X*(ONE - EXP1(-HALF*TAUSQ*U**2))
40   SUM1 = SIN(HALF*SUM1)*X
      SUM2 = HALF*SUM2*X
      AINTL = AINTL + SUM1
      ERSM = ERSM + SUM2
      K = K - 1
      GO TO 10
50   RETURN
END
C
      double precision FUNCTION CFE(N,ALB,ANC,ITH,X)
C
C      Coefficient of TAUSQ in error when convergence factor of
C      exp(-0.5 * TAUSQ * U**2) is used when df is evaluated at X.
C
      implicit double precision (a-h,o-z)
      double precision X
      INTEGER N(*),ITH(*)
      double precision ALB(*),ANC(*)
      double precision AXL,AXL1,AXL2,SXL,SUM1,ALJ
      INTEGER J,K,IT,ITK
      DOUBLE PRECISION AINTL,ERSM
      double precision PI,ALN28,SIGSQ,ALMAX,ALMIN,AMEAN,C,ZERO,ONE,
&FOUR,HUNDRED

```

```
INTEGER ICOUNT, IR, LIM
LOGICAL NDT SRT, FAIL
COMMON /QFCOM/ AINTL, ERSM, PI, ALN28, SIGSQ, ALMAX, ALMIN, AMEAN, C,
1 ICOUNT, IR, NDT SRT, FAIL, LIM
DATA ZERO/0.0d0/, ONE/1.0d0/, FOUR/4.0d0/, HUNDRD/100.0d0/
C
CALL GCNTR
IF (.NOT.(NDTSRT)) GO TO 10
CALL ORDER (ALB, ITH)
10 AXL = ABS(X)
SXL = SIGN(ONE, X)
SUM1 = ZERO
J = IR
20 IF (.NOT.(J.GT.0)) GO TO 70
IT = ITH(J)
IF (.NOT.(ALB(IT)*SXL.GT.ZERO)) GO TO 60
ALJ = ABS(ALB(IT))
AXL1 = AXL - ALJ*(N(IT) + ANC(IT))
AXL2 = ALJ/ALN28
IF (.NOT.(AXL1.GT.AXL2)) GO TO 30
AXL = AXL1
GO TO 60
30 IF (.NOT.(AXL.GT.AXL2)) GO TO 40
AXL = AXL2
40 SUM1 = (AXL - AXL1)/ALJ
K = J - 1
50 IF (.NOT.(K.GT.0)) GO TO 70
ITK = ITH(K)
SUM1 = SUM1 + (N(ITK) + ANC(ITK))
K = K - 1
GO TO 50
60 J = J - 1
GO TO 20
70 IF (.NOT.(SUM1.GT.HUNDRD)) GO TO 80
CFE = ONE
FAIL = .TRUE.
GO TO 90
80 CFE = 2** (SUM1/FOUR) / (PI*AXL**2)
90 RETURN
END
```